# Machine learning and time: "time accounting" learning 

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14 mars 2014

[^0]- Finite network with nodes $\{1, \ldots, d\}$ : users of a social network, of an e-commerce platform, etc.
- For each node $j \in\{1, \ldots, d\}$ we observe the timestamps $\left\{t_{j, 1}, t_{j, 2}, \ldots\right\}$ of nodes' actions
- Goal: recover levels of interactions between users based on the timestamps patterns


## Introduction

From


Quantify interactions between users


- Do inference directly from actions of users
- Understand the community structure of users underlying the actions
- Exploit the hidden lower-dimensional structure of the network for inference/prediction
- Counting process $N_{j}(t)=\sum_{i \geq 1} \mathbf{1}_{t_{j, i} \leq t}$
- Data: a $d$-dimensional counting process $N=\left[N_{1}, \ldots, N_{d}\right]^{\top}$
- $d$ is large
- Observed on $[0, T]$. "Asymptotics" in $T \rightarrow+\infty$
- $N_{j}$ has intensity $\lambda_{j}$, namely
$\mathbb{P}(j$ does something at time $t$ knowning the past $)$

$$
=\mathbb{P}\left(N_{j} \text { has a jump in }[t, t+d t] \mid \mathcal{F}_{t}\right)=\lambda_{j}(t) d t
$$

for $j=1, \ldots, d$ where $\mathcal{F}_{t}$ some filtration

## Model: Multivariate Hawkes Process (MHP)

- MHP assumes an autoregressive structure on the intensities:

$$
\lambda_{j}(t)=\mu_{j}(t)+\int_{(0, t)} \sum_{k=1}^{d} \varphi_{j, k}(t-s) d N_{k}(s),
$$

- $\mu_{j}(t) \geq 0$ baseline intensity of the $j$-th coordinate
- $\varphi_{j, k}: \mathbb{R}^{+} \rightarrow \mathbb{R}^{+}$self-exciting component: influence of $k \rightarrow j$
- Write this in matrix form

$$
\lambda(t)=\mu+\int_{(0, t)} \varphi(t-s) d N(s)
$$

with $\boldsymbol{\mu}=\left[\mu_{1}, \ldots, \mu_{d}\right]^{\top}$ and $\varphi(t)=\left[\varphi_{j, k}(t)\right]_{1 \leq j, k \leq d}$.

- Notation:

$$
\int_{(0, t)} \varphi(t-s) d N_{j}(s)=\sum_{i: t_{j, i}<t} \varphi\left(t-t_{j, i}\right)
$$

Introduced by Hawkes in 1971

- Earthquakes and geophysics: Kagan and Knopoff (1981), Zhuang, Harte, Werner, Hainzl and Zhou (2012)
- Genomics: Reynaud-Bouret and Schbath (2010)
- High-frequency Finance: Bacry Delattre Hoffmann and Muzy (2013)
- Terrorist activity : Porter and White (2012)
- Neurobiology: Hansen, Reynaud-Bouret and Rivoirard (2012)
- Social networks : Carne and Sornette (2008), Simma and Jordan (2010), Zhou Song and Zha (2013)
- And even FPGA-based implementation: Guo and Luk (2013)


## A brief history of MHP

## THE GENESIS <br> BLOCK

Digital currency research and data
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Analyzing Trade Clustering To Predict Price Movement In Bitcoin Trading
Sep 19, 2013 Posted By Jonathan Heusser in Bitcoin 201, Economics, Featured, News, Trading Tagged Analysis, Bitcoin Trading

Parametric estimation (Maximum likelihood)

- First work: Ogata 78
- Simma and Jordan (2010), Zhou Song and Zha (2013) $\rightarrow$ Expected Maximization (EM) algorithms, with priors


## Non parametric estimation

- Marsan Lengliné (2008), generalized by Lewis, Mohler (2010) $\rightarrow$ EM for penalized likelihood function
$\rightarrow$ Monovariate Hawkes processes, Small amount of data, No theoretical results
- Reynaud-Bouret and Schbath (2010)
$\rightarrow$ Developed for small amount of data (Sparse penalization)
- Bacry and Muzy (2014)
$\rightarrow$ Larger amount of data

Dimension $d$ is large:

- Need a simple parametric model on $\mu$ and $\varphi$
- We want a convex optimization problem with smooth loss
- We want to encode some prior assumptions by penalizing this loss

Simple parametrization:

- Constant baselines $\mu_{j}(\cdot) \equiv \mu_{j}$
- Take

$$
\varphi_{j, k}(t)=a_{j, k} e^{-\alpha_{j, k} t}
$$

- $a_{j, k}=$ level of interaction between nodes $j$ and $k$
- $\alpha_{j, k}=$ lifetime of instantaneous excitation of node $j$ by node $k$ The matrix

$$
\boldsymbol{A}=\left[a_{j, k}\right]_{1 \leq j, k \leq d}
$$

is understood has a weighted adjacency matrix of mutual excitement of th nodes $\{1, \ldots, d\}$

- $\boldsymbol{A}$ is non-symmetric


## A simple parametrization of the MHP

We end up with intensities

$$
\lambda_{j, \theta}(t)=\mu_{j}+\int_{(0, t)} \sum_{k=1}^{d} a_{j, k} e^{-\alpha_{j, k}(t-s)} d N_{k}(s)
$$

for $j \in\{1, \ldots, d\}$ where

$$
\theta=[\mu, \boldsymbol{A}, \boldsymbol{\alpha}]
$$

with

- baselines $\mu=\left[\mu_{1}, \ldots, \mu_{d}\right]^{\top} \in \mathbb{R}_{+}^{d}$
- adjacencies $\boldsymbol{A}=\left[a_{j, k}\right]_{1 \leq j, k \leq d} \in \mathbb{R}_{+}^{d \times d}$
- decays $\boldsymbol{\alpha}=\left[\alpha_{j, k}\right]_{1 \leq j, k \leq d} \in \mathbb{R}_{+}^{d \times d}$


## A simple parametrization of the MHP

For $d=1$, intensity $\lambda_{\theta}$ looks like this:


Goodness-of-fit $=-\log$-likelihood is given by:

$$
-\ell_{T}(\theta)=\sum_{j=1}^{d}\left\{\int_{0}^{T}\left(\lambda_{j, \theta}(t)-1\right) d t-\int_{0}^{T} \log \lambda_{j, \theta}(t) d N_{j}(t)\right\}
$$

with

$$
\lambda_{j, \theta}(t)=\mu_{j}+\sum_{k=1}^{d} a_{j, k} \int_{(0, t)} \exp \left(-\alpha_{j, k}(t-s)\right) d N_{k}(s)
$$

where $\theta=(\boldsymbol{\mu}, \boldsymbol{A}, \boldsymbol{\alpha})$ with $\boldsymbol{\mu}=\left[\mu_{j}\right], \boldsymbol{A}=\left[A_{j, k}\right], \boldsymbol{\alpha}=\left[\alpha_{j, k}\right]$

## Prior encoding by penalization

## Prior assumptions

- Some users are basically inactive and react only if stimulated:
$\boldsymbol{\mu}$ is sparse
- Everybody does not interact with everybody:
$\boldsymbol{A}$ is sparse
- Interactions have community structure, possibly overlapping, a small number of factors explain interactions:
$\boldsymbol{A}$ is low-rank

- Decays $\alpha$ not sparse, but $\alpha_{j, k}$ should be regularized proportionaly to $a_{j, k}$

Standard convex relaxations [Tibshirani 01, ..., Srebro et al. 05, Bach 08, Candès \& Recht 08, ...]

- Tightest convex relaxation of $\|\boldsymbol{A}\|_{0}=\sum_{j, k} \mathbf{1}_{\boldsymbol{A}_{j, k}>0}$ is $\ell_{1}$-norm:

$$
\|\boldsymbol{A}\|_{1}=\sum_{j, k}\left|\boldsymbol{A}_{j, k}\right|
$$

- Tightest convex relaxation of rank is trace-norm:

$$
\|A\|_{*}=\sum_{j} \sigma_{j}(A)=\|\sigma(A)\|_{1}
$$

where $\sigma_{1}(A) \geq \cdots \geq \sigma_{d}(A)$ singular values of $\boldsymbol{A}$

## Prior encoding by penalization

So, we use the following penalizations

- Use $\ell_{1}$ penalization on $\boldsymbol{\mu}$
- Use $\ell_{1}$ penalization on $\boldsymbol{A}$
- Use trace-norm penalization on $\boldsymbol{A}$
- Use $\ell_{2}^{2}$ penalization on $\boldsymbol{\alpha}$, weighted by $\boldsymbol{A}$
[but other choices might be interesting...]
NB1: to induce sparsity AND low-rank on $\boldsymbol{A}$, we use the mixed penalization

$$
\boldsymbol{A} \mapsto w_{*}\|\boldsymbol{A}\|_{*}+w_{1}\|\boldsymbol{A}\|_{1}
$$

NB2: recent works by Richard et al (2013, 2014): better way to induce sparsity and low-rank than the sum, but not-scalable / non-convex


$\left\{\boldsymbol{A}:\|\boldsymbol{A}\|_{1} \leq 1\right\} \quad\left\{\boldsymbol{A}:\|\boldsymbol{A}\|_{1}+\|\boldsymbol{A}\|_{*} \leq 1\right\}$

The balls are computed on the set of $2 \times 2$ symmetric matrices, which is identified with $\mathbb{R}^{3}$.
[show video]

Finally, consider

$$
\begin{aligned}
\hat{\theta} \in \underset{\theta=(\mu, \boldsymbol{A}, \boldsymbol{\alpha})}{\operatorname{argmin}}\left\{-\frac{1}{T} \ell_{T}(\theta)+\right. & \tau\|\boldsymbol{\mu}\|_{1}+\gamma_{1}\|\boldsymbol{A}\|_{1} \\
& \left.+\gamma_{*}\|\boldsymbol{A}\|_{*}+\frac{\kappa}{2}\|\boldsymbol{A} \odot \boldsymbol{\alpha}\|_{F}^{2}\right\}
\end{aligned}
$$

where we recall

$$
-\frac{1}{T} \ell_{T}(\theta)=\frac{1}{T} \sum_{j=1}^{d}\left\{\int_{0}^{T} \lambda_{j, \theta}(t) d t-\int_{0}^{T} \log \lambda_{j, \theta}(t) d N_{j}(t)\right\}
$$

with

$$
\lambda_{j, \theta}(t)=\mu_{j}+\sum_{k=1}^{d} a_{j, k} \int_{(0, t)} \exp \left(-\alpha_{j, k}(t-s)\right) d N_{k}(s)
$$

## Penalized maximum likelihood: a problem

Problem: $\theta \mapsto \lambda_{j, \theta}(t)$ not convex! Indeed

$$
(a, \alpha) \mapsto a h_{\alpha}(t)
$$

never convex when $\alpha \mapsto h_{\alpha}(t)$ is convex


We want convexity for:

- Convergence to a global optimum
- Plethora of optimization algorithms

Generic in the chosen penalization [if proximal operator easy to compute]

A solution: the perspective function trick:

- If $\alpha \mapsto h_{\alpha}(t)$ is convex, then

$$
(a, \alpha) \mapsto a h_{\alpha / a}(t)
$$

is convex

- Reparametrization $\boldsymbol{\beta}=\boldsymbol{A} \circ \boldsymbol{\alpha}$, leading to

$$
\lambda_{j, \theta}(t)=\mu_{j}+\sum_{k=1}^{d} a_{j, k} \int_{(0, t)} \exp \left(-\frac{\beta_{j, k}}{a_{j, k}}(t-s)\right) d N_{k}(s)
$$

with $\theta=(\boldsymbol{\mu}, \boldsymbol{A}, \boldsymbol{\beta})$ for $\boldsymbol{\beta}=\left[\beta_{j, k}\right]_{1 \leq j, k \leq d}$

- With this reparametrization

$$
\theta \mapsto \lambda_{j, \theta}(t)
$$

is convex

The reparametrization $\boldsymbol{\beta}=\boldsymbol{A} \odot \boldsymbol{\alpha}$ leads to

$$
\begin{align*}
\hat{\theta} \in \underset{\theta=(\boldsymbol{\mu}, \boldsymbol{A}, \boldsymbol{\beta})}{\operatorname{argmin}}\left\{-\frac{1}{T} \ell_{T}(\theta)+\right. & \tau\|\boldsymbol{\mu}\|_{1}+\gamma_{1}\|\boldsymbol{A}\|_{1}  \tag{1}\\
& \left.+\gamma_{*}\|\boldsymbol{A}\|_{*}+\frac{\kappa}{2}\|\boldsymbol{\beta}\|_{F}^{2}\right\}
\end{align*}
$$

where

$$
-\frac{1}{T} \ell_{T}(\theta)=\frac{1}{T} \sum_{j=1}^{d}\left\{\int_{0}^{T} \lambda_{j, \theta}(t) d t-\int_{0}^{T} \log \lambda_{j, \theta}(t) d N_{j}(t)\right\}
$$

with

$$
\lambda_{j, \theta}(t)=\mu_{j}+\sum_{k=1}^{d} a_{j, k} \int_{(0, t)} \exp \left(-\frac{\beta_{j, k}}{a_{j, k}}(t-s)\right) d N_{k}(s)
$$

## Convex optimization - numerical aspects

- Can be solved using first-order routines:

Fista [Beck Teboulle (2009)], Prisma [Orabona et al (2012)], GFB [Peyre et al. (2011)], Primal-Dual [Chambolle et al. (2009), Condat et al. (2013)], ADMM [Boyd (2012)], etc...

- Gradient of $-\ell_{T}(\theta)$ using a recursion formula
$\rightarrow$ Naively $O\left(n^{2} d\right)$ with $n=$ number of events (very large) but $O(n d)$ when careful (using recursion formulas)
$\rightarrow$ Parallelized code for this: gradient of each node $j \in\{1, \ldots, d\}$ computed in parallel
- Computation bootleneck: exp and log, accelerated using ugly hacking
- Trace norm penalization, truncated SVD: default's Lanczos's implementation of Python is fast enough for $d \approx 1 K$, use a non-convex factorized formulation $\boldsymbol{A}=\boldsymbol{U} \boldsymbol{V}^{\top}$ for $d \gg 1 K$

Toy example: take matrix $\boldsymbol{A}$ as



We consider a simplified framework

- Fix a set $\left\{h_{j, k}: 1 \leq j, k \leq d\right\}$ and intensities

$$
\lambda_{j, \theta}(t)=\mu_{j}+\int_{(0, t)} \sum_{k=1}^{d} a_{j, k} h_{j, k}(t-s) d N_{k}(s)
$$

where $\theta=[\mu, \boldsymbol{A}]$ with $\mu=\left[\mu_{1}, \ldots, \mu_{d}\right]^{\top}$ and $\boldsymbol{A}=\left[a_{j, k}\right]_{1 \leq j, k \leq d}$

- Instead of - log likelihood, consider least squares

$$
R_{T}(\theta)=\frac{1}{T} \sum_{j=1}^{d}\left\{\int_{0}^{T} \lambda_{j, \theta}(t)^{2} d t-2 \int_{0}^{T} \lambda_{j, \theta}(t) d N_{j}(t)\right\}
$$

Introduce

$$
\hat{\theta} \in \underset{\theta \in \mathbb{R}_{+}^{d} \times \mathbb{R}_{+}^{d \times d}}{\operatorname{argmin}}\left\{R_{T}(\theta)+\operatorname{pen}(\theta)\right\},
$$

with

$$
\operatorname{pen}(\theta)=\|\mu\|_{1, \hat{w}}+\|\boldsymbol{A}\|_{1, \hat{w}}+\hat{w}_{*}\|\boldsymbol{A}\|_{*}
$$

- Penalization tuned by data-driven weights $\hat{w}, \hat{W}$ and $\hat{w}_{*}$
- Comes from sharp controls of the noise terms
- Solves the scaling problem for this model (e.g. feature scaling)
$\ell_{1}$-penalization of $\mu$

$$
\|\mu\|_{1, \hat{w}}=\sum_{j=1}^{d} \hat{w}_{j}\left|\mu_{j}\right|
$$

with

$$
\hat{w}_{j} \approx \sqrt{\frac{(x+\log d) N_{j}([0, T]) / T}{T}}
$$

where $N_{j}([0, T])=\#$ events for node $j$

- Each $\mu_{j}$ penalized by its average events intensity
$\ell_{1}$-penalization of $\boldsymbol{A}$

$$
\|\boldsymbol{A}\|_{1, \hat{\boldsymbol{w}}}=\sum_{1 \leq j, k \leq d} \hat{\boldsymbol{w}}_{j, k}\left|\boldsymbol{A}_{j, k}\right|
$$

with

$$
\hat{\boldsymbol{W}}_{j, k} \approx \sqrt{\frac{(x+\log d) \hat{\boldsymbol{V}}_{j, k}(T)}{T}}
$$

where

$$
\hat{\boldsymbol{V}}_{j, k}(t)=\frac{1}{t} \int_{0}^{t}\left(\int_{(0, s)} h_{j, k}(s-u) d N_{k}(u)\right)^{2} d N_{j}(s)
$$

$=$ variance estimation of the self-excitement for $k \rightarrow j$

Trace-norm penalization of $\boldsymbol{A}$ [difficult]

$$
\hat{w}_{*}\|\boldsymbol{A}\|_{*}=\hat{w}_{*} \sum_{j=1}^{d} \sigma_{j}(\boldsymbol{A})
$$

with

$$
\hat{w}_{*} \approx \sqrt{\frac{(x+\log d)\left(\left\|\hat{\boldsymbol{V}}_{1}(T)\right\|_{\mathrm{op}} \vee\left\|\hat{\boldsymbol{V}}_{2}(T)\right\|_{\mathrm{op}}\right)}{T}}
$$

where $\|\cdot\|_{\text {op }}=$ operator norm

## Solving the "feature scaling" problem

and where $\hat{\boldsymbol{V}}_{1}(t)$ diagonal matrix with entries

$$
\left(\hat{\boldsymbol{V}}_{1}(t)\right)_{j, j}=\frac{1}{t} \int_{0}^{t}\|\boldsymbol{H}(s)\|_{2, \infty}^{2} d N_{j}(s)
$$

$\hat{\boldsymbol{V}}_{2}(t)$ matrix with entries

$$
\left(\hat{\boldsymbol{V}}_{2}(t)\right)_{j, k}=\frac{1}{t} \int_{0}^{t}\|\boldsymbol{H}(s)\|_{2, \infty}^{2} \sum_{l=1}^{d} \frac{H_{j, l}(s) H_{k, l}(s)}{\left\|\boldsymbol{H}_{l, \bullet}(s)\right\|_{2}^{2}} d N_{l}(s)
$$

with $\|\cdot\|_{2, \infty}=$ maximum $\ell_{2}$ row norm and $\boldsymbol{H}(t)$ matrix with entries

$$
\boldsymbol{H}_{j, k}(t)=\int_{(0, t)} h_{j, k}(t-s) d N_{k}(s)
$$

- $\hat{\boldsymbol{V}}_{j, k}(t),\left\|\hat{\boldsymbol{V}}_{1}(t)\right\|_{\text {op }}$ and $\left\|\hat{\boldsymbol{V}}_{2}(t)\right\|_{\text {op }}$ are estimations (based on optional variation) of non-observable variance terms
- It comes from new Bernstein's concentration inequalities, used on the noise term
- We develop a new probabilistic tool: non-commutative concentration inequality for random matrix martingales in continuous time (theory given by Tropp (2011) applies to discrete time only, and depend on unobserved variance terms)

These tools give a sharp data-driven tuning of the penalizations, solving the scaling problem

Define

$$
\left\langle\lambda_{1}, \lambda_{2}\right\rangle_{T}=\frac{1}{T} \sum_{j=1}^{d} \int_{0}^{T} \lambda_{1, j}(t) \lambda_{2, j}(t) d t
$$

and $\|\lambda\|_{T}^{2}=\langle\lambda, \lambda\rangle_{T}$

- We use a standard assumption to obtain fast rates for the Lasso: the RE (Restricted Eigenvalue) Assumption [Bickel et al. (2009), Koltchinkii (2011), ...]


## Theorem 1

We have

$$
\begin{aligned}
&\left\|\lambda_{\hat{\theta}}-\lambda_{0}\right\|_{T}^{2} \leq \inf _{\theta}\left\{\left\|\lambda_{\theta}-\lambda_{0}\right\|_{T}^{2}+\kappa(\theta)^{2}\left(\frac{5}{4}\left\|(\hat{w})_{\operatorname{supp}(\mu)}\right\|_{2}^{2}\right.\right. \\
&\left.\left.+\frac{9}{8}\left\|(\hat{\boldsymbol{W}})_{\operatorname{supp}(\boldsymbol{A})}\right\|_{F}^{2}+\frac{9}{8} \hat{w}_{*}^{2} \operatorname{rank}(\boldsymbol{A})\right)\right\}
\end{aligned}
$$

with a probability larger than $1-146 e^{-x}$

- $\kappa(\theta)$ : RE constant
- Sharp: leading constant 1


## Some theory: sharp oracle inequalities

Take-home message: $\hat{\theta}$ achieves an optimal tradeoff between approximation and complexity given by

$$
\begin{aligned}
& \frac{\|\mu\|_{0}(x+\log d)}{T} \max _{j} N_{j}([0, T]) / T \\
& \quad+\frac{\|\boldsymbol{A}\|_{0}(x+\log d)}{T} \max _{j, k} \hat{\boldsymbol{V}}_{j, k}(T) \\
& \quad+\frac{\operatorname{rank}(A)(x+\log d)}{T}\left\|\hat{\boldsymbol{V}}_{1}(T)\right\|_{\mathrm{op}} \vee\left\|\hat{\boldsymbol{V}}_{2}(T)\right\|_{\mathrm{op}} .
\end{aligned}
$$

- Complexity measured by sparsity and rank
- Convergence has shape $(\log d) / T$, where $T=$ length of the observation interval
- Terms balanced by empirical variance terms
- New Bernstein's empirical concentration inequality for continuous-time matrix martingale
- Consider the random matrix $\boldsymbol{Z}(t)$ with entries

$$
\boldsymbol{Z}_{j, k}(t)=\int_{0}^{t} \int_{(0, s)} h_{j, k}(s-u) d N_{k}(u) d M_{j}(s)
$$

where $M_{j}(t)=N_{j}(t)-\int_{0}^{t} \lambda_{j}(s) d s$ are martingales obtained by compensation

- This is the noise term in our problem


## A new concentration inequality

A classical concentration inequality for $\boldsymbol{Z}_{j, k}$
[Lipster Shiryayev 1986] is

$$
\frac{1}{t}(\boldsymbol{Z}(t))_{j, k} \leq \sqrt{\frac{2 v x}{t}}+\frac{b x}{3 t}
$$

for any $x>0$, with a probability $\geq 1-e^{-x}$ whenever

$$
\frac{1}{t}\left\langle\boldsymbol{Z}_{j, k}\right\rangle_{t}=\frac{1}{t} \int_{0}^{t}\left(\int_{(0, s)} h_{j, k}(s-u) d N_{k}(u)\right)^{2} \lambda_{j}(s) d s \leq v
$$

and

$$
\sup _{s \in[0, t]} \int_{(0, s)} h_{j, k}(s-u) d N_{k}(u) \leq b
$$

- Predictable variation $\left\langle\boldsymbol{Z}_{j, k}\right\rangle_{t}$ depends on non-observed $\lambda_{j}$ : this concentration is useless for statistics
- Need an empirical Bernstein's inequality, with a variance term using the optional variation

$$
\frac{1}{t}\left[\boldsymbol{Z}_{j, k}\right]_{t}=\frac{1}{t} \int_{0}^{t}\left(\int_{(0, s)} h_{j, k}(s-u) d N_{k}(u)\right)^{2} d N_{j}(s)
$$

- We need also to remove the event $\left\{\left\langle\boldsymbol{Z}_{j, k}\right\rangle_{t} \leq t v\right\}$ from this inequality
We provide:
- A control of all the entries $\boldsymbol{Z}_{j, k}$ of $\boldsymbol{Z}$
- A control of $\left\|\boldsymbol{Z}_{t}\right\|_{\mathrm{op}}$


## A new concentration inequality

## Theorem 2

We have

$$
\begin{aligned}
\frac{1}{t}\left|Z_{j, k}(t)\right| \leq & 2 \sqrt{2} \sqrt{\frac{\left(x+2 \log d+\hat{\boldsymbol{L}}_{j, k}(t)\right) \hat{\boldsymbol{V}}_{j, k}(t)}{t}} \\
& +9.31 \frac{\left(x+2 \log d+\hat{\boldsymbol{L}}_{j, k}(t)\right) \boldsymbol{B}_{j, k}(t)}{t}
\end{aligned}
$$

for any $1 \leq j, k \leq d$, with a probability larger than $1-30.55 e^{-x}$.

- Based on a previous result by G. and Guilloux (2011), see also Hansen et al (2012)
- Reminiscent of previous works by Audibert (2008)


## A new concentration inequality

## Theorem 3

For any $x>0$, we have

$$
\begin{aligned}
& \frac{\|\boldsymbol{Z}(t)\|_{\mathrm{op}}}{t} \\
& \quad \leq 4 \sqrt{\frac{\left(x+\log d+\hat{\ell}_{x}(t)\right)\left\|\hat{\boldsymbol{V}}_{1}(t)\right\|_{\mathrm{op}} \vee\left\|\hat{\boldsymbol{V}}_{2}(t)\right\|_{\mathrm{op}}}{t}} \\
& \quad+\frac{\left(x+\log d+\hat{\ell}_{x}(t)\right)\left(10.34+2.65 \sup _{t \in[0, T]}\|\boldsymbol{H}(t)\|_{2, \infty}\right)}{t}
\end{aligned}
$$

with a probability larger than $1-84.9 e^{-x}$

- First non-commutative Bernstein's inequality for countinuous time martingales
- Can be extended to a wider class of martingales
- Extension of [Tropp (2012)] results


## Consequence: a sharp scaling of penalizations



Ground Truth


NoPen


L1

## Consequence: a sharp scaling of penalizations

## L1 vs wL1




## Consequence: a sharp scaling of penalizations

## Nuclear vs wNuclear




## Consequence: a sharp scaling of penalizations



Error for L1 and wL1


Error for L1Nuclear and wL1Nuclear


AUC for L1 and wL1

## Consequence: a sharp scaling of penalizations



Error for L1 and wL1


Error for L1Nuclear and wL1Nuclear


AUC for L1 and wL1

- Reparametrization of the problem
- Theoretical analysis gives insight to choose the correct scaling of the penalizations
- First oracle inequality for this problem
- This required new probabilistic tools for matrix martingales in continuous time
- Larger scale: factorized form $\boldsymbol{A}=\boldsymbol{U} \boldsymbol{V}^{\top}$
- Incorporation of features (text, time-varying graph-features, etc.)
- Time varying baseline $\mu(t)$ for non-stationarity


## Thank you！


[^0]:    ${ }^{1}$ CMAP - Ecole Polytechnique

