# Advanced Statistics III : Bayesian Ideas

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### Outline

Introduction

Somes Results

### A random parameter



### Theorem (Bayes)

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A)\,\mathbb{P}(A)}{\mathbb{P}(B)}$$

 Probablities as a partial belief

## An example: mamographies

- 1% of women at age forty who participate in routine screening have breast cancer
- 80% of women with breast cancer will get positive mammographies
- 9.6% of women without breast cancer will also get positive mammographies
- What is the probability that a women with positive mammography in a routine screening actually has breast cancer?

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- Two visions of the probabilities: from *outside* (frequency) and from *inside* (partial belief).

# Generic formulation - prior and posterior

- model: experiment  $X \sim dP(x|\theta) = p(x|\theta)d\mu(x)$ ,  $\theta \in \Theta$
- prior distribution:  $d\pi(\theta)$  contains knowledge on  $\theta$  anterior to the experiment
- *posterior* distribution  $d\Pi(\theta)$  with density:

$$\Pi(\theta) = p(\theta|X) = \frac{\pi(\theta)p(X|\theta)}{\int_{\theta} \pi(\theta)p(X|\theta)d\theta} \propto \pi(\theta)p(X|\theta)$$

- Idea: the experiment modifies the beliefs on  $\theta$
- conjugate prior:  $\pi(\theta)$  and  $\Pi(\theta|X)$  have a common pattern Example:  $X \sim \mathcal{N}(\theta, \sigma^2)$ ,  $\theta \sim \mathcal{N}(m, \tau)$ .
- Confidence interval, tests...

### Example: Binomial variables

Conjugate prior family: Beta(a, b)

$$\pi(\theta) = \frac{\theta^{a-1} (1-\theta)^{b-1}}{\beta(a,b)}$$

where

$$\beta(a,b) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}, \quad \Gamma(x) = \int_0^\infty t^{x-1}e^{-t}dt$$

- $\mathbb{E}\left[Beta(a,b)\right] = \frac{a}{a+b}$ ,  $\operatorname{Var}\left[Beta(a,b)\right] = \frac{ab}{(a+b)^2(a+b+1)}$
- Posterior distribution:

$$\Pi( heta) \sim Beta(x+a,n-x+b)$$
  $\hat{ heta}_{\pi}(X) = rac{X+a}{n+a+b}$ 

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## Consistency result

#### Definition

The posterior distribution  $\Pi$  is said to be consistent at  $\theta_0$  if for every neighbourhood U of  $\theta_0$ ,  $\Pi(U|X_1,\ldots,X_n)$  goes to 1 almost surely as  $X_1,\ldots,X_n\sim^{iid}P_{\theta_0}$ .

### Theorem (Doob)

Suppose  $\mathbb{P}(\cdot|\theta) \neq \mathbb{P}(\cdot|\theta')$  for  $\theta \neq \theta'$ . For any prior  $\pi$ , the posterior is consistent at every  $\theta$  except possibly on a set of  $\pi$ -measure zero.

### Theorem (Bernstein - Von Mises)

Under appropriate conditions [see Bickel-Docksum Section 5.5],

$$\mathcal{L}\left(\sqrt{n}\left(\hat{\theta}_{\pi}-\theta\right)|X_{1},\ldots,X_{n}\right)\rightarrow\mathcal{N}\left(0,\operatorname{Var}[P_{\theta}]\right)$$

almost-surely under  $P_{\theta}$  for all  $\theta$ .

# Risk, Bayesian and minimax approachs

#### Definition

• The (quadratic) risk of estimator  $\hat{\theta}$  under paramater  $\theta$  is:

$$R(\theta, \hat{\theta}) = \mathbb{E}_{X \sim P_{\theta}} \left\| \hat{\theta} - \theta \right\|^{2}$$

- Frequentist approach: worst case
  - ightharpoonup worst case risk:  $\overline{R}(\hat{\theta}) = \sup_{\theta \in \Theta} R(\theta, \hat{\theta})$
  - $ightharpoonup minimax risk: \overline{R} = \inf_{\hat{\theta}} \overline{R}(\hat{\theta}) \ (\Longrightarrow minimax estimator)$
- Bayesian approach: prior  $\pi$ 
  - ightharpoonup average risk under  $\pi$ :  $\underline{R}(\pi,\hat{\theta}) = \mathbb{E}_{\theta \sim \pi} R(\theta,\hat{\theta})$
  - $\triangleright$  bayesian risk under  $\pi$ :  $\underline{R}(\pi) = \inf_{\hat{\theta}} \underline{R}(\pi, \hat{\theta})$
  - $ightharpoonup maximin risk: <math>\underline{R} = \sup_{\pi} \underline{R}(\pi)$  ( $\Longrightarrow$  least favorable prior)

## **Properties**

#### Lemma

The bayesian risk is always smaller than the minimax risk:  $\underline{R} \leq \overline{R}$ 

#### Lemma

The bayesian risk:

$$\underline{R}(\pi) \triangleq \min_{\hat{\theta}} \mathbb{E}_{\Pi} \left[ \|\hat{\theta} - \theta\|^2 \right]$$

is reached by the the posterior mean  $\hat{\theta}_{\pi} = \mathbb{E}[\Pi]$ .

Theorem ("a Bayes rule with constant risk is minimax") If  $\hat{\theta}_{\pi}$  is a Bayes estimator with respect to a prior  $\pi$  and if  $R(\theta, \hat{\theta}_{\pi}) = \underline{R}(\pi)$  for all  $\theta$ , then  $\hat{\theta}_{\pi}$  is minimax and  $\pi$  is a least favorable prior.

## Minimax estimator in the Binomial setup

#### **Theorem**

Let  $X \sim \mathcal{B}(n, \theta), \theta \in [0, 1]$ 

• The minimax estimator is

$$\hat{\theta}_n(X) = \frac{X + \sqrt{n}/2}{n + \sqrt{n}}$$

- It has quadratic loss  $\bar{R}_n = \frac{1}{4(1+\sqrt{n})^2}$
- The least favorable prior is  $\pi_n = \text{Beta}(\sqrt{n}/2, \sqrt{n}/2)$

