# Machine learning and time: "time accounting" learning

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- Finite network with nodes  $\{1, \ldots, d\}$ : users of a social network, of an e-commerce platform, etc.
- For each node  $j \in \{1, ..., d\}$  we observe the timestamps  $\{t_{j,1}, t_{j,2}, ...\}$  of nodes' actions
- Goal: recover **levels of interactions** between users based on the timestamps patterns

## Introduction

## From



Quantify interactions between users





- Do inference directly from actions of users
- Understand the community structure of users underlying the actions
- Exploit the hidden lower-dimensional structure of the network for inference/prediction

- Counting process  $N_j(t) = \sum_{i\geq 1} \mathbf{1}_{t_{j,i}\leq t}$
- Data: a *d*-dimensional counting process  $N = [N_1, \ldots, N_d]^\top$
- *d* is large
- Observed on [0, T]. "Asymptotics" in  $T \to +\infty$
- $N_j$  has intensity  $\lambda_j$ , namely

 $\mathbb{P}(j ext{ does something at time } t ext{ knowning the past}) = \mathbb{P}(N_j ext{ has a jump in } [t, t + dt] | \mathcal{F}_t) = \lambda_j(t)dt$ 

for  $j = 1, \ldots, d$  where  $\mathcal{F}_t$  some filtration

## Model: Multivariate Hawkes Process (MHP)

• MHP assumes an autoregressive structure on the intensities:

$$\lambda_{j}(t) = \mu_{j}(t) + \int_{(0,t)} \sum_{k=1}^{d} \varphi_{j,k}(t-s) dN_{k}(s),$$

- $\mu_j(t) \ge 0$  baseline intensity of the *j*-th coordinate
- $\varphi_{j,k} : \mathbb{R}^+ \to \mathbb{R}^+$  self-exciting component: influence of  $k \to j$
- Write this in matrix form

$$\lambda(t) = \mu + \int_{(0,t)} \varphi(t-s) dN(s),$$

with  $\boldsymbol{\mu} = [\mu_1, \dots, \mu_d]^\top$  and  $\boldsymbol{\varphi}(t) = [\varphi_{\boldsymbol{j}, \boldsymbol{k}}(t)]_{1 \leq \boldsymbol{j}, \boldsymbol{k} \leq d}$ .

• Notation:

$$\int_{(0,t)} \varphi(t-s) dN_j(s) = \sum_{i:t_{j,i} < t} \varphi(t-t_{j,i})$$

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Introduced by Hawkes in 1971

- Earthquakes and geophysics : Kagan and Knopoff (1981), Zhuang, Harte, Werner, Hainzl and Zhou (2012)
- **Genomics** : Reynaud-Bouret and Schbath (2010)
- **High-frequency Finance** : Bacry Delattre Hoffmann and Muzy (2013)
- **Terrorist activity** : Porter and White (2012)
- **Neurobiology** : Hansen, Reynaud-Bouret and Rivoirard (2012)
- Social networks : Carne and Sornette (2008), Simma and Jordan (2010), Zhou Song and Zha (2013)
- And even FPGA-based implementation : Guo and Luk (2013)

## A brief history of MHP



Home / Bitcoin 201 / Analyzing Trade Clustering To Predict Price Movement In Bitcoin Trading



## Analyzing Trade Clustering To Predict Price Movement In Bitcoin Trading

Sep 19, 2013 Posted By Jonathan Heusser In Bitcoin 201, Economics, Featured, News, Trading Tagged Analysis, Bitcoin Trading,

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Hawkes Process, Jonathan Heusser, London, Price, Trading

#### **Parametric estimation** (Maximum likelihood)

- First work : Ogata 78
- Simma and Jordan (2010), Zhou Song and Zha (2013)
  - $\rightarrow$  Expected Maximization (EM) algorithms, with priors

#### Non parametric estimation

- Marsan Lengliné (2008), generalized by Lewis, Mohler (2010)
  - $\rightarrow$  EM for penalized likelihood function
  - $\rightarrow$  Monovariate Hawkes processes, Small amount of data, No theoretical results
- Reynaud-Bouret and Schbath (2010)
  - $\rightarrow$  Developed for small amount of data (Sparse penalization)
- Bacry and Muzy (2014)
  - $\rightarrow$  Larger amount of data

Dimension *d* is large:

- ullet Need a simple parametric model on  $\mu$  and  $\varphi$
- We want a **convex** optimization problem with smooth loss
- We want to encode some prior assumptions by penalizing this loss

#### Simple parametrization:

- Constant baselines  $\mu_j(\cdot) \equiv \mu_j$
- Take

$$\varphi_{j,k}(t) = a_{j,k} e^{-\alpha_{j,k}t}$$

•  $a_{j,k}$  = level of interaction between nodes j and k

•  $\alpha_{j,k} =$  lifetime of instantaneous excitation of node j by node kThe matrix

$$oldsymbol{A} = [a_{j,k}]_{1 \leq j,k \leq d}$$

is understood has a **weighted adjacency matrix** of mutual excitement of th nodes  $\{1, \ldots, d\}$ 

• **A** is non-symmetric

We end up with intensities

$$\lambda_{j,\theta}(t) = \mu_j + \int_{(0,t)} \sum_{k=1}^d a_{j,k} e^{-\alpha_{j,k}(t-s)} dN_k(s)$$

for  $j \in \{1, \ldots, d\}$  where

$$heta = [\mu, oldsymbol{A}, oldsymbol{lpha}]$$

with

- baselines  $\mu = [\mu_1, \dots, \mu_d]^\top \in \mathbb{R}^d_+$
- adjacencies  $\mathbf{A} = [a_{j,k}]_{1 \leq j,k \leq d} \in \mathbb{R}^{d \times d}_+$
- decays  $\boldsymbol{\alpha} = [\alpha_{j,k}]_{1 \leq j,k \leq d} \in \mathbb{R}^{d \times d}_+$

## For d = 1, intensity $\lambda_{\theta}$ looks like this:



 **Goodness-of-fit** =  $-\log$ -likelihood is given by:

$$-\ell_{T}(\theta) = \sum_{j=1}^{d} \left\{ \int_{0}^{T} (\lambda_{j,\theta}(t) - 1) dt - \int_{0}^{T} \log \lambda_{j,\theta}(t) dN_{j}(t) \right\}$$

with

$$\lambda_{j,\theta}(t) = \mu_j + \sum_{k=1}^d a_{j,k} \int_{(0,t)} \exp\left(-\alpha_{j,k}(t-s)\right) dN_k(s)$$

where  $heta=(oldsymbol{\mu},oldsymbol{A},oldsymbol{lpha})$  with  $oldsymbol{\mu}=[\mu_j]$ ,  $oldsymbol{A}=[A_{j,k}]$ ,  $oldsymbol{lpha}=[lpha_{j,k}]$ 

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## Prior encoding by penalization

## **Prior** assumptions

• Some users are basically inactive and react only if stimulated:

 $\mu$  is sparse

• Everybody does not interact with everybody:

A is sparse

• Interactions have community structure, possibly overlapping, a small number of factors explain interactions:



• Decays  $\alpha$  not sparse, but  $\alpha_{j,k}$  should be regularized proportionally to  $a_{j,k}$ 

**A** is low-rank

**Standard convex relaxations** [Tibshirani 01, ..., Srebro et al. 05, Bach 08, Candès & Recht 08, ...]

• Tightest convex relaxation of  $\|\mathbf{A}\|_0 = \sum_{j,k} \mathbf{1}_{\mathbf{A}_{j,k}>0}$  is  $\ell_1$ -norm:

$$\|oldsymbol{A}\|_1 = \sum_{j,k} |oldsymbol{A}_{j,k}|$$

• Tightest convex relaxation of rank is trace-norm:

$$\|A\|_* = \sum_j \sigma_j(A) = \|\sigma(A)\|_1$$

where  $\sigma_1(A) \geq \cdots \geq \sigma_d(A)$  singular values of **A** 

#### So, we use the following penalizations

- Use  $\ell_1$  penalization on  $oldsymbol{\mu}$
- Use  $\ell_1$  penalization on  $\boldsymbol{A}$
- Use trace-norm penalization on A
- Use  $\ell_2^2$  penalization on  $\alpha$ , weighted by **A**

[but other choices might be interesting...]

NB1: to induce **sparsity AND low-rank** on **A**, we use the mixed penalization

$$oldsymbol{A}\mapsto w_*\|oldsymbol{A}\|_*+w_1\|oldsymbol{A}\|_1$$

NB2: recent works by Richard et al (2013, 2014): better way to induce sparsity and low-rank than the sum, but not-scalable / non-convex

#### Sparse and low-rank matrices



 $\{ \boldsymbol{A} : \|\boldsymbol{A}\|_* \leq 1 \} \qquad \qquad \{ \boldsymbol{A} : \|\boldsymbol{A}\|_1 \leq 1 \} \qquad \{ \boldsymbol{A} : \|\boldsymbol{A}\|_1 + \|\boldsymbol{A}\|_* \leq 1 \}$ 

The balls are computed on the set of  $2 \times 2$  symmetric matrices, which is identified with  $\mathbb{R}^3$ .

[show video]

Finally, consider

$$\hat{\theta} \in \underset{\theta=(\boldsymbol{\mu},\boldsymbol{A},\boldsymbol{\alpha})}{\operatorname{argmin}} \left\{ -\frac{1}{T} \ell_{T}(\theta) + \tau \|\boldsymbol{\mu}\|_{1} + \gamma_{1} \|\boldsymbol{A}\|_{1} \right\}$$

where we recall

$$-\frac{1}{T}\ell_{T}(\theta) = \frac{1}{T}\sum_{j=1}^{d} \left\{ \int_{0}^{T} \lambda_{j,\theta}(t)dt - \int_{0}^{T} \log \lambda_{j,\theta}(t)dN_{j}(t) \right\}$$

with

$$\lambda_{j,\theta}(t) = \mu_j + \sum_{k=1}^d a_{j,k} \int_{(0,t)} \exp\left(-\alpha_{j,k}(t-s)\right) dN_k(s)$$

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Problem:  $\theta \mapsto \lambda_{j,\theta}(t)$  not convex! Indeed

 $(a, \alpha) \mapsto ah_{\alpha}(t)$ 

**never convex** when  $\alpha \mapsto h_{\alpha}(t)$  is convex

We want convexity for:

- Convergence to a global optimum
- Plethora of optimization algorithms

Generic in the chosen penalization [if proximal operator easy to compute]



A solution: the **perspective function** trick:

• If  $\alpha \mapsto h_{\alpha}(t)$  is convex, then

 $(a, \alpha) \mapsto ah_{\alpha/a}(t)$ 

#### is convex

• Reparametrization  $\boldsymbol{\beta} = \boldsymbol{A} \circ \boldsymbol{\alpha}$ , leading to

$$\lambda_{j,\theta}(t) = \mu_j + \sum_{k=1}^d a_{j,k} \int_{(0,t)} \exp\left(-\frac{\beta_{j,k}}{a_{j,k}}(t-s)\right) dN_k(s)$$

with  $\theta = (\mu, \boldsymbol{A}, \boldsymbol{\beta})$  for  $\boldsymbol{\beta} = [\beta_{j,k}]_{1 \leq j,k \leq d}$ 

• With this reparametrization

$$heta\mapsto\lambda_{j, heta}(t)$$

is convex

## Penalized maximum likelihood: reparametrization

The reparametrization  $\boldsymbol{\beta} = \boldsymbol{A} \odot \boldsymbol{\alpha}$  leads to

$$\hat{\theta} \in \underset{\theta=(\mu,\boldsymbol{A},\boldsymbol{\beta})}{\operatorname{argmin}} \left\{ -\frac{1}{T} \ell_{T}(\theta) + \tau \|\boldsymbol{\mu}\|_{1} + \gamma_{1} \|\boldsymbol{A}\|_{1} + \gamma_{*} \|\boldsymbol{A}\|_{*} + \frac{\kappa}{2} \|\boldsymbol{\beta}\|_{F}^{2} \right\}$$
(1)

where

$$-\frac{1}{T}\ell_{T}(\theta) = \frac{1}{T}\sum_{j=1}^{d} \left\{ \int_{0}^{T} \lambda_{j,\theta}(t)dt - \int_{0}^{T} \log \lambda_{j,\theta}(t)dN_{j}(t) \right\}$$

with

$$\lambda_{j,\theta}(t) = \mu_j + \sum_{k=1}^d a_{j,k} \int_{(0,t)} \exp\left(-\frac{\beta_{j,k}}{a_{j,k}}(t-s)\right) dN_k(s)$$

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• Can be solved using first-order routines:

Fista [Beck Teboulle (2009)], Prisma [Orabona et al (2012)], GFB [Peyre et al. (2011)], Primal-Dual [Chambolle et al. (2009), Condat et al. (2013)], ADMM [Boyd (2012)], etc...

• Gradient of  $-\ell_T(\theta)$  using a recursion formula

 $\rightarrow$  Naively  $O(n^2d)$  with n = number of events (very large) but O(nd) when careful (using recursion formulas)

 $\rightarrow$  Parallelized code for this: gradient of each node  $j \in \{1, \dots, d\}$  computed **in parallel** 

- Computation bootleneck: exp and log, accelerated using ugly hacking
- Trace norm penalization, truncated SVD: default's Lanczos's implementation of Python is fast enough for  $d \approx 1K$ , use a non-convex factorized formulation  $\mathbf{A} = \mathbf{U}\mathbf{V}^{\top}$  for  $d \gg 1K$

Toy example: take matrix **A** as



#### Numerical experiment: dimension 100, 20100 parameters



We consider a simplified framework

• Fix a set  $\{h_{j,k} : 1 \leq j, k \leq d\}$  and intensities

$$\lambda_{j,\theta}(t) = \mu_j + \int_{(0,t)} \sum_{k=1}^d a_{j,k} h_{j,k}(t-s) dN_k(s),$$

where  $\theta = [\mu, \mathbf{A}]$  with  $\mu = [\mu_1, \dots, \mu_d]^\top$  and  $\mathbf{A} = [a_{j,k}]_{1 \leq j,k \leq d}$ 

• Instead of - log likelihood, consider least squares

$$R_{T}(\theta) = \frac{1}{T} \sum_{j=1}^{d} \left\{ \int_{0}^{T} \lambda_{j,\theta}(t)^{2} dt - 2 \int_{0}^{T} \lambda_{j,\theta}(t) dN_{j}(t) \right\}$$

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#### Introduce

$$\hat{\theta} \in \underset{\theta \in \mathbb{R}^{d}_{+} \times \mathbb{R}^{d \times d}_{+}}{\operatorname{argmin}} \{ R_{T}(\theta) + \operatorname{pen}(\theta) \},$$

with

$$\mathsf{pen}(\theta) = \|\mu\|_{1,\hat{w}} + \|\boldsymbol{A}\|_{1,\hat{\boldsymbol{W}}} + \hat{w}_*\|\boldsymbol{A}\|_*$$

- Penalization tuned by data-driven weights  $\hat{w}$ ,  $\hat{W}$  and  $\hat{w}_*$
- Comes from sharp controls of the noise terms
- Solves the scaling problem for this model (e.g. feature scaling)

## $\ell_1\text{-penalization}$ of $\mu$

$$\|\mu\|_{1,\hat{w}} = \sum_{j=1}^d \hat{w}_j |\mu_j|$$

#### with

$$\hat{w}_j pprox \sqrt{rac{(x+\log d)N_j([0, T])/T}{T}}$$

where  $N_i([0, T]) = \#$  events for node j

• Each  $\mu_j$  penalized by its average events intensity

 $\ell_1\text{-penalization}$  of  $\boldsymbol{A}$ 

$$\| \boldsymbol{A} \|_{1, \hat{\boldsymbol{W}}} = \sum_{1 \leq j,k \leq d} \hat{\boldsymbol{W}}_{j,k} | \boldsymbol{A}_{j,k} |$$

with

$$\hat{\boldsymbol{W}}_{j,k} \approx \sqrt{rac{(x+\log d)\hat{\boldsymbol{V}}_{j,k}(T)}{T}}$$

where

$$\hat{V}_{j,k}(t) = \frac{1}{t} \int_0^t \left( \int_{(0,s)} h_{j,k}(s-u) dN_k(u) \right)^2 dN_j(s)$$

= variance estimation of the self-excitement for  $k \rightarrow j$ 

## Trace-norm penalization of **A** [difficult]

$$\hat{w}_* \| \boldsymbol{A} \|_* = \hat{w}_* \sum_{j=1}^d \sigma_j(\boldsymbol{A})$$

#### with

$$\hat{w}_* \approx \sqrt{\frac{(x + \log d)(\|\hat{\boldsymbol{V}}_1(T)\|_{\mathrm{op}} \vee \|\hat{\boldsymbol{V}}_2(T)\|_{\mathrm{op}})}{T}}$$

where  $\|\cdot\|_{\mathrm{op}} = \mathsf{operator} \mathsf{ norm}$ 

and where  $\hat{\boldsymbol{V}}_1(t)$  diagonal matrix with entries

$$(\hat{\boldsymbol{V}}_{1}(t))_{j,j} = \frac{1}{t} \int_{0}^{t} \|\boldsymbol{H}(s)\|_{2,\infty}^{2} dN_{j}(s),$$

 $\hat{\boldsymbol{V}}_2(t)$  matrix with entries

$$(\hat{\boldsymbol{V}}_{2}(t))_{j,k} = rac{1}{t} \int_{0}^{t} \|\boldsymbol{H}(s)\|_{2,\infty}^{2} \sum_{l=1}^{d} rac{H_{j,l}(s)H_{k,l}(s)}{\|\boldsymbol{H}_{l,\bullet}(s)\|_{2}^{2}} dN_{l}(s),$$

with  $\|\cdot\|_{2,\infty} = \max \lim_{t \to \infty} \ell_2$  row norm and H(t) matrix with entries

$$\boldsymbol{H}_{j,k}(t) = \int_{(0,t)} h_{j,k}(t-s) dN_k(s)$$

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- $\hat{V}_{j,k}(t)$ ,  $\|\hat{V}_1(t)\|_{op}$  and  $\|\hat{V}_2(t)\|_{op}$  are estimations (based on optional variation) of non-observable variance terms
- It comes from new Bernstein's concentration inequalities, used on the noise term
- We develop a new probabilistic tool: non-commutative concentration inequality for random matrix martingales in continuous time (theory given by Tropp (2011) applies to discrete time only, and depend on unobserved variance terms)

These tools give a **sharp data-driven tuning** of the penalizations, solving the scaling problem

#### Define

$$\langle \lambda_1, \lambda_2 \rangle_T = \frac{1}{T} \sum_{j=1}^d \int_0^T \lambda_{1,j}(t) \lambda_{2,j}(t) dt$$

and  $\|\lambda\|_T^2 = \langle \lambda, \lambda \rangle_T$ 

 We use a standard assumption to obtain fast rates for the Lasso: the RE (Restricted Eigenvalue) Assumption [Bickel et al. (2009), Koltchinkii (2011), ...]

#### Theorem 1

#### We have

$$\begin{split} \|\lambda_{\hat{\theta}} - \lambda_0\|_{\mathcal{T}}^2 &\leq \inf_{\theta} \left\{ \|\lambda_{\theta} - \lambda_0\|_{\mathcal{T}}^2 + \kappa(\theta)^2 \Big(\frac{5}{4} \|(\hat{w})_{\mathsf{supp}(\mu)}\|_2^2 \\ &+ \frac{9}{8} \|(\hat{\boldsymbol{W}})_{\mathsf{supp}(\boldsymbol{A})}\|_F^2 + \frac{9}{8} \hat{w}_*^2 \operatorname{rank}(\boldsymbol{A})\Big) \right\} \end{split}$$

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with a probability larger than  $1 - 146e^{-x}$ 

- $\kappa(\theta)$ : RE constant
- Sharp: leading constant 1

Take-home message:  $\hat{\theta}$  achieves an optimal tradeoff between approximation and complexity given by

$$\frac{\|\mu\|_0(x+\log d)}{T} \max_j N_j([0,T])/T$$

$$+ \frac{\|\boldsymbol{A}\|_0(x+\log d)}{T} \max_{j,k} \hat{\boldsymbol{V}}_{j,k}(T)$$

$$+ \frac{\operatorname{rank}(A)(x+\log d)}{T} \|\hat{\boldsymbol{V}}_1(T)\|_{\operatorname{op}} \vee \|\hat{\boldsymbol{V}}_2(T)\|_{\operatorname{op}}.$$

- Complexity measured by sparsity and rank
- Convergence has shape  $(\log d)/T$ , where T = length of the observation interval
- Terms balanced by empirical variance terms

- New Bernstein's empirical concentration inequality for continuous-time matrix martingale
- Consider the random matrix Z(t) with entries

$$Z_{j,k}(t) = \int_0^t \int_{(0,s)} h_{j,k}(s-u) dN_k(u) dM_j(s)$$

where  $M_j(t) = N_j(t) - \int_0^t \lambda_j(s) ds$  are martingales obtained by compensation

• This is the noise term in our problem

A classical concentration inequality for  $Z_{j,k}$ [Lipster Shiryayev 1986] is

$$rac{1}{t}(oldsymbol{Z}(t))_{j,k} \leq \sqrt{rac{2vx}{t}} + rac{bx}{3t}$$

for any x > 0, with a probability  $\geq 1 - e^{-x}$  whenever

$$\frac{1}{t}\langle \boldsymbol{Z}_{j,k}\rangle_t = \frac{1}{t}\int_0^t \Big(\int_{(0,s)} h_{j,k}(s-u)dN_k(u)\Big)^2\lambda_j(s)ds \leq v$$

and

$$\sup_{s\in[0,t]}\int_{(0,s)}h_{j,k}(s-u)dN_k(u)\leq b$$

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## A new concentration inequality

- Predictable variation  $\langle \mathbf{Z}_{j,k} \rangle_t$  depends on non-observed  $\lambda_j$ : this concentration is **useless** for statistics
- Need an **empirical Bernstein's inequality**, with a variance term using the **optional variation**

$$\frac{1}{t}[\boldsymbol{Z}_{j,k}]_t = \frac{1}{t} \int_0^t \Big( \int_{(0,s)} h_{j,k}(s-u) dN_k(u) \Big)^2 dN_j(s)$$

• We need also to remove the event  $\{\langle \mathbf{Z}_{j,k} \rangle_t \leq tv\}$  from this inequality

We provide:

- A control of all the entries  $Z_{j,k}$  of Z
- A control of  $\|\boldsymbol{Z}_t\|_{\mathrm{op}}$

#### Theorem 2

We have

$$\begin{aligned} \frac{1}{t} |\boldsymbol{Z}_{j,k}(t)| &\leq 2\sqrt{2} \sqrt{\frac{(x+2\log d + \hat{\boldsymbol{L}}_{j,k}(t))\hat{\boldsymbol{V}}_{j,k}(t)}{t}} \\ &+ 9.31 \frac{(x+2\log d + \hat{\boldsymbol{L}}_{j,k}(t))\boldsymbol{B}_{j,k}(t)}{t} \end{aligned}$$

for any  $1 \le j, k \le d$ , with a probability larger than  $1 - 30.55e^{-x}$ .

- Based on a previous result by G. and Guilloux (2011), see also Hansen et al (2012)
- Reminiscent of previous works by Audibert (2008)

#### Theorem 3

For any x > 0, we have

$$\frac{\|\boldsymbol{Z}(t)\|_{\text{op}}}{t} \leq 4\sqrt{\frac{(x+\log d+\hat{\ell}_{x}(t))\|\hat{\boldsymbol{V}}_{1}(t)\|_{\text{op}} \vee \|\hat{\boldsymbol{V}}_{2}(t)\|_{\text{op}}}{t}} + \frac{(x+\log d+\hat{\ell}_{x}(t))(10.34+2.65\sup_{t\in[0,T]}\|\boldsymbol{H}(t)\|_{2,\infty})}{t}$$

with a probability larger than  $1 - 84.9e^{-x}$ 

- First non-commutative Bernstein's inequality for countinuous time martingales
- Can be extended to a wider class of martingales
- Extension of [Tropp (2012)] results

## Consequence: a sharp scaling of penalizations



# L1 vs wL1





wL1

## Nuclear vs wNuclear





wL1Nuclear

#### Consequence: a sharp scaling of penalizations







wL1

Error for L1 and Error for L1Nuclear AUC for L1 and and wL1Nuclear

wL1

#### Consequence: a sharp scaling of penalizations







wL1

Error for L1 and Error for L1Nuclear AUC for L1 and and wL1Nuclear

wL1

- Reparametrization of the problem
- Theoretical analysis gives insight to choose the correct scaling of the penalizations
- First oracle inequality for this problem
- This required new probabilistic tools for matrix martingales in continuous time

- Larger scale: factorized form  $\mathbf{A} = \mathbf{U}\mathbf{V}^{\top}$
- Incorporation of features (text, time-varying graph-features, etc.)
- Time varying baseline  $\mu(t)$  for non-stationarity

# Thank you!

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