Machine Learning – Course n°2

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Unsupervised Learning: Clustering



Unsupervised learning

- Marketing: finding groups of customers with similar behavior given a large database of customer data containing their properties and past buying records
- Biology: classification of plants and animals given their features
- **Insurance:** identifying groups of motor insurance policy holders with a high average claim cost, identifying frauds
- **City-planning:** identifying groups of houses according to their house type, value and geographical location
- Internet: document classification, clustering weblog data to discover groups of similar access patterns



Two main directions

- Data: base of customer data containing their properties and past buying records
- Goal: Use the customers similarities to find groups

Two directions:

- Clustering: propose an explicit grouping of the customers
- Visualization: propose a representation of the customers so that the groups are visibles





Supervised learning reminder

- Training data $D_n = [(x_1, y_1), \dots, (x_n, y_n)]$
- (x_i, y_i) i.i.d \mathbb{P} on $\mathcal{X} \times \mathcal{Y}$
- Construct a predictor $\hat{f} : \mathcal{X} \to \mathcal{Y}$ using D_n
- Loss $\ell(y, f(x))$ measures how well f(x) predicts y well
- Aim is to minimize the generalization error

$$\mathcal{E}_{X,Y}(\ell(Y,\hat{f}(X))|D_n) = \int \ell(y,\hat{f}(x))d\mathbb{P}(x,y).$$

The goal is clear

• Predict y based on feature x

Heard on the street

• Supervised learning is solved. Unsupervised learning isn't



Unsupervised learning

- Training data $D_n = [x_1, \ldots, x_n]$
- Loss: Not Clear
- Aim: Not Clear

The goal is unclear.

Classical tasks

- Clustering: construct groups of data in homogeneous classes
- **Dimension reduction**: construct a *map* of the data in a *low-dimension* space without *distorting* it too much

Motivations

- Interpretation of the groups
- Use of the groups in further processing



Clustering



Clustering

- Training data $D_n = \{x_1, \dots, x_n\}$ with $x_i \in \mathbb{R}^d$
- Recover Latent groups
- Construct $f : \mathbb{R}^d \to \{1, \dots, K\}$ which affects cluster number to x_i

$$f: x_i \mapsto k_i$$

• No ground truth for k_i

Warning

• Choice of K is hard

Roughly two approaches

- Partition-based
- Model-based



K-means

K-means

- Fix $K \ge 2$, *n* data points $x_i \in \mathbb{R}^d$
- Find centroids $c_1, \ldots c_K$ that minimizes the **quantification** error

$$\sum_{i=1}^{n} \min_{k=1,...,K} \|x_i - c_k\|_2^2$$

• Impossible to find the exact solution (NP Complete)



K-means algorithm

Lloyd (1981) proposes a way of finding local solutions

K-means algorithm

- Choose at random K centroids $\{c_1, \ldots, c_K\}$
- For each $k \in \{1, ..., K\}$, find the set C_k of points that are closer to c_k than any $c_{k'}$ for $k' \neq k$
- Update the centroids:

$$c_k = \frac{1}{|C_k|} \sum_{i \in C_k} x_i$$

• Repeat the two previous steps until the sets C_k don't change



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Remark: *K*-means computes a Voronoi partitioning, it implicitly assumes convex clusters, that are uniquely defined by their centroids.



Voronoi partitioning













K-means





K-means





Initialization problem

K-means very sensitive to the choice of initial points





An easy solution:

- Pick a first point at random
- Choose the next initial point the farthest from the previous ones























But, very sensitive to outliers





But, very sensitive to outliers





But, very sensitive to outliers





Robustness to outliers: *K*-means++

Principle

Pick the initial centroids as follows

- Pick uniformly at random $i \in \{x_1, \ldots, x_n\}$, put $c_1 \leftarrow x_i$
- 2 $k \leftarrow k+1$

Sample $i \in \{X_1, \ldots, X_n\}$ with probability

$$\frac{\min_{k'=1,\dots,k-1} \left\| x_i - c_{k'} \right\|_2^2}{\sum_{i'=1}^n \min_{k'=1,\dots,k-1} \left\| X_{i'} - c_{k'} \right\|_2^2}$$

• Put $c_k \leftarrow x_i$

• If k < K go back to step 2.

Then use K-means based on these initial clusters

This is between random initialization and furthest point initialization



In expectation, leads to a solution **close to the** optimum. Define the quantification error

$$Q_n(c_1,\ldots,c_K) = \sum_{i=1}^n \min_{k=1,\ldots,K} \|x_i - c_k\|_2^2.$$

[Arthur and Vassilvitskii, 2006]

If c_1, \ldots, c_K are centroids obtained with *K*-means++, then

$$\mathcal{E}[\mathsf{Q}_n(c_1,\ldots,c_K)] \leq 8(\log K+2)\min_{c_1',\ldots,c_K'}\mathsf{Q}_n(c_1',\ldots,c_K')$$

where $\ensuremath{\mathcal{E}}$ is with respect to random choice of initial centroids

Complexity

$$O(n \times K \times n_{it})$$



Pros • Simple: easy to implement

- Efficient: guaranteed to converge in finite number of iterations $O(n \times K \times n_{it})$
- Popular
- **Cons** Notion of mean: means need to be defined
 - Number of clusters: K needs to be specified
 - Sensitive to outliers
 - \mapsto can be fixed by subsampling and/or outlier detection
 - Roundish clusters: not suited for spherical data, fails if clusters are not convex/round



Mixture models

Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean



Figure 1: Gaussian Mixture in dimension 1



Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean





Model-based clustering

- use a model on data with clusters
- using a mixture of distributions with different location/mean



Figure 1: Poisson Mixture



Gaussian Mixture Models (GMM)

from K-means to GMM





GMM




Mixture Models

Mixture of densities f_1, \ldots, f_K

• $K \in \mathbb{N}^*$ (number of clusters)

•
$$(p_1,\ldots,p_K)\in (\mathbb{R}^+)^K$$
 s.t. $\sum_{k=1}^{K}p_k=1$

Mixture density

$$f = \sum_{k=1}^{K} p_k f_k$$

Gaussian Mixtures Model (GMM)

• Put $f_k = \varphi_{\mu_k, \Sigma_k}$ = density of $N(\mu_k, \Sigma_k)$, where we recall

$$\varphi_{\mu_k,\Sigma_k}(x) = \frac{1}{(2\pi)^{d/2}\sqrt{\det(\Sigma_k)}} \exp\left(-\frac{1}{2}(x-\mu_k)^\top \Sigma_k^{-1}(x-\mu_k)\right)$$



with $\Sigma_k \succ 0$

Latent variables

Proposition (Latent variable)

- Let $\{P_{\theta} = f_{\theta}\mu : \theta \in \Theta\}$ be a statistical model (for r.v. in \mathbb{R}^d) dominated by μ , K be a positive integer, $(\theta_1, \dots, \theta_K) \in \Theta^K$
- Let (p_1, \ldots, p_K) be a probability vector.
- $\bullet \ \ Let \ now \ Z \sim \mathcal{M}(1, p_1, \dots, p_K)$ be a multinomial variable

•
$$\mathbf{Y} := \sum_{k=1}^{K} k \mathbb{1}_{Z_k=1}$$
.

Then

$$\forall k \in [K] \colon \mathbb{P}(Y = k) = p_k.$$

In addition, let X be a random variable such that $X|Y \sim P_{\theta_Y}$. Then

$$X \sim \sum_{k=1}^{K} p_k \mathsf{P}_{ heta_k}.$$



Proposition bridges the gap between mixture models and clustering:

- when X is distributed w.r.t to a mixture model with k components, we describe it with k clusters defined by a latent variable Y ∈ [k].
- Conversely, clustering is naturally modeled by a mixture model: clusters are distributed w.r.t conditional variables X|Y?

Thus, we focus on the marginal distribution of *X*, which is, by Bayes' theorem:

$$\forall x \in \mathbb{R}^d$$
: $f(x) = \sum_{k=1}^{K} p_k f_{\theta_k}(x),$

where $p_k = \mathbb{P}(Y = k)$ is the prior probability of a cluster.



Then, the Bayes rule for clustering is given by

$$g^{\star} \colon x \mapsto \arg \max_{1 \leq k \leq K} \mathbb{P}_{(p_1, \dots, p_K, \theta)} \left(Y = k | X = x \right) = \arg \max_{1 \leq k \leq K} p_k f_{\theta_k}(x).$$

The final partitioning $\{C_1, \ldots, C_K\}$ is $C_k = \{x \in \mathbb{R}^d : g^*(x) = k\}.$

 \rightsquigarrow explains how to sample X according to a mixture model.





Gaussian Mixtures Model

• Statistical model with density

$$f_{\theta} = \sum_{k=1}^{K} p_k \varphi_{\mu_k, \Sigma_k},$$

- Parameter $\theta = (p_1, \dots, p_K, \mu_1, \dots, \mu_K, \Sigma_1, \dots, \Sigma_K)$
- Goodness-of-fit is:

$$R_n(\theta) = -\log$$
-likelihood $= -\sum_{i=1}^n \log \left(\sum_{k=1}^K p_k \varphi_{\mu_k, \Sigma_k}(x_i) \right)$

• A local minimizer $\hat{\theta}$ is typically obtained using an algorithm called **Expectation-Minimization (EM) algorithm**



Gaussian Mixture Models (GMM)

EM algorithm

Idea:

- there is a hidden structure in the model
- knowing this structure, the optimization problem is easier



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- there is a hidden structure in the model
- knowing this structure, the optimization problem is easier

Indeed, each point X_i belongs to an unknown class $k \in \{1, \ldots, K\}$

- Put $C_{i,k} = 1$ when *i* belongs to class *k*, $C_{i,k} = 0$ otherwise.
- We don't observe $\{C_{i,k}\}_{1 \le i \le n, 1 \le k \le K}$
- We say that these are latent variables
- Put $C_k = \{i : C_{i,k} = 1\}$, then C_1, \ldots, C_K is a partition of $\{1, \ldots, n\}$.



Generative model:

• *i* belongs to class C_k with probability p_k , namely

$$C_i = (C_{i,1}, \ldots, C_{i,K}) \sim \mathcal{M}(1, p_1, \ldots, p_K)$$

[multinomial distribution with parameter $(1, p_1, \ldots, p_K)$].

•
$$X_i \sim \varphi_{\mu_k, \Sigma_k}$$
 if $C_{i,k} = 1$



Marginal

• The joint distribution of (X, C) is, according to this model

$$f_{\theta}(x,c) = \prod_{k=1}^{K} (p_k \varphi_{\mu_k, \Sigma_k}(x))^{c_k}$$

 $[c_k = 1$ for only one k and 0 elsewhere] and the marginal density in X is the one of the mixture:

$$f_{\theta}(x) = \sum_{k=1}^{K} p_k \varphi_{\mu_k, \Sigma_k}(x)$$

 This generative model adds a latent variable C, in such a way that the marginal distribution of (X, C) in X is indeed the one of the mixture f_θ



Complete Likelihood

- Put $X = (X_1, ..., X_n)$ and $C = (C_1, ..., C_n)$
- Do as if we observed the latent variables C
- Write a **completed** likelihood for these "virtual" observations (joint distribution of (**X**, **C**)):

$$L_{c}(\theta; \mathbf{X}, \mathbf{C}) = \prod_{i=1}^{n} \prod_{k=1}^{K} (p_{k}\varphi_{\mu_{k}, \Sigma_{k}}(X_{i}))^{C_{i,k}}$$

and the completed log-likelihood:

$$\ell_c(heta; \mathbf{X}, \mathbf{C}) = \sum_{i=1}^n \sum_{k=1}^K C_{i,k}(\log p_k + \log \varphi_{\mu_k, \Sigma_k}(X_i)).$$



EM-Algorithm

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[Dempster et al. (77)]
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(E=Expectation, M=Maximization)
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Initialize $\theta^{(0)}$

for t = 0, ..., until *convergence*, repeat:

• (E)-step: [Expectation with respect to the latent variables, for the previous value of θ]

Compute

$$\theta \mapsto \mathsf{Q}(\theta, \theta^{(t)}) = \mathcal{E}_{\theta^{(t)}}\Big[\ell_{\mathsf{c}}(\theta; \mathsf{X}, \mathsf{C}) \Big| \mathsf{X}\Big]$$

(M)-step: [Maximize this expectation]
 Compute

$$heta^{(t+1)} \in rg\max_{ heta \in \Theta} Q(heta, heta^{(t)})$$



• (E) and (M) steps often have explicit solutions

Theorem

The sequence $\theta^{(t)}$ obtained using EM Algorithm satisfies

 $\ell(\theta^{(t+1)}; \mathbf{X}) \geq \ell(\theta^{(t)}; \mathbf{X})$

for any t.

- A each step, EM increases the likelihood
- Initialization will be very important (usually done using K-Means or K-Means++)



Remains to check that $Q_1(\theta^{(t+1)}, \theta^{(t)}) - Q_1(\theta^{(t)}, \theta^{(t)}) \le 0$:

$$\begin{split} Q_1(\theta^{(t+1)}, \theta^{(t)}) - Q_1(\theta^{(t)}, \theta^{(t)}) &= \mathcal{E}_{\theta^{(t)}} \big[\ell(\theta^{(t+1)}; \mathbf{C} | \mathbf{X}) - \ell(\theta^{(t)}; \mathbf{C} | \mathbf{X}) | \mathbf{X} \big] \\ &= \int \log \Big(\frac{f_{\theta^{(t+1)}}(c | \mathbf{X})}{f_{\theta^{(t)}}(c | \mathbf{X})} \Big) f_{\theta^{(t)}}(c | \mathbf{X}) \mu(dc) \\ &\leq \\ (\text{Jensen})} \log \int f_{\theta^{(t+1)}}(c | \mathbf{X}) \mu(dc) = \mathbf{0}, \end{split}$$

This proves $\ell(\theta^{(t+1)}; \mathbf{X}) \geq \ell(\theta^{(t)}; \mathbf{X})$ for any t



So what is the EM Algorithm, and where do we use this?

- It is an algorithm that allows to optimize a likelihood with missing or latent data
- For a mixture distribution, we come up with natural latent variables, that simplify the original optimization problem



Gaussian Mixture Models (GMM)

EM and GMM

Consider the completed likelihood

$$\ell_{c}(\theta; \mathbf{X}, \mathbf{C}) = \sum_{i=1}^{n} \sum_{k=1}^{K} C_{i,k}(\log p_{k} + \log \varphi_{\mu_{k}, \Sigma_{k}}(X_{i})).$$

where

$$\theta = (p_1, \ldots, p_K, \mu_1, \ldots, \mu_K, \Sigma_1, \ldots, \Sigma_K).$$

What are the (E) and (M) steps in this case?





Compute

$$\mathcal{E}_{\theta^{(t)}}\Big[\ell_{c}(\theta;\mathbf{X},\mathbf{C})\Big|\mathbf{X}\Big] = \sum_{i=1}^{n}\sum_{k=1}^{K}\mathcal{E}_{\theta^{(t)}}[C_{i,k}|\mathbf{X}](\log p_{k} + \log \varphi_{\mu_{k},\Sigma_{k}}(X_{i})),$$

which is simply:

$$\mathcal{E}_{\theta}[C_{i,k}|\mathbf{X}] = \mathbb{P}_{\theta}(C_{i,k} = \mathbf{1}|X_i) =: \pi_{i,k}(\theta),$$

where

$$\pi_{i,k}(\theta) = \frac{p_k \varphi_{\mu_k, \Sigma_k}(X_i)}{\sum_{k'=1}^{K} p_{k'} \varphi_{\mu_{k'}, \Sigma_{k'}}(X_i)}.$$

We call $\pi_{i,k}(\theta)$ the "soft-assignment" of *i* in class *k*.



(E)-Step



Figure 2: Soft-assignments :
$$x \mapsto \frac{p_k \varphi_{\mu_k, \Sigma_k}(x)}{\sum_{k'=1}^K p_{k'} \varphi_{\mu_{k'}, \Sigma_{k'}}(x)}$$
.



(M)-Step

Compute

$$\theta^{(t+1)} = (p_1^{(t+1)}, \dots, p_K^{(t+1)}, \mu_1^{(t+1)}, \dots, \mu_K^{(t+1)}, \Sigma_1^{(t+1)}, \dots, \Sigma_K^{(t+1)})$$

using:

$$\begin{split} p_k^{(t+1)} &= \frac{1}{n} \sum_{i=1}^n \pi_{i,k}(\theta^{(t)}), \\ \mu_k^{(t+1)} &= \frac{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)}) X_i}{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})} \\ \Sigma_k^{(t+1)} &= \frac{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)}) (X_i - \mu_k^{(t+1)}) (X_i - \mu_k^{(t+1)})^\top}{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})}. \end{split}$$

This is natural: estimation for the means and covariances, weighted by the soft-assignments



So, for t = 1, ..., iterate (E) and (M) until *convergence*:

$$\pi_{i,k}(\theta^{(t)}) = \frac{p_k^{(t)}\varphi_{\mu_k^{(t)}, \Sigma_k^{(t)}}(X_i)}{\sum_{k'=1}^{K} p_{k'}^{(t)}\varphi_{\mu_{k'}^{(t)}, \Sigma_{k'}^{(t)}}(X_i)}$$

$$p_k^{(t+1)} = \frac{1}{n} \sum_{i=1}^n \pi_{i,k}(\theta^{(t)})$$

$$\mu_k^{(t+1)} = \frac{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})X_i}{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})(X_i - \mu_k^{(t+1)})(X_i - \mu_k^{(t+1)})^\top}{\sum_{i=1}^n \pi_{i,k}(\theta^{(t)})}.$$

We obtain an estimator $\hat{\theta} = (\hat{p}_1, \dots, \hat{p}_K, \hat{\mu}_1, \dots, \hat{\mu}_K, \hat{\Sigma}_1, \dots, \hat{\Sigma}_K).$

Complexity

 $O(n \times K \times n_{it})$

with *n_{it}* number of iterations

EM-Algorithm on one picture





- Given $\hat{\theta}$, how to affect a cluster number *k* to a given $x \in \mathbb{R}^d$?
- Compute the the soft-assignments

$$\pi_k(\mathbf{x}) = \frac{\hat{p}_k \varphi_{\hat{\mu}_k, \hat{\Sigma}_k}(\mathbf{x})}{\sum_{k'=1}^K \hat{p}_{k'} \varphi_{\hat{\mu}_{k'}, \hat{\Sigma}_{k'}}(\mathbf{x})}$$

• Consider

$$i \in \mathcal{C}_k$$
 if $\pi_k(x) > \pi_{k'}(x)$ for any $k' \neq k$



Soft-Assignment

The MAP rule



Mixture $f_{\theta}(x)$

Soft-assignement $\pi_k(x)$







Gaussian Mixtures Model





Gaussian Mixtures Model





Point-based objectives

Contrarily to center-based objectives, point-based objectives do not require to compute a cluster center.

The distortion $D(C_1, ..., C_k)$ to minimize is computed based on pair of points belonging to clusters. For example, the sum of in-cluster distances is

$$\hat{D}(C_1,\ldots,C_K) = \sum_{k=1}^K \sum_{X,Y \in \hat{C}_k} d(X,Y).$$

with $\hat{C}_k = C_k \cap \{X_i : i = 1; ..., n\}$



Let $s : \mathcal{X} \times \mathcal{X} \rightarrow [0, 1]$ be a similarity measure. Another example lies in the distortion defined by the sum of interclass similarities:

$$D(C_1,\ldots,C_K) = \mathbb{E}\left(\sum_{k=1}^K s(X,Y)\mathbb{1}_{X \in C_k \cap Y \notin C_k}\right)$$

- center-based approach: making sure that points in the same cluster are similar
- point-based objectives approach: points separated into different clusters should be dissimilar



Representation by a similarity graph

- each vertex represents a data point X_i
- vertices are connected by an edge whose weight is their similarity

Such a graph can be defined by the similarity (or adjacency) matrix

$$W = (s(X_i, X_j))_{1 \le i,j \le n}$$

Given the index sets I_k of each empirical cluster \hat{C}_k , the previous point-based objective function has an empirical twin given by:

$$\hat{D}(C_1,\ldots,C_k) = \sum_{j=1}^k \sum_{i \in I_j \atop \ell
eq I_j} W_{i,\ell}.$$

Minimizing $\hat{D}(C_1, \ldots, C_k)$ is often referred as the graph cut problem.



Consider a similarity graph G = (V, E), for which the vertices $V = (v_1, \ldots, v_n)$ represent the points (X_1, \ldots, X_n)

- Two vertices v_i and v_j are connected if the similarity s(X_i, X_j) > 0 (or > τ with τ a threshold)
- The edge between these two vertices is weighted by their similarity s(X_i, X_i).
- The weighted adjacency matrix is $W = (s(X_i, X_j))_{1 \le i,j \le n}$.
- The graph *G* is assumed undirected, which is equivalent to *W* being symmetric (via a symmetric similarity measure *s*).



Definition

• The degree of a vertex $v_i \in V$ is $d_i = \sum_{\ell=1}^n W_{i,\ell}$.

Given $A \subset V$,

- the size of A = |A| = the number of its vertices
- the volume of A, $\operatorname{vol}(A) = \sum_{i \in [n]: v_i \in A} d_i$
- A is said **connected** if any two vertices of A can be joined by a path such that all intermediate points also lie in A.
- A is called a **connected component** if it is connected and if there are no connections between vertices in A and V\A.



When constructing a similarity graph, the goal is to model the local neighborhood relationships between data points.

 $\stackrel{\rightsquigarrow}{\to} \textbf{3 popular similarity graphs based on a given distance} \\ d: \mathcal{X} \times \mathcal{X} \to \mathbb{R}_+.$

The ϵ -neighborhood graph

- X_i and X_j are connected iff $d(X_i, X_j) \leq \epsilon$.
- If ϵ is small enough, all connected points are roughly at the same distance.
- The weights assigned are W_{i,j} = 1 (if d(X_i, X_j) ≤ ε and 0 otherwise).
- Usually considered as an unweighted graph



Similarity graphs

k-nearest neighbor graph

- X_i and X_j are connected iff $X_i \in kNN(X_j)$ or $X_j \in kNN(X_i)$.
- $W_{i,j} = 1$ if X_i and X_j are connected and 0 otherwise.



Similarity graphs

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Mutual k-nearest neighbor graph

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- $W_{i,j} = 1$ if X_i and X_j are connected and 0 otherwise.

The fully connected graph

- Points are connected if they have a similarity $s(X_i, X_j) > 0$
- The edges are weighted by $s(X_i, X_j)$
- A popular choice is the Gaussian similarity: $s(x, x') = e^{-\frac{d(x, x')^2}{2\sigma^2}}$, in which σ^2 plays a role similar to ϵ and k



Spectral clustering

Normalized graph-cut

- For *k* = 2, finding a minimal cut of a graph can be done efficiently: Stoer-Wagner algorithm
- Problem: often results is separating a vertex from the rest

One solution

Normalizing the empirical distortion either by the size of the clusters:

$$\hat{D}_r(C_1,\ldots,C_K) = \sum_{k=1}^K rac{1}{|\hat{C}_k|} \sum_{i \in I_k \atop \ell \notin I_k} W_{i,\ell},$$

(let us remind that $|\hat{C}_k| = |I_k|$) or by their volume:

$$\hat{D}_n(C_1,\ldots,C_K) = \sum_{k=1}^K \frac{1}{\operatorname{vol}(\hat{C}_k)} \sum_{i \in I_k \atop \ell \notin I_k} W_{i,\ell}.$$

hese objectives are respectively called ratio cut and normalized cut.

- Problem: the balancing introduced by the cluster importance makes the minimization problem computationally hard to solve
- A relaxation procedure: spectral clustering algorithm.

Definition (Unnormalized graph Laplacian)

Let $W \in \mathbb{R}^{n \times n}$ be a symmetric matrix.

- The diagonal matrix $D \in \mathbb{R}^{n \times n}$ such that $D_{i,i} = \sum_{j=1}^{n} W_{i,j}, \forall i \in [n]$ is called the *degree matrix* of the graph defined by W
- L = D W is called the *Laplacian* of the graph defined by W



Proposition

Let W and L be respectively the adjacency matrix and the Laplacian of the similarity graph of (X_1, \ldots, X_n) . For any positive integer K and for all partitioning (C_1, \ldots, C_K) of (X_1, \ldots, X_n) , we have

$$\hat{D}_r(C_1,\ldots,C_K)=\mathrm{tr}(H^{\top}LH),$$

where $H = \left(\frac{1}{\sqrt{|I_k|}}\mathbbm{1}_{i \in I_k}\right)_{\substack{1 \leq i \leq n \\ 1 \leq k \leq K}}$.

In addition, the columns of H are orthonormal to each other ($H^{\top}H = I$).



First, denoting $h_j \in \mathbb{R}^n$ the columns of H (for $j \in [K]$), we have

$$tr(H^{\top}LH) = tr((L^{1/2}H)^{\top}(L^{1/2}H)) = \sum_{j=1}^{K} (L^{1/2}h_j)^{\top}(L^{1/2}h_j) = \sum_{j=1}^{K} h_j^{\top}Lh_j.$$



Proof ii

In addition, for all $u \in \mathbb{R}^n$,

$$\begin{split} u^{\top}Lu &= u^{\top}Du - u^{\top}Wu \\ &= \sum_{1 \leq i \leq n} D_{i,i}u_i^2 - \sum_{1 \leq i,\ell \leq n} W_{i,\ell}u_iu_\ell \\ &= \frac{1}{2} \left(\sum_{1 \leq i \leq n} D_{i,i}u_i^2 + \sum_{1 \leq \ell \leq n} D_{\ell,\ell}u_\ell^2 - 2\sum_{1 \leq i,\ell \leq n} W_{i,\ell}u_iu_\ell \right) \\ &= \frac{1}{2} \left(\sum_{1 \leq i,\ell \leq n} W_{i,\ell}u_i^2 + \sum_{1 \leq i,\ell \leq n} W_{i,\ell}u_\ell^2 - 2\sum_{1 \leq i,\ell \leq n} W_{i,\ell}u_iu_\ell \right) \quad (W_{i,\ell} \text{ symm.}) \\ &= \frac{1}{2} \sum_{1 \leq i,\ell \leq n} W_{i,\ell}(u_i - u_\ell)^2. \end{split}$$



Proof iii

Therefore, for all $j \in [K]$,

$$egin{aligned} & h_j^{ op} \mathsf{L} h_j = rac{1}{2} \sum_{1 \leq i, \ell \leq n} \mathsf{W}_{i,\ell} (\mathsf{H}_{i,j} - \mathsf{H}_{\ell,j})^2 \ &= rac{1}{2} \sum_{i \in l_j} rac{\mathsf{W}_{i,\ell}}{|l_j|} + rac{1}{2} \sum_{i \notin l_j \atop \ell \in l_j} rac{\mathsf{W}_{i,\ell}}{|l_j|} \ &= rac{1}{|l_j|} \sum_{i \in l_j \atop \ell \notin l_j} \mathsf{W}_{i,\ell}, \end{aligned}$$

since $H_{i,j} - H_{\ell,j}$ is nonzero only if $i \in I_j$ and $\ell \notin I_j$ or the other way around.



Gathering everything, we have:

$$\operatorname{tr}(H^{\top}LH) = \sum_{j=1}^{K} h_{j}^{\top}Lh_{j} = \sum_{j=1}^{K} \frac{1}{|I_{j}|} \sum_{i \in I_{j} \atop \ell \notin I_{j}} W_{i,\ell} = \hat{D}_{r}(C_{1}, \ldots, C_{K}).$$



one-hot-encoding

- Up to normalization, *H* represents the one-hot-encoding.
- For example, for K = 3, if we reorganize the sample $(X_1, ..., X_n)$ such that \hat{C}_1 appears first, then \hat{C}_2 and so on, we get

$$H = \begin{pmatrix} \frac{1}{|\hat{c}_1|} & 0 & 0\\ \vdots & \vdots & \vdots\\ \frac{1}{|\hat{c}_1|} & 0 & 0\\ 0 & \frac{1}{|\hat{c}_2|} & 0\\ \vdots & \vdots & \vdots\\ 0 & \frac{1}{|\hat{c}_2|} & 0\\ 0 & 0 & \frac{1}{|\hat{c}_3|}\\ \vdots & \vdots & \vdots\\ 0 & 0 & \frac{1}{|\hat{c}_2|} \end{pmatrix}.$$



Ratio-cut problem

The ratio cut problem

$$\min_{(\hat{C}_{1},...,\hat{C}_{K})\in P(\{X_{1},...,X_{n}\})}\sum_{k=1}^{K}\frac{1}{|I_{k}|}\sum_{i\in I_{k}\atop\ell\notin I_{k}}W_{i,\ell}$$

is equivalent to

$$\min_{H \in \mathbb{R}^{n \times K}} \operatorname{tr}(H^{\top} L H)$$

s.t.
$$\begin{cases} H^{\top} H = I \\ \forall j \in [K], \forall i \in [n] : H_{i,j} \in \left\{0, \frac{1}{\sqrt{|I_j|}}\right\} \end{cases}$$



Relaxation for ratio-cut

Problems:

- this is an integer programming problem which we may not be able to solve efficiently.
- the values $(|I_1|, \ldots, |I_K|)$ are not known in advance.

Idea: discard the last constraint

```
\min_{H \in \mathbb{R}^{n \times K}} \operatorname{tr}(H^{\top}LH)
s.t. H^{\top}H = I.
```



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\min_{H \in \mathbb{R}^{n \times K}} \operatorname{tr}(H^{\top}LH)s.t. H^{\top}H = I.
```

- Solved by the matrix *H* for which the columns are the minor eigenvectors of *L*.
- Resulting algorithm: unnormalized spectral clustering.



- maps data (X_1, \ldots, X_n) to rows of the K minor eigenvectors of L
- then performs a vanilla K-means

Unnormalized spectral clustering

Require: $W \in \mathbb{R}^{n \times n}$ (adjacency matrix).

 $L \leftarrow \text{Laplacian of } W$

 $H \leftarrow K$ minor eigenvectors of L as columns

 $Y_i \leftarrow i^{th}$ row of *H* (for all $i \in [n]$) $\{Y_i \in \mathbb{R}^K\}$

 $(\hat{C}_1, \dots, \hat{C}_K) \leftarrow \text{output of } K\text{-means algorithm based on } (Y_1, \dots, Y_n)$ Ensure: $(\hat{C}_1, \dots, \hat{C}_K)$.

Idea

Dimension reduction

Ø K-means



Proposition

Let W and L be respectively the adjacency matrix and the Laplacian of the similarity graph of (X_1, \ldots, X_n) . For any positive integer k and for all partitioning (C_1, \ldots, C_K) of (X_1, \ldots, X_n) , we have

$$\hat{D}_n(C_1,\ldots,C_K) = tr(H^\top L H)$$

where $H = \left(\frac{1}{\sqrt{\operatorname{vol}(C_j)}}\mathbb{1}_{i \in I_j}\right)_{\substack{1 \le i \le n \\ 1 \le j \le K}}$.

In addition, the columns of $D^{\frac{1}{2}}H$ are orthonormal to each other $(H^{\top}DH = I)$.



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In addition, the columns of $D^{\frac{1}{2}}H$ are orthonormal to each other $(H^{\top}DH = I)$.

Proof: similar to the one of the previous Proposition except that we have for all $j \in [K]$,

$$h_j^{\top} L h_j = \frac{1}{\operatorname{vol}(\hat{C}_j)} \sum_{i \in I_j \atop \ell \notin I_j} W_{i,\ell}.$$



Relaxation for normalized-cut problem

The normalized cut problem

$$\min_{(\hat{C}_1,\ldots,\hat{C}_K)\in P(\{X_1,\ldots,X_n\})}\sum_{j=1}^K \frac{1}{\operatorname{vol}(\hat{C}_j)}\sum_{i\in I_j\atop\ell\notin I_j}W_{i,\ell}$$

is equivalent to

$$\begin{split} \min_{H \in \mathbb{R}^{n \times K}} \operatorname{tr}(H^{\top} L H) \\ H^{\top} D H &= I \\ \text{s.t.} \quad \forall j \in [K], \forall i \in [n] \colon H_{i,j} \in \left\{0, \frac{1}{\sqrt{\operatorname{vol}(\hat{C}_j)}}\right\}, \end{split}$$

and can be relaxed to

$$\min_{H \in \mathbb{R}^{n \times k}} \operatorname{tr}(H^{\top}LH)$$
(1)

s.t.
$$H^{\top}DH = I.$$
 (2)



Relaxation for normalized-cut problem

$$\min_{H \in \mathbb{R}^{n \times k}} \operatorname{tr}(H^{\top}LH)$$
(3)

s.t.
$$H^{\top}DH = I.$$
 (4)

can be reformulated

$$\begin{split} \min_{H \in \mathbb{R}^{n \times k}} \operatorname{tr}(U^\top L_s U) \\ \text{s.t.} \quad \begin{array}{l} H = D^{-\frac{1}{2}} U \\ U^\top U = I, \end{array} \end{split}$$

where $L_s = D^{-\frac{1}{2}}LD^{-\frac{1}{2}}$.



Relaxation for normalized-cut problem

$$\min_{H \in \mathbb{R}^{n \times k}} \operatorname{tr}(H^{\top}LH)$$
(3)

s.t.
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can be reformulated

 $\min_{H \in \mathbb{R}^{n \times k}} \operatorname{tr}(U^{\top} L_{s} U)$ s.t. $\begin{array}{l} H = D^{-\frac{1}{2}} U \\ U^{\top} U = I, \end{array}$

where $L_s = D^{-\frac{1}{2}}LD^{-\frac{1}{2}}$.

- Solved by U for which the columns are minor eigenvectors of L_s
- corr. to H for which columns are minor eigenvectors of $L_w = D^{-1}L$
- resulting algorithm: normalized spectral clustering.

Normalized spectral clustering (with L_w)

Require: $W \in \mathbb{R}^{n \times n}$ (adjacency matrix).

 $L_w \leftarrow Laplacian of W$

 $H \leftarrow K$ minor eigenvectors of L_w as columns {similar to the generalized eigenproblem $Lu = \lambda Du$ }

 $Y_i \leftarrow i^{th} \text{ row of } H \text{ (for all } i \in [n]) \{Y_i \in \mathbb{R}^K\}$

 $(\hat{C}_1, \dots, \hat{C}_K) \leftarrow$ output of k-means algorithm based on (Y_1, \dots, Y_n) Ensure: $(\hat{C}_1, \dots, \hat{C}_K)$.

Remark: $\lambda \in \mathbb{R}_+$ is eigenvalue of L_w with eigenvector u iff λ and u solve the generalized eigenvalue problem $Lu = \lambda Du$.



Ratio cut vs. normalized cut i

- Both objective functions: points separated into different clusters should be dissimilar
- Both take into account the *importance* of the clusters (by their size or their volume)

Different behavior on cluster importance:

• it is easy to see that, for all $j \in [K]$:

$$\sum_{i \in I_j \atop \ell \in I_j} W_{i,\ell} = \operatorname{vol}(\hat{C}_j) - \sum_{i \in I_j \atop \ell \notin I_j} W_{i,\ell}.$$

In other words, the intra-cluster similarity is maximized as soon as the volume of the cluster is maximized and the cut with the rest



of the vertices is minimized; which is what is achieved by normalized cut minimization.

- On the other hand, the size $|\hat{C}_j|$ of a cluster is not necessarily related to the intra-cluster similarity.
- ~> normalized cut minimization addresses both parts of clustering



- (+) Normalized spectral clustering: L_w behaves as expected when $n \to \infty$
- (-) *L* can lead to completely unreliable results, even for small sample size Von Luxburg, 2007



(+) Normalized spectral clustering: L_w behaves as expected when $n \to \infty$

 (-) L can lead to completely unreliable results, even for small sample size Von Luxburg, 2007

There exists another spectral algo:

Normalized spectral clustering (with L_s) Require: $W \in \mathbb{R}^{n \times n}$ (adjacency matrix). $L_s \leftarrow$ Laplacian of W $H \leftarrow K$ minor eigenvectors of L_s as columns $Y_i \leftarrow i^{th}$ row of H normalized to 1 (for all $i \in [n]$) { $Y_i \in \mathbb{R}^K$, $\sum_{j=1}^{K} (Y_i)_j^2 = 1$ } $(\hat{C}_1, \dots, \hat{C}_K) \leftarrow$ output of k-means algorithm based on (Y_1, \dots, Y_n) Ensure: $(\hat{C}_1, \dots, \hat{C}_K)$.



- There is no theoretical guarantees concerning the "quality" of these two relaxations.
- Second, there exist many other relaxations: relying on semidefinite programming
- Spectral relaxations are not appealing for the quality of the solutions they provide but for the simplicity of the problem in which they result (standard linear algebra – eigenvalue – problems).



- $W \in \mathbb{R}^{n \times n}_+$ a symmetric adjacency matrix
- $D \in \mathbb{R}^{n \times n}$ its degree matrix

So far, we have seen three Laplacians:

Definition Unnormalized Laplacian: L = D - W; Nomalized Laplacian 1: $L_s = D^{-\frac{1}{2}}LD^{-\frac{1}{2}} = I - D^{-\frac{1}{2}}WD^{-\frac{1}{2}}$; Nomalized Laplacian 2: $L_w = D^{-1}L = I - D^{-1}W$.

Indexed by *s* and *w*: they are respectively symmetrically normalized by $D^{-\frac{1}{2}}$ (on left and right) and whitened by *D*.



Properties of graph Laplacian

- Bridging gap between graph cut and eigenvalue dec.
- Discover that 0 is an eigenvalue of *L* and *L_w* with eigenvector 1.

Proposition

$$\begin{split} u^{\top}Lu &= \frac{1}{2}\sum_{1\leq i,\ell\leq n}W_{i,\ell}(u_i - u_\ell)^2\\ u^{\top}L_s u &= \frac{1}{2}\sum_{1\leq i,\ell\leq n}W_{i,\ell}\left(\frac{u_i}{\sqrt{D_{i,i}}} - \frac{u_\ell}{\sqrt{D_{\ell,\ell}}}\right)^2 \end{split}$$



• $\lambda \in \mathbb{R}_+$ is eigenvalue of L_w with eigenvector u if and only if λ is eigenvalue of L_s with eigenvector $D^{\frac{1}{2}}u$.

\overline{\mathfrak{B}} (a) (a) (b) (c) (c)

Properties of graph Laplacian

Proof :

- See above.
- Obvious.

Symmetry comes from symmetry of W. SDPness comes from Point 1 and Point 3.



Proof :

- See above.
- Obvious.

Symmetry comes from symmetry of W. SDPness comes from Point 1 and Point 3.

Proposition

Let G be an undirected graph with non-negative weights.

- the multiplicities of the eigenvalue 0 of L, L_s and L_w are the same and equal the number k of connected components (A_1, \ldots, A_k) in G.
- the eigenspace of 0 for both L and L_w is spanned by $\{\mathbb{1}_{A_1}, \ldots, \mathbb{1}_{A_k}\}$ and the eigenspace of 0 for L_s is spanned by $\{D^{-\frac{1}{2}}\mathbb{1}_{A_1}, \ldots, D^{-\frac{1}{2}}\mathbb{1}_{A_k}\}$.



Drawback of K-means

Lack of hierarchy in clusters (i.e. decreasing *K* does not lead to merging clusters)

 \rightsquigarrow Hierachical clustering to address this issue



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How?

• Introduce very simple methods based on measuring the similarity (or linkage) between clusters.



Drawback of K-means

Lack of hierarchy in clusters (i.e. decreasing *K* does not lead to merging clusters)

 \rightsquigarrow Hierachical clustering to address this issue

How?

• Introduce very simple methods based on measuring the similarity (or linkage) between clusters.

What?

- Focus on agglomerative approaches (which is based on merging clusters) → bottom-up
- We put divisive ones aside (based on splitting clusters) → top-down



Agglomerative approaches

Linkage-based methods are probably the simplest and most intuitive paradigm of clustering.

Agglomerative version

- start from the partitioning of training set (X₁,...,X_n) in which each cluster is a unit set {X_i} (for i ∈ [n]])
- merge successively the *closest* clusters

Straightforwardly,

- the number of clusters decreases at each iteration
- clusters are nested
- each cluster \hat{C}^t at iteration t is either the same $\hat{C}^t = \hat{C}^{t-1}$ or the union of two previous clusters $\hat{C}^t = \hat{C}_1^{t-1} \cup \hat{C}_2^{t-1}$.

Two parameters need to be defined in such a procedure:

- the (dis)similarity (or linkage) between two clusters
- the merging stopping rule

Some dissimilarities

Let $d : \mathcal{X} \times \mathcal{X} \to \mathbb{R}_+$ be a dissimilarity and consider two subsets A and B of (X_1, \ldots, X_n) .

Here are some cluster dissimilarities $D: P(\{X_1, \ldots, X_n\})^2 \to \mathbb{R}_+$.

Single linkage

$$D(A,B) = \min_{x \in A, y \in B} d(x,y).$$

Complete linkage

$$D(A,B) = \max_{x \in A, y \in B} d(x,y).$$





Some dissimilarities

Average linkage

$$D(A,B) = \frac{1}{|A||B|} \sum_{x \in A, y \in B} d(x,y).$$

Ward's minimum variance

Given the intraclass inertia for a generic subset $C \subset (X_1, \ldots, X_n)$:

$$I(C) = \sum_{x \in C} d(x, m_C)^2,$$

where $m_C = \frac{1}{|C|} \sum_{y \in C} y$, the cluster distance in Ward's method is

$$D(A,B) = I(A \cup B) - I(A) - I(B),$$

which is the increase of intraclass inertia when merging A and B.


Linkage

For the Euclidean distance,

$$D(A,B) = \frac{|A||B|}{|A| + |B|} ||m_A - m_B||^2.$$

Since Ward's method merges clusters by minimizing the increase in the total intraclass inertia, it is very similar k-means but greedy procedure

Linkage methods can be used with a variety of distances (or affinities), in particular:

- Euclidean distance (or I2);
- Manhattan distance (or Cityblock, or I1);
- cosine distance;
- any precomputed affinity matrix.



Stopping criterion

- If the agglomerative procedure runs until the end, all points share the same large cluster.
- The resulting sequence of partitioning can be represented as a tree, called a dendrogram
 - the root = unique cluster that gathers all points (the final cluster)
 - the leaves = the unit set clusters (algorithm initialization)

Stopping rules

- a fixed number of clusters



Complexity

- $O(n^3)$ if no restriction on the merging possibilities
- $O(n^2)$ if only a bounded number of merging is possible for a given cluster



Agglomerative clustering on some pictures





Cluster Dendogram

T





Concluding remarks

K-Means and GMM don't work for "embedded" cluster structures



