# Machine Learning - Course $\mathrm{n}^{\circ} 2$ 

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## Unsupervised Learning: Clustering

## Unsupervised learning

- Marketing: finding groups of customers with similar behavior given a large database of customer data containing their properties and past buying records
- Biology: classification of plants and animals given their features
- Insurance: identifying groups of motor insurance policy holders with a high average claim cost, identifying frauds
- City-planning: identifying groups of houses according to their house type, value and geographical location
- Internet: document classification, clustering weblog data to discover groups of similar access patterns


## Two main directions

- Data: base of customer data containing their properties and past buying records
- Goal: Use the customers similarities to find groups


## Two directions:

- Clustering: propose an explicit grouping of the customers
- Visualization: propose a representation of the customers so that the groups are visibles



## Supervised learning reminder

- Training data $D_{n}=\left[\left(x_{1}, y_{1}\right), \ldots,\left(x_{n}, y_{n}\right)\right]$
- $\left(x_{i}, y_{i}\right)$ i.i.d $\mathbb{P}$ on $\mathcal{X} \times \mathcal{Y}$
- Construct a predictor $\hat{f}: \mathcal{X} \rightarrow \mathcal{Y}$ using $D_{n}$
- Loss $\ell(y, f(x))$ measures how well $f(x)$ predicts $y$ well
- Aim is to minimize the generalization error

$$
\mathcal{E}_{X, Y}\left(\ell(Y, \hat{f}(X)) \mid D_{n}\right)=\int \ell(y, \hat{f}(x)) d \mathbb{P}(x, y)
$$

The goal is clear

- Predict $y$ based on feature $x$

Heard on the street

- Supervised learning is solved. Unsupervised learning isn't


## Unsupervised learning

- Training data $D_{n}=\left[x_{1}, \ldots, x_{n}\right]$
- Loss: Not Clear
- Aim: Not Clear

The goal is unclear.

## Classical tasks

- Clustering: construct groups of data in homogeneous classes
- Dimension reduction: construct a map of the data in a low-dimension space without distorting it too much


## Motivations

- Interpretation of the groups
- Use of the groups in further processing


## Clustering

## Clustering

- Training data $D_{n}=\left\{x_{1}, \ldots, x_{n}\right\}$ with $x_{i} \in \mathbb{R}^{d}$
- Recover Latent groups
- Construct $f: \mathbb{R}^{d} \rightarrow\{1, \ldots, K\}$ which affects cluster number to $x_{i}$

$$
f: x_{i} \mapsto k_{i}
$$

- No ground truth for $k_{i}$


## Warning

- Choice of $K$ is hard

Roughly two approaches

- Partition-based
- Model-based


## K-means

## K-means problem

## K-means

- Fix $K \geq 2$, $n$ data points $x_{i} \in \mathbb{R}^{d}$
- Find centroids $c_{1}, \ldots c_{K}$ that minimizes the quantification error

$$
\sum_{i=1}^{n} \min _{k=1, \ldots, K}\left\|x_{i}-c_{k}\right\|_{2}^{2}
$$

- Impossible to find the exact solution (NP Complete)


## K-means algorithm

Lloyd (1981) proposes a way of finding local solutions
K-means algorithm

- Choose at random $K$ centroids $\left\{c_{1}, \ldots, c_{K}\right\}$
- For each $k \in\{1, \ldots, K\}$, find the set $C_{k}$ of points that are closer to $c_{k}$ than any $c_{k^{\prime}}$ for $k^{\prime} \neq k$
- Update the centroids:

$$
c_{k}=\frac{1}{\left|C_{k}\right|} \sum_{i \in C_{k}} x_{i}
$$

- Repeat the two previous steps until the sets $C_{k}$ don't change


## K-means algorithm

Lloyd (1981) proposes a way of finding local solutions

## K-means algorithm

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Remark: K-means computes a Voronoi partitioning, it implicitly assumes convex clusters, that are uniquely defined by their centroids.

## Voronoi partitioning


(U) (m)

## Bryant park


(U) 些

## K-means



## K-means



## Initialization problem

K-means very sensitive to the choice of initial points




## An easy solution:

- Pick a first point at random
- Choose the next initial point the farthest from the previous ones

Farthest point initialization


$$
\begin{aligned}
& \text { OOO } \\
& 0 \text { OOO } \\
& 0000
\end{aligned}
$$

$$
0{ }_{0}^{0}{ }_{0}^{0}{ }_{0}^{0} 0_{0}^{0} 0_{0}^{0} 0
$$



Farthest point initialization


$$
\begin{aligned}
& 0_{0}^{0} 00 \\
& 0000 \\
& 0000
\end{aligned}
$$




Farthest point initialization


Farthest point initialization





Farthest point initialization





But, very sensitive to outliers


But, very sensitive to outliers


But, very sensitive to outliers


## Robustness to outliers: K-means++

## Principle

Pick the initial centroids as follows
(1) Pick uniformly at random $i \in\left\{x_{1}, \ldots, x_{n}\right\}$, put $c_{1} \leftarrow x_{i}$
(2) $k \leftarrow k+1$
(3) Sample $i \in\left\{X_{1}, \ldots, X_{n}\right\}$ with probability

$$
\frac{\min _{k^{\prime}=1, \ldots, k-1}\left\|x_{i}-c_{k^{\prime}}\right\|_{2}^{2}}{\sum_{i^{\prime}=1}^{n} \min _{k^{\prime}=1, \ldots, k-1}\left\|X_{i^{\prime}}-c_{k^{\prime}}\right\|_{2}^{2}}
$$

(9) Put $c_{k} \leftarrow x_{i}$
(3) If $k<K$ go back to step 2 .

Then use $K$-means based on these initial clusters
This is between random initialization and furthest point initialization

## K-means++

In expectation, leads to a solution close to the optimum. Define the quantification error

$$
Q_{n}\left(c_{1}, \ldots, c_{K}\right)=\sum_{i=1}^{n} \min _{k=1, \ldots, K}\left\|x_{i}-c_{k}\right\|_{2}^{2}
$$

## [Arthur and Vassilvitskii, 2006]

If $c_{1}, \ldots, c_{K}$ are centroids obtained with $K$-means++, then

$$
\mathcal{E}\left[Q_{n}\left(c_{1}, \ldots, c_{K}\right)\right] \leq 8(\log K+2) \min _{c_{1}^{\prime}, \ldots, c_{k}^{\prime}} Q_{n}\left(c_{1}^{\prime}, \ldots, c_{K}^{\prime}\right)
$$

where $\mathcal{E}$ is with respect to random choice of initial centroids

## Complexity

$$
O\left(n \times K \times n_{i t}\right)
$$

## K-means: Pros and Cons

Pros - Simple: easy to implement

- Efficient: guaranteed to converge in finite number of iterations $O\left(n \times K \times n_{i t}\right)$
- Popular

Cons - Notion of mean: means need to be defined

- Number of clusters: K needs to be specified
- Sensitive to outliers
$\mapsto$ can be fixed by subsampling and/or outlier detection
- Roundish clusters: not suited for spherical data, fails if clusters are not convex/round

Mixture models

## Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean


Figure 1: Gaussian Mixture in dimension 1

## Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean

Density mixture - plane [12]


## Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean


Figure 1: Poisson Mixture

## Gaussian Mixture Models (GMM)

## from K-means to GMM


... to GMM


远

## GMM



## Mixture Models

Mixture of densities $f_{1}, \ldots, f_{K}$

- $K \in \mathbb{N}^{*}$ (number of clusters)
- $\left(p_{1}, \ldots, p_{K}\right) \in\left(\mathbb{R}^{+}\right)^{K}$ s.t. $\sum_{k=1}^{K} p_{k}=1$


## Mixture density

$$
f=\sum_{k=1}^{K} p_{k} f_{k}
$$

## Gaussian Mixtures Model (GMM)

- Put $f_{k}=\varphi_{\mu_{k}, \Sigma_{k}}=$ density of $N\left(\mu_{k}, \Sigma_{k}\right)$, where we recall

$$
\varphi_{\mu_{k}, \Sigma_{k}}(x)=\frac{1}{(2 \pi)^{d / 2} \sqrt{\operatorname{det}\left(\Sigma_{k}\right)}} \exp \left(-\frac{1}{2}\left(x-\mu_{k}\right)^{\top} \Sigma_{k}^{-1}\left(x-\mu_{k}\right)\right)
$$

를 with $\Sigma_{k} \succ 0$

## Latent variables

## Proposition (Latent variable)

- Let $\left\{P_{\theta}=f_{\theta} \mu: \theta \in \Theta\right\}$ be a statistical model (for r.v. in $\mathbb{R}^{d}$ ) dominated by $\mu, K$ be a positive integer, $\left(\theta_{1}, \ldots, \theta_{K}\right) \in \Theta^{K}$
- Let $\left(p_{1}, \ldots, p_{K}\right)$ be a probability vector.
- Let now $Z \sim \mathcal{M}\left(1, p_{1}, \ldots, p_{K}\right)$ be a multinomial variable
- $Y:=\sum_{k=1}^{K} k \mathbb{1}_{Z_{k}=1}$.

Then

$$
\forall k \in[K]: \mathbb{P}(Y=k)=p_{k} .
$$

In addition, let $X$ be a random variable such that $X \mid Y \sim P_{\theta_{\gamma}}$. Then

$$
X \sim \sum_{k=1}^{K} p_{k} P_{\theta_{k}} .
$$

## On GMM i

Proposition bridges the gap between mixture models and clustering:

- when $X$ is distributed w.r.t to a mixture model with $k$ components, we describe it with $k$ clusters defined by a latent variable $Y \in[k]$.
- Conversely, clustering is naturally modeled by a mixture model: clusters are distributed w.r.t conditional variables $X \mid Y$ ?

Thus, we focus on the marginal distribution of $X$, which is, by Bayes' theorem:

$$
\forall x \in \mathbb{R}^{d}: f(x)=\sum_{k=1}^{K} p_{k} f_{\theta_{k}}(x)
$$

where $p_{k}=\mathbb{P}(Y=k)$ is the prior probability of a cluster.

## On GMM ii

Then, the Bayes rule for clustering is given by

$$
g^{\star}: x \mapsto \arg \max _{1 \leq k \leq K} \mathbb{P}_{\left(p_{1}, \ldots, p_{k}, \theta\right)}(Y=k \mid X=x)=\arg \max _{1 \leq k \leq K} p_{k} f_{\theta_{k}}(x) .
$$

The final partitioning $\left\{C_{1}, \ldots, C_{K}\right\}$ is $C_{k}=\left\{x \in \mathbb{R}^{d}: g^{\star}(x)=k\right\}$.
$\rightsquigarrow$ explains how to sample $X$ according to a mixture model.

## GMM

## Gaussian Mixtures Model

- Statistical model with density

$$
f_{\theta}=\sum_{k=1}^{K} p_{k} \varphi_{\mu_{k}, \Sigma_{k}}
$$

- Parameter $\theta=\left(p_{1}, \ldots, p_{K}, \mu_{1}, \ldots, \mu_{K}, \Sigma_{1}, \ldots, \Sigma_{K}\right)$
- Goodness-of-fit is:

$$
R_{n}(\theta)=-\log \text {-likelihood }=-\sum_{i=1}^{n} \log \left(\sum_{k=1}^{K} p_{k} \varphi_{\mu_{k}, \Sigma_{k}}\left(x_{i}\right)\right)
$$

- A local minimizer $\hat{\theta}$ is typically obtained using an algorithm called Expectation-Minimization (EM) algorithm


## Gaussian Mixture Models (GMM)

EM algorithm

## EM Algorithm

Idea:

- there is a hidden structure in the model
- knowing this structure, the optimization problem is easier


## EM Algorithm

Idea:

- there is a hidden structure in the model
- knowing this structure, the optimization problem is easier

Indeed, each point $X_{i}$ belongs to an unknown class $k \in\{1, \ldots, K\}$

- Put $C_{i, k}=1$ when $i$ belongs to class $k, C_{i, k}=0$ otherwise.
- We don't observe $\left\{C_{i, k}\right\}_{1 \leq i \leq n, 1 \leq k \leq K}$
- We say that these are latent variables
- Put $\mathcal{C}_{k}=\left\{i: C_{i, k}=1\right\}$, then $\mathcal{C}_{1}, \ldots, \mathcal{C}_{K}$ is a partition of $\{1, \ldots, n\}$.


## Generative model

Generative model:

- i belongs to class $\mathcal{C}_{k}$ with probability $p_{k}$, namely

$$
C_{i}=\left(C_{i, 1}, \ldots, C_{i, K}\right) \sim \mathcal{M}\left(1, p_{1}, \ldots, p_{K}\right)
$$

[multinomial distribution with parameter $\left(1, p_{1}, \ldots, p_{K}\right)$ ].

- $X_{i} \sim \varphi_{\mu_{k}, \Sigma_{k}}$ if $C_{i, k}=1$


## Marginal

- The joint distribution of $(X, C)$ is, according to this model

$$
f_{\theta}(x, c)=\prod_{k=1}^{K}\left(p_{k} \varphi_{\mu_{k}, \Sigma_{k}}(x)\right)^{c_{k}}
$$

[ $c_{k}=1$ for only one $k$ and 0 elsewhere] and the marginal density in $X$ is the one of the mixture:

$$
f_{\theta}(x)=\sum_{k=1}^{K} p_{k} \varphi_{\mu_{k}, \Sigma_{k}}(x)
$$

- This generative model adds a latent variable $C$, in such a way that the marginal distribution of $(X, C)$ in $X$ is indeed the one of the mixture $f_{\theta}$


## Complete Likelihood

- Put $\mathbf{X}=\left(X_{1}, \ldots, X_{n}\right)$ and $\mathbf{C}=\left(C_{1}, \ldots, C_{n}\right)$
- Do as if we observed the latent variables $\mathbf{C}$
- Write a completed likelihood for these "virtual" observations (joint distribution of (X, C)):

$$
L_{c}(\theta ; \mathbf{X}, \mathbf{C})=\prod_{i=1}^{n} \prod_{k=1}^{K}\left(p_{k} \varphi_{\mu_{k}, \Sigma_{k}}\left(X_{i}\right)\right)^{c_{i, k}}
$$

and the completed log-likelihood:

$$
\ell_{c}(\theta ; \mathbf{X}, \mathbf{C})=\sum_{i=1}^{n} \sum_{k=1}^{K} C_{i, k}\left(\log p_{k}+\log \varphi_{\mu_{k}, \Sigma_{k}}\left(X_{i}\right)\right)
$$

## EM-Algorithm

[Dempster et al. (77)]
(E=Expectation, M=Maximization)

Initialize $\theta^{(0)}$ for $t=0, \ldots$, until convergence, repeat:
(1) (E)-step: [Expectation with respect to the latent variables, for the previous value of $\theta$ ]
Compute

$$
\theta \mapsto Q\left(\theta, \theta^{(t)}\right)=\mathcal{E}_{\theta^{(t)}}\left[\ell_{c}(\theta ; \mathbf{X}, \mathbf{C}) \mid \mathbf{X}\right]
$$

(2) (M)-step: [Maximize this expectation]

Compute

$$
\theta^{(t+1)} \in \arg \max _{\theta \in \Theta} Q\left(\theta, \theta^{(t)}\right)
$$

## EM-Algorithm

- (E) and (M) steps often have explicit solutions


## Theorem

The sequence $\theta^{(t)}$ obtained using EM Algorithm satisfies

$$
\ell\left(\theta^{(t+1)} ; \mathbf{X}\right) \geq \ell\left(\theta^{(t)} ; \mathbf{X}\right)
$$

for any $t$.

- A each step, EM increases the likelihood
- Initialization will be very important (usually done using K-Means or K-Means++)


## Proof

Remains to check that $\mathrm{Q}_{1}\left(\theta^{(t+1)}, \theta^{(t)}\right)-\mathrm{Q}_{1}\left(\theta^{(t)}, \theta^{(t)}\right) \leq 0$ :

$$
\begin{aligned}
\mathbf{Q}_{1}\left(\theta^{(t+1)}, \theta^{(t)}\right)-\mathbf{Q}_{1}\left(\theta^{(t)}, \theta^{(t)}\right) & =\mathcal{E}_{\theta^{(t)}}\left[\ell\left(\theta^{(t+1)} ; \mathbf{C} \mid \mathbf{X}\right)-\ell\left(\theta^{(t)} ; \mathbf{C} \mid \mathbf{X}\right) \mid \mathbf{X}\right] \\
& =\int \log \left(\frac{f_{\theta^{(t+1)}}(c \mid \mathbf{X})}{\left.f_{\theta^{(t)}(c \mid \mathbf{X})}\right)}\right) f_{\theta^{(t)}}(c \mid \mathbf{X}) \mu(d c) \\
& \leq \log \int f_{\theta^{(t+1)}}(c \mid \mathbf{X}) \mu(d c)=0
\end{aligned}
$$

This proves $\ell\left(\theta^{(t+1)} ; \mathbf{X}\right) \geq \ell\left(\theta^{(t)} ; \mathbf{X}\right)$ for any $t$

## EM: what? and when?

So what is the EM Algorithm, and where do we use this?

- It is an algorithm that allows to optimize a likelihood with missing or latent data
- For a mixture distribution, we come up with natural latent variables, that simplify the original optimization problem

Gaussian Mixture Models (GMM)
EM and GMM

## EM and GMM (soft Kmeans)

Consider the completed likelihood

$$
\ell_{c}(\theta ; \mathbf{X}, \mathbf{C})=\sum_{i=1}^{n} \sum_{k=1}^{K} C_{i, k}\left(\log p_{k}+\log \varphi_{\mu_{k}, \Sigma_{k}}\left(X_{i}\right)\right)
$$

where

$$
\theta=\left(p_{1}, \ldots, p_{\kappa}, \mu_{1}, \ldots, \mu_{\kappa}, \Sigma_{1}, \ldots, \Sigma_{\kappa}\right) .
$$

What are the ( E ) and ( M ) steps in this case?

## (E)-Step

Compute

$$
\mathcal{E}_{\theta^{(t)}}\left[\ell_{c}(\theta ; \mathbf{X}, \mathbf{C}) \mid \mathbf{X}\right]=\sum_{i=1}^{n} \sum_{k=1}^{K} \mathcal{E}_{\theta^{(t)}}\left[C_{i, k} \mid \mathbf{X}\right]\left(\log p_{k}+\log \varphi_{\mu_{k}, \Sigma_{k}}\left(X_{i}\right)\right)
$$

which is simply:

$$
\mathcal{E}_{\theta}\left[C_{i, k} \mid \mathbf{X}\right]=\mathbb{P}_{\theta}\left(C_{i, k}=1 \mid X_{i}\right)=: \pi_{i, k}(\theta),
$$

where

$$
\pi_{i, k}(\theta)=\frac{p_{k} \varphi_{\mu_{k}, \Sigma_{k}}\left(X_{i}\right)}{\sum_{k^{\prime}=1}^{K} p_{k^{\prime}} \varphi_{\mu_{k^{\prime}}, \Sigma_{k^{\prime}}}\left(X_{i}\right)} .
$$

We call $\pi_{i, k}(\theta)$ the "soft-assignment" of $i$ in class $k$.

## (E)-Step



Figure 2: Soft-assignments : $x \mapsto \frac{p_{k} \varphi_{\mu_{k}, \Sigma_{k}}(x)}{\sum_{k^{\prime}=1}^{K} p_{k^{\prime}} \varphi_{\mu_{k^{\prime}}}, \Sigma_{k^{\prime}}(x)}$.

## (M)-Step

Compute

$$
\theta^{(t+1)}=\left(p_{1}^{(t+1)}, \ldots, p_{K}^{(t+1)}, \mu_{1}^{(t+1)}, \ldots, \mu_{K}^{(t+1)}, \Sigma_{1}^{(t+1)}, \ldots, \Sigma_{K}^{(t+1)}\right)
$$

using:

$$
\begin{aligned}
p_{k}^{(t+1)} & =\frac{1}{n} \sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right), \\
\mu_{k}^{(t+1)} & =\frac{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right) X_{i}}{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)} \\
\Sigma_{k}^{(t+1)} & =\frac{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)\left(X_{i}-\mu_{k}^{(t+1)}\right)\left(X_{i}-\mu_{k}^{(t+1)}\right)^{\top}}{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)} .
\end{aligned}
$$

This is natural: estimation for the means and covariances, weighted by the soft-assignments

So, for $t=1, \ldots$, iterate ( E ) and (M) until convergence:

$$
\begin{aligned}
\pi_{i, k}\left(\theta^{(t)}\right) & =\frac{p_{k}^{(t)} \varphi_{\mu_{k}^{(t)}, \Sigma_{k}^{(t)}}\left(X_{i}\right)}{\sum_{k^{\prime}=1}^{K} p_{k^{\prime}}^{(t)} \varphi_{\mu_{k^{\prime}}(t)}^{(t)} \Sigma_{k^{\prime}}^{(t)}\left(X_{i}\right)} \\
p_{k}^{(t+1)} & =\frac{1}{n} \sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right) \\
\mu_{k}^{(t+1)} & =\frac{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right) X_{i}}{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)} \\
\Sigma_{k}^{(t+1)} & =\frac{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)\left(X_{i}-\mu_{k}^{(t+1)}\right)\left(X_{i}-\mu_{k}^{(t+1)}\right)^{\top}}{\sum_{i=1}^{n} \pi_{i, k}\left(\theta^{(t)}\right)} .
\end{aligned}
$$

We obtain an estimator $\hat{\theta}=\left(\hat{p}_{1}, \ldots, \hat{p}_{K}, \hat{\mu}_{1}, \ldots, \hat{\mu}_{K}, \hat{\Sigma}_{1}, \ldots, \hat{\Sigma}_{K}\right)$.

## Complexity

$$
O\left(n \times K \times n_{i t}\right)
$$



EM-Algorithm on one picture


## Clustering: the maximum a posteriori (MAP) rule

- Given $\hat{\theta}$, how to affect a cluster number $k$ to a given $x \in \mathbb{R}^{d}$ ?
- Compute the the soft-assignments

$$
\pi_{k}(x)=\frac{\hat{p}_{k} \varphi_{\hat{\mu}_{k}, \hat{\Sigma}_{k}}(x)}{\sum_{k^{\prime}=1}^{K} \hat{p}_{k^{\prime}} \varphi_{\hat{\mu}_{k^{\prime}}, \hat{\Sigma}_{k^{\prime}}}(x)}
$$

- Consider

$$
i \in \mathcal{C}_{k} \text { if } \pi_{k}(x)>\pi_{k^{\prime}}(x) \text { for any } k^{\prime} \neq k
$$

## Soft-Assignment

## The MAP rule



Mixture $f_{\theta}(x)$


Soft-assignement $\pi_{k}(x)$

## Soft-Assignment



Figure 3: Soft-assignments : $x \mapsto \frac{p_{k} \varphi_{\mu_{k}, \Sigma_{k}}(x)}{\sum_{k^{\prime}=1}^{K} p_{k^{\prime}} \varphi_{\mu_{k^{\prime}}, \Sigma_{k^{\prime}}}(x)}$.

Gaussian Mixtures Model


## Gaussian Mixtures Model

spherical

diag
Train accuracy: 88.6
Test accuracy: 94.4

tied
Train accuracy: 95.6
Test accuracy: 100.0


## Point-based objectives

## Point-based objectives

Contrarily to center-based objectives, point-based objectives do not require to compute a cluster center.

The distortion $D\left(C_{1}, \ldots, C_{k}\right)$ to minimize is computed based on pair of points belonging to clusters. For example, the sum of in-cluster distances is

$$
\hat{D}\left(C_{1}, \ldots, C_{K}\right)=\sum_{k=1}^{K} \sum_{X, Y \in \hat{C}_{k}} d(X, Y)
$$

with $\hat{C}_{k}=C_{k} \cap\left\{X_{i}: i=1 ; \ldots, n\right\}$

## Point-based objectives

Let $s: \mathcal{X} \times \mathcal{X} \rightarrow[0,1]$ be a similarity measure. Another example lies in the distortion defined by the sum of interclass similarities:

$$
D\left(C_{1}, \ldots, C_{K}\right)=\mathbb{E}\left(\sum_{k=1}^{K} s(X, Y) \mathbb{1}_{X \in C_{k} \cap Y \notin C_{k}}\right)
$$

- center-based approach: making sure that points in the same cluster are similar
- point-based objectives approach: points separated into different clusters should be dissimilar


## Graph-cut problem

Representation by a similarity graph

- each vertex represents a data point $X_{i}$
- vertices are connected by an edge whose weight is their similarity

Such a graph can be defined by the similarity (or adjacency) matrix

$$
W=\left(s\left(X_{i}, X_{j}\right)\right)_{1 \leq i, j \leq n}
$$

Given the index sets $I_{k}$ of each empirical cluster $\hat{C}_{k}$, the previous point-based objective function has an empirical twin given by:

$$
\hat{D}\left(C_{1}, \ldots, C_{k}\right)=\sum_{j=1}^{k} \sum_{\substack{\left.i \in l_{j} \\ \ell \notin\right|_{j}}} W_{i, \ell} .
$$

Minimizing $\hat{D}\left(C_{1}, \ldots, C_{k}\right)$ is often referred as the graph cut problem.

## Similarity graph

Consider a similarity graph $G=(V, E)$, for which the vertices $V=\left(v_{1}, \ldots, V_{n}\right)$ represent the points $\left(X_{1}, \ldots, X_{n}\right)$

- Two vertices $v_{i}$ and $v_{j}$ are connected if the similarity $s\left(X_{i}, X_{j}\right)>0$ (or $>\tau$ with $\tau$ a threshold)
- The edge between these two vertices is weighted by their similarity $s\left(X_{i}, X_{j}\right)$.
- The weighted adjacency matrix is $W=\left(s\left(X_{i}, X_{j}\right)\right)_{1 \leq i, j \leq n}$.
- The graph $G$ is assumed undirected, which is equivalent to $W$ being symmetric (via a symmetric similarity measure s).


## Similarity graph: vocabulary

## Definition

- The degree of a vertex $v_{i} \in V$ is $d_{i}=\sum_{\ell=1}^{n} W_{i, \ell}$.

Given $A \subset V$,

- the size of $A=|A|=$ the number of its vertices
- the volume of $A, \operatorname{vol}(A)=\sum_{i \in[n]: v_{i} \in A} d_{i}$
- $A$ is said connected if any two vertices of $A$ can be joined by a path such that all intermediate points also lie in $A$.
- A is called a connected component if it is connected and if there are no connections between vertices in $A$ and $V \backslash A$.


## Similarity graph

When constructing a similarity graph, the goal is to model the local neighborhood relationships between data points.
$\rightsquigarrow 3$ popular similarity graphs based on a given distance $d: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}_{+}$.

## The $\epsilon$-neighborhood graph

- $X_{i}$ and $X_{j}$ are connected iff $d\left(X_{i}, X_{j}\right) \leq \epsilon$.
- If $\epsilon$ is small enough, all connected points are roughly at the same distance.
- The weights assigned are $W_{i, j}=1$ (if $d\left(X_{i}, X_{j}\right) \leq \epsilon$ and 0 otherwise).
- Usually considered as an unweighted graph


## Similarity graphs

## k-nearest neighbor graph

- $X_{i}$ and $X_{j}$ are connected iff $X_{i} \in k N N\left(X_{j}\right)$ or $X_{j} \in k N N\left(X_{i}\right)$.
- $W_{i, j}=1$ if $X_{i}$ and $X_{j}$ are connected and 0 otherwise.


## Similarity graphs

## k-nearest neighbor graph

- $X_{i}$ and $X_{j}$ are connected iff $X_{i} \in k N N\left(X_{j}\right)$ or $X_{j} \in k N N\left(X_{i}\right)$.
- $W_{i, j}=1$ if $X_{i}$ and $X_{j}$ are connected and 0 otherwise.


## Mutual k-nearest neighbor graph

- $X_{i}$ and $X_{j}$ are connected iff $X_{i} \in \operatorname{kNN}\left(X_{j}\right)$ and $X_{j} \in \operatorname{kNN}\left(X_{i}\right)$.
- $W_{i, j}=1$ if $X_{i}$ and $X_{j}$ are connected and 0 otherwise.


## Similarity graphs

## k-nearest neighbor graph

- $X_{i}$ and $X_{j}$ are connected iff $X_{i} \in \operatorname{kNN}\left(X_{j}\right)$ or $X_{j} \in \operatorname{kNN}\left(X_{i}\right)$.
- $W_{i, j}=1$ if $X_{i}$ and $X_{j}$ are connected and 0 otherwise.


## Mutual k-nearest neighbor graph

- $X_{i}$ and $X_{j}$ are connected iff $X_{i} \in \operatorname{kNN}\left(X_{j}\right)$ and $X_{j} \in \operatorname{kNN}\left(X_{i}\right)$.
- $W_{i, j}=1$ if $X_{i}$ and $X_{j}$ are connected and 0 otherwise.


## The fully connected graph

- Points are connected if they have a similarity $s\left(X_{i}, X_{j}\right)>0$
- The edges are weighted by $s\left(X_{i}, X_{j}\right)$
- A popular choice is the Gaussian similarity: $s\left(x, x^{\prime}\right)=e^{-\frac{d\left(x, x^{\prime}\right)^{2}}{2 \sigma^{2}}}$, in which $\sigma^{2}$ plays a role similar to $\epsilon$ and $k$


## Spectral clustering

## Normalized graph-cut

- For $k=2$, finding a minimal cut of a graph can be done efficiently: Stoer-Wagner algorithm
- Problem: often results is separating a vertex from the rest


## One solution

Normalizing the empirical distortion either by the size of the clusters:

$$
\hat{\mathrm{D}}_{r}\left(C_{1}, \ldots, c_{K}\right)=\sum_{k=1}^{K} \frac{1}{\left|\hat{C}_{k}\right|} \sum_{\substack{i \in \neq k \\ \ell \notin \neq k}} W_{i, \ell},
$$

(let us remind that $\left|\hat{C}_{k}\right|=\left|I_{k}\right|$ ) or by their volume:

$$
\hat{D}_{n}\left(C_{1}, \ldots, C_{K}\right)=\sum_{k=1}^{K} \frac{1}{\operatorname{vol}\left(\hat{C}_{k}\right)} \sum_{\substack{\left.i \in l_{k} \\ \ell \notin\right|_{k}}} W_{i, \ell} .
$$

릉 hese objectives are respectively called ratio cut and normalized cut.

## Towards spectral clustering

- Problem: the balancing introduced by the cluster importance makes the minimization problem computationally hard to solve
- A relaxation procedure: spectral clustering algorithm.


## Definition (Unnormalized graph Laplacian)

Let $W \in \mathbb{R}^{n \times n}$ be a symmetric matrix.

- The diagonal matrix $D \in \mathbb{R}^{n \times n}$ such that $D_{i, i}=\sum_{j=1}^{n} W_{i, j}, \forall i \in[n]$ is called the degree matrix of the graph defined by $W$
- $L=D-W$ is calledthe Laplacian of the graph defined by $W$


## Reformulation of ratio-cut

## Proposition

Let $W$ and $L$ be respectively the adjacency matrix and the Laplacian of the similarity graph of $\left(X_{1}, \ldots, X_{n}\right)$. For any positive integer $K$ and for all partitioning $\left(C_{1}, \ldots, C_{K}\right)$ of $\left(X_{1}, \ldots, X_{n}\right)$, we have

$$
\hat{D}_{r}\left(C_{1}, \ldots, C_{K}\right)=\operatorname{tr}\left(H^{\top} L H\right),
$$

where $H=\left(\frac{1}{\sqrt{|k|}} \mathbb{1}_{\left.i \in\right|_{k}}\right)_{\substack{1 \leq \leq \leq n \\ 1 \leq k \leq K}}$.
In addition, the columns of $H$ are orthonormal to each other $\left(H^{\top} H=I\right)$.

## Proof i

First, denoting $h_{j} \in \mathbb{R}^{n}$ the columns of $H$ (for $j \in[K]$ ), we have

$$
\operatorname{tr}\left(H^{\top} L H\right)=\operatorname{tr}\left(\left(L^{1 / 2} H\right)^{\top}\left(L^{1 / 2} H\right)\right)=\sum_{j=1}^{K}\left(L^{1 / 2} h_{j}\right)^{\top}\left(L^{1 / 2} h_{j}\right)=\sum_{j=1}^{K} h_{j}^{\top} L h_{j} .
$$

## Proof ii

In addition, for all $u \in \mathbb{R}^{n}$,

$$
\begin{aligned}
u^{\top} L u & =u^{\top} D u-u^{\top} W u \\
& =\sum_{1 \leq i \leq n} D_{i, i} u_{i}^{2}-\sum_{1 \leq i, \ell \leq n} W_{i, \ell} u_{i} u_{\ell} \\
& =\frac{1}{2}\left(\sum_{1 \leq i \leq n} D_{i, j} u_{i}^{2}+\sum_{1 \leq \ell \leq n} D_{\ell, \ell} u_{\ell}^{2}-2 \sum_{1 \leq i, \ell \leq n} W_{i, \ell} u_{i} u_{\ell}\right) \\
& =\frac{1}{2}\left(\sum_{1 \leq i, \ell \leq n} W_{i, \ell} u_{i}^{2}+\sum_{1 \leq i, \ell \leq n} W_{i, \ell} u_{\ell}^{2}-2 \sum_{1 \leq i, \ell \leq n} W_{i, \ell} u_{i} u_{\ell}\right) \\
& =\frac{1}{2} \sum_{1 \leq i, \ell \leq n} W_{i, \ell}\left(u_{i}-u_{\ell}\right)^{2} .
\end{aligned}
$$

## Proof iii

Therefore, for all $j \in[K]$,

$$
\begin{aligned}
h_{j}^{\top} L h_{j} & =\frac{1}{2} \sum_{1 \leq i, \ell \leq n} W_{i, \ell}\left(H_{i, j}-H_{\ell, j}\right)^{2} \\
& =\frac{1}{2} \sum_{\substack{i \in l_{j} \\
\ell \notin I_{j}}} \frac{W_{i, \ell}}{\left|I_{j}\right|}+\frac{1}{2} \sum_{\substack{i \neq\left.\right|_{j} \\
\ell \in I_{j}}} \frac{W_{i, \ell}}{\left|I_{j}\right|} \\
& =\frac{1}{\left|I_{j}\right|} \sum_{\substack{i \in l_{j} \\
\ell \notin j_{j}}} W_{i, \ell},
\end{aligned}
$$

since $H_{i, j}-H_{\ell, j}$ is nonzero only if $i \in I_{j}$ and $\ell \notin I_{j}$ or the other way around.

## Proof iv

Gathering everything, we have:

$$
\operatorname{tr}\left(H^{\top} L H\right)=\sum_{j=1}^{K} h_{j}^{\top} L h_{j}=\sum_{j=1}^{K} \frac{1}{\left|I_{j}\right|} \sum_{\substack{\left.\left.i \in\right|_{j} \\ \ell \notin\right|_{j}}} W_{i, \ell}=\hat{D}_{r}\left(C_{1}, \ldots, C_{K}\right) .
$$

## one-hot-encoding

- Up to normalization, $H$ represents the one-hot-encoding.
- For example, for $K=3$, if we reorganize the sample $\left(X_{1}, \ldots, X_{n}\right)$ such that $\hat{C}_{1}$ appears first, then $\hat{C}_{2}$ and so on, we get

$$
H=\left(\begin{array}{ccc}
\frac{1}{\left|\hat{c}_{1}\right|} & 0 & 0 \\
\vdots & \vdots & \vdots \\
\frac{1}{\left|\hat{C}_{1}\right|} & 0 & 0 \\
0 & \frac{1}{\left|\hat{c}_{2}\right|} & 0 \\
\vdots & \vdots & \vdots \\
0 & \frac{1}{\left|\hat{c}_{2}\right|} & 0 \\
0 & 0 & \frac{1}{\left|\hat{C}_{3}\right|} \\
\vdots & \vdots & \vdots \\
0 & 0 & \frac{1}{\left|\hat{c}_{3}\right|}
\end{array}\right) .
$$

## Ratio-cut problem

The ratio cut problem

$$
\min _{\left(\hat{c}_{1}, \ldots, \hat{c}_{K}\right) \in P\left(\left\{X_{1}, \ldots, X_{n}\right\}\right)} \sum_{k=1}^{K} \frac{1}{\left|I_{k}\right|} \sum_{\substack{i \in 1_{k} \\ \ell \notin \dot{k}_{k}}} W_{i, \ell}
$$

is equivalent to

$$
\begin{aligned}
& \min _{H \in \mathbb{R}^{n \times K}} \operatorname{tr}\left(H^{\top} L H\right) \\
& \text { s.t. }\left\{\begin{array}{l}
H^{\top} H=1 \\
\forall j \in[K], \forall i \in[n]: H_{i, j} \in\left\{0, \frac{1}{\sqrt{\mid i, j}}\right\} .
\end{array}\right.
\end{aligned}
$$

## Relaxation for ratio-cut

Problems:

- this is an integer programming problem which we may not be able to solve efficiently.
- the values $\left(\left|I_{1}\right|, \ldots,\left|I_{K}\right|\right)$ are not known in advance.

Idea: discard the last constraint

$$
\begin{gathered}
\min _{H \in \mathbb{R}^{n \times K}} \operatorname{tr}\left(H^{\top} L H\right) \\
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\end{gathered}
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\begin{gathered}
\min _{H \in \mathbb{R}^{n \times K}} \operatorname{tr}\left(H^{\top} L H\right) \\
\text { s.t. } H^{\top} H=l .
\end{gathered}
$$

- Solved by the matrix $H$ for which the columns are the minor eigenvectors of $L$.
- Resulting algorithm: unnormalized spectral clustering.
- maps data $\left(X_{1}, \ldots, X_{n}\right)$ to rows of the $K$ minor eigenvectors of $L$
- then performs a vanilla K-means


## Unnormalized spectral clustering

```
Unnormalized spectral clustering
Require: \(W \in \mathbb{R}^{n \times n}\) (adjacency matrix).
    \(L \leftarrow\) Laplacian of \(W\)
    \(H \leftarrow K\) minor eigenvectors of \(L\) as columns
    \(Y_{i} \leftarrow i^{\text {th }}\) row of \(H\) (for all \(\left.i \in[n]\right)\left\{Y_{i} \in \mathbb{R}^{K}\right\}\)
    \(\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right) \leftarrow\) output of \(K\)-means algorithm based on \(\left(Y_{1}, \ldots, Y_{n}\right)\)
Ensure: \(\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right)\).
```

Idea
(1) Dimension reduction
(2) K-means

## Reformulation for normalized cut

## Proposition

Let $W$ and $L$ be respectively the adjacency matrix and the Laplacian of the similarity graph of $\left(X_{1}, \ldots, X_{n}\right)$. For any positive integer $k$ and for all partitioning $\left(C_{1}, \ldots, C_{K}\right)$ of $\left(X_{1}, \ldots, X_{n}\right)$, we have

$$
\hat{D}_{n}\left(C_{1}, \ldots, C_{K}\right)=\operatorname{tr}\left(H^{\top} L H\right),
$$

where $H=\left(\frac{1}{\sqrt{\operatorname{vol}\left(C_{j}\right)}} \mathbb{1}_{i \in l_{j}}\right)_{\substack{1 \leq \leq \leq n \\ 1 \leq \leq i}}$.
In addition, the columns of $D^{\frac{1}{2}} \mathrm{H}$ are orthonormal to each other ( $H^{\top} D H=I$ ).

## Reformulation for normalized cut

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Let $W$ and $L$ be respectively the adjacency matrix and the Laplacian of the similarity graph of $\left(X_{1}, \ldots, X_{n}\right)$. For any positive integer $k$ and for all partitioning $\left(C_{1}, \ldots, C_{K}\right)$ of $\left(X_{1}, \ldots, X_{n}\right)$, we have

$$
\hat{D}_{n}\left(C_{1}, \ldots, C_{K}\right)=\operatorname{tr}\left(H^{\top} L H\right),
$$

where $H=\left(\frac{1}{\sqrt{\operatorname{vol}\left(G_{j}\right)}} \mathbb{1}_{i \in I_{j}}\right)_{\substack{1 \leq i \leq n \\ 1 \leq \leq \leq K}}$.
In addition, the columns of $D^{\frac{1}{2}} \mathrm{H}$ are orthonormal to each other ( $H^{\top} D H=I$ ).

Proof: similar to the one of the previous Proposition except that we have for all $j \in[K]$,

$$
h_{j}^{\top} L h_{j}=\frac{1}{\operatorname{vol}\left(\hat{C}_{j}\right)} \sum_{\substack{\left.i \in\right|_{j} \\ \ell \notin j}} W_{i, \ell} .
$$

## Relaxation for normalized-cut problem

The normalized cut problem

$$
\min _{\left(\hat{c}_{1}, \ldots, \hat{C}_{k}\right) \in P\left(\left\{X_{1}, \ldots, X_{n}\right\}\right)} \sum_{j=1}^{K} \frac{1}{\operatorname{vol}\left(\hat{C}_{j}\right)} \sum_{\substack{\left.i \in \epsilon_{j} \\ \ell \notin\right|_{j}}} W_{i, \ell}
$$

is equivalent to

$$
\begin{aligned}
\min _{H \in \mathbb{R}^{n \times K}} & \operatorname{tr}\left(H^{\top} L H\right) \\
& H^{\top} D H=1 \\
\text { s.t. } & \forall j \in[K], \forall i \in[n]: H_{i, j} \in\left\{0, \frac{1}{\sqrt{\operatorname{vol}\left(\hat{C}_{\mathrm{i}}\right)}}\right\},
\end{aligned}
$$

and can be relaxed to

$$
\begin{align*}
\min _{H \in \mathbb{R}^{n \times k}} \operatorname{tr}\left(H^{\top} L H\right)  \tag{1}\\
\text { s.t. } H^{\top} D H=I . \tag{2}
\end{align*}
$$

## Relaxation for normalized-cut problem

$$
\begin{align*}
& \min _{H \in \mathbb{R}^{n \times k}} \operatorname{tr}\left(H^{\top} L H\right)  \tag{3}\\
& \text { s.t. } H^{\top} D H=l . \tag{4}
\end{align*}
$$

can be reformulated

$$
\begin{aligned}
\min _{H \in \mathbb{R}^{n \times k}} & \operatorname{tr}\left(U^{\top} L_{s} U\right) \\
\text { s.t. } & H=D^{-\frac{1}{2}} U \\
& U^{\top} U=l
\end{aligned}
$$

where $L_{s}=D^{-\frac{1}{2}} L D^{-\frac{1}{2}}$.

## Relaxation for normalized-cut problem

$$
\begin{align*}
& \min _{H \in \mathbb{R}^{n \times k}} \operatorname{tr}\left(H^{\top} L H\right)  \tag{3}\\
& \text { s.t. } H^{\top} D H=l . \tag{4}
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\text { s.t. } & H=D^{-\frac{1}{2}} U \\
& U^{\top} U=I,
\end{aligned}
$$

where $L_{s}=D^{-\frac{1}{2}} L D^{-\frac{1}{2}}$.

- Solved by $U$ for which the columns are minor eigenvectors of $L_{s}$
- corr. to $H$ for which columns are minor eigenvectors of $L_{w}=D^{-1} L$

릉 - resulting algorithm: normalized spectral clustering.

## Normalized spectral clustering

Normalized spectral clustering (with $L_{w}$ )
Require: $W \in \mathbb{R}^{n \times n}$ (adjacency matrix).
$L_{w} \leftarrow$ Laplacian of $W$
$H \leftarrow K$ minor eigenvectors of $L_{w}$ as columns \{similar to the
generalized eigenproblem $L u=\lambda D u\}$
$Y_{i} \leftarrow i^{\text {th }}$ row of $H$ (for all $\left.i \in[n]\right)\left\{Y_{i} \in \mathbb{R}^{K}\right\}$
$\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right) \leftarrow$ output of $k$-means algorithm based on $\left(Y_{1}, \ldots, Y_{n}\right)$
Ensure: $\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right)$.

Remark: $\lambda \in \mathbb{R}_{+}$is eigenvalue of $L_{w}$ with eigenvector $u$ iff $\lambda$ and $u$ solve the generalized eigenvalue problem $L u=\lambda D u$.

## Ratio cut vs. normalized cut i

- Both objective functions: points separated into different clusters should be dissimilar
- Both take into account the importance of the clusters (by their size or their volume)

Different behavior on cluster importance:

- it is easy to see that, for all $j \in[K]$ :

$$
\sum_{\substack{i \in j_{j} \\ \ell \in j_{j}}} W_{i, \ell}=\operatorname{vol}\left(\hat{C}_{j}\right)-\sum_{\substack{\left.i \in j_{j} \\ \ell \notin\right|_{j}}} W_{i, \ell} .
$$

In other words, the intra-cluster similarity is maximized as soon as the volume of the cluster is maximized and the cut with the rest

## Ratio cut vs. normalized cut ii

of the vertices is minimized; which is what is achieved by normalized cut minimization.

- On the other hand, the size $\left|\hat{C}_{j}\right|$ of a cluster is not necessarily related to the intra-cluster similarity.
$\rightsquigarrow$ normalized cut minimization addresses both parts of clustering
(+) Normalized spectral clustering: $L_{w}$ behaves as expected when $n \rightarrow \infty$
(-) L can lead to completely unreliable results, even for small sample size Von Luxburg, 2007
(+) Normalized spectral clustering: $L_{w}$ behaves as expected when $n \rightarrow \infty$
(-) L can lead to completely unreliable results, even for small sample size Von Luxburg, 2007

There exists another spectral algo:
Normalized spectral clustering (with $L_{s}$ )
Require: $W \in \mathbb{R}^{n \times n}$ (adjacency matrix).
$L_{s} \leftarrow$ Laplacian of $W$
$H \leftarrow K$ minor eigenvectors of $L_{s}$ as columns
$Y_{i} \leftarrow i^{\text {th }}$ row of $H$ normalized to 1 (for all $\left.i \in[n]\right)\left\{Y_{i} \in \mathbb{R}^{K}\right.$,
$\left.\sum_{j=1}^{K}\left(Y_{i}\right)_{j}^{2}=1\right\}$
$\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right) \leftarrow$ output of $k$-means algorithm based on $\left(Y_{1}, \ldots, Y_{n}\right)$
Ensure: $\left(\hat{C}_{1}, \ldots, \hat{C}_{K}\right)$.

## On spectral relaxations

- There is no theoretical guarantees concerning the "quality" of these two relaxations.
- Second, there exist many other relaxations: relying on semidefinite programming
- Spectral relaxations are not appealing for the quality of the solutions they provide but for the simplicity of the problem in which they result (standard linear algebra - eigenvalue problems).


## Properties of graph Laplacian

- $W \in \mathbb{R}_{+}^{n \times n}$ a symmetric adjacency matrix
- $D \in \mathbb{R}^{n \times n}$ its degree matrix

So far, we have seen three Laplacians:

## Definition

Unnormalized Laplacian: $L=D-W$;
Nomalized Laplacian 1: $L_{s}=D^{-\frac{1}{2}} L D^{-\frac{1}{2}}=I-D^{-\frac{1}{2}} W D^{-\frac{1}{2}}$;
Nomalized Laplacian 2: $L_{w}=D^{-1} L=I-D^{-1} W$.
Indexed by s and w: they are respectively symmetrically normalized by $D^{-\frac{1}{2}}$ (on left and right) and whitened by $D$.

## Properties of graph Laplacian

- Bridging gap between graph cut and eigenvalue dec.
- Discover that 0 is an eigenvalue of $L$ and $L_{w}$ with eigenvector $\mathbb{1}$.


## Proposition

(1) $\forall u \in \mathbb{R}^{n}$ :

$$
\begin{aligned}
u^{\top} L u & =\frac{1}{2} \sum_{1 \leq i, \ell \leq n} W_{i, \ell}\left(u_{i}-u_{\ell}\right)^{2} \\
u^{\top} L_{s} u & =\frac{1}{2} \sum_{1 \leq i, \ell \leq n} W_{i, \ell}\left(\frac{u_{i}}{\sqrt{D_{i, i}}}-\frac{u_{\ell}}{\sqrt{D_{\ell, \ell}}}\right)^{2} .
\end{aligned}
$$

(2) 0 is eigenvalue of $L$ and $L_{w}$ with eigenvector $\mathbb{1}$. 0 is eigenvalue of $L_{s}$ with eigenvector $D^{\frac{1}{2}} \mathbb{1}$.
(3) $\lambda \in \mathbb{R}_{+}$is eigenvalue of $L_{w}$ with eigenvector $u$ if and only if $\lambda$ is eigenvalue of $L_{s}$ with eigenvector $D^{\frac{1}{2}} u$.
릉 $L, L_{s}$ and $L_{w}$ are symmetric SDP matrices.

## Properties of graph Laplacian

Proof:
(1) See above.
(2) Obvious.
(3) $\lambda u=L_{w} u \Longleftrightarrow \lambda u=D^{-1} L u \Longleftrightarrow \lambda\left(D^{\frac{1}{2}} u\right)=D^{-\frac{1}{2}} L D^{-\frac{1}{2}}\left(D^{\frac{1}{2}} u\right)$.
(9) Symmetry comes from symmetry of W. SDPness comes from Point 1 and Point 3.

## Properties of graph Laplacian

Proof :
(1) See above.
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(9) Symmetry comes from symmetry of W. SDPness comes from Point 1 and Point 3.

## Proposition

Let $G$ be an undirected graph with non-negative weights.

- the multiplicities of the eigenvalue 0 of $L, L_{s}$ and $L_{w}$ are the same and equal the number $k$ of connected components $\left(A_{1}, \ldots, A_{k}\right)$ in $G$.
- the eigenspace of 0 for both $L$ and $L_{w}$ is spanned by $\left\{\mathbb{1}_{A_{1}}, \ldots, \mathbb{1}_{A_{k}}\right\}$ and the eigenspace of 0 for $L_{s}$ is spanned by $\left\{D^{-\frac{1}{2}} \mathbb{1}_{A_{1}}, \ldots, D^{-\frac{1}{2}} \mathbb{1}_{A_{k}}\right\}$.


## Hierarchical clustering

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## Drawback of K-means

Lack of hierarchy in clusters (i.e. decreasing $K$ does not lead to merging clusters)
$\rightsquigarrow$ Hierachical clustering to address this issue

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## How?

- Introduce very simple methods based on measuring the similarity (or linkage) between clusters.


## Hierarchical clustering

## Drawback of K-means

Lack of hierarchy in clusters (i.e. decreasing $K$ does not lead to merging clusters)
$\rightsquigarrow$ Hierachical clustering to address this issue
How?

- Introduce very simple methods based on measuring the similarity (or linkage) between clusters.

What?

- Focus on agglomerative approaches (which is based on merging clusters) $\rightsquigarrow$ bottom-up
- We put divisive ones aside (based on splitting clusters) $\rightsquigarrow$ top-down


## Agglomerative approaches

Linkage-based methods are probably the simplest and most intuitive paradigm of clustering.

## Agglomerative version

- start from the partitioning of training set $\left(X_{1}, \ldots, X_{n}\right)$ in which each cluster is a unit set $\left\{X_{i}\right\}$ (for $\left.i \in[n)\right]$ )
- merge successively the closest clusters

Straightforwardly,

- the number of clusters decreases at each iteration
- clusters are nested
- each cluster $\hat{C}^{t}$ at iteration $t$ is either the same $\hat{C}^{t}=\hat{C}^{t-1}$ or the union of two previous clusters $\hat{C}^{t}=\hat{C}_{1}^{t-1} \cup \hat{C}_{2}^{t-1}$.
Two parameters need to be defined in such a procedure:
- the (dis)similarity (or linkage) between two clusters
- the merging stopping rule


## Some dissimilarities

Let $d: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}_{+}$be a dissimilarity and consider two subsets $A$ and $B$ of $\left(X_{1}, \ldots, X_{n}\right)$.
Here are some cluster dissimilarities $D: P\left(\left\{X_{1}, \ldots, X_{n}\right\}\right)^{2} \rightarrow \mathbb{R}_{+}$.
Single linkage

$$
D(A, B)=\min _{x \in A, y \in B} d(x, y) .
$$

Complete linkage

$$
D(A, B)=\max _{x \in A, y \in B} d(x, y)
$$



## Some dissimilarities

Average linkage

$$
D(A, B)=\frac{1}{|A||B|} \sum_{x \in A, y \in B} d(x, y) .
$$

## Ward's minimum variance

Given the intraclass inertia for a generic subset $C \subset\left(X_{1}, \ldots, X_{n}\right)$ :

$$
I(C)=\sum_{x \in C} d\left(x, m_{C}\right)^{2}
$$

where $m_{C}=\frac{1}{|C|} \sum_{y \in C} y$, the cluster distance in Ward's method is

$$
D(A, B)=I(A \cup B)-I(A)-I(B),
$$

which is the increase of intraclass inertia when merging $A$ and $B$.

## Linkage

For the Euclidean distance,

$$
D(A, B)=\frac{|A||B|}{|A|+|B|}\left\|m_{A}-m_{B}\right\|^{2} .
$$

Since Ward's method merges clusters by minimizing the increase in the total intraclass inertia, it is very similar k-means but greedy procedure

Linkage methods can be used with a variety of distances (or affinities), in particular:

- Euclidean distance (or I2);
- Manhattan distance (or Cityblock, or I1);
- cosine distance;
- any precomputed affinity matrix.


## Stopping criterion

- If the agglomerative procedure runs until the end, all points share the same large cluster.
- The resulting sequence of partitioning can be represented as a tree, called a dendrogram
- the root = unique cluster that gathers all points (the final cluster)
- the leaves $=$ the unit set clusters (algorithm initialization)


## Stopping rules

- a fixed number of clusters
- a distance upper bound $\bar{D}$ (or alternatively a scaled distance upper bound $\alpha \in \mathbb{R}_{+}$such that $\bar{D}=\alpha \max _{1 \leq i, j \leq n} d\left(X_{i}, X_{j}\right)$ for single, complete and average linkages)


## Agglomerative clustering

## Complexity

- $O\left(n^{3}\right)$ if no restriction on the merging possibilities
- $O\left(n^{2}\right)$ if only a bounded number of merging is possible for a given cluster


## Agglomerative clustering on some pictures



## Cluster Dendogram



## Concluding remarks

K-Means and GMM don't work for "embedded" cluster structures


