

## TD1: Playing with definitions (corrected version)

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**Exercise 1.***Statistical distance*


**Definition 1** (Statistical distance). Let  $X$  and  $Y$  be two discrete random variables over a countable set  $A$ . The statistical distance between  $X$  and  $Y$  is the quantity

$$\Delta(X, Y) = \frac{1}{2} \sum_{a \in A} |\Pr[X = a] - \Pr[Y = a]|.$$

The statistical distance verifies usual properties of distance function, i.e., it is a positive definite binary symmetric function that satisfies the triangle inequality:

- $\Delta(X, Y) \geq 0$ , with equality if and only if  $X$  and  $Y$  are identically distributed,
- $\Delta(X, Y) = \Delta(Y, X)$ ,
- $\Delta(X, Z) \leq \Delta(X, Y) + \Delta(Y, Z)$ .

1. Show that if  $\Delta(X, Y) = 0$ , then for any deterministic adversary  $\mathcal{A}$ , we have  $\text{Adv}_{\mathcal{A}}(X, Y) = 0$ .

 By definition,  $\text{Adv}_{\mathcal{A}}(X, Y) = |\Pr_{a \leftarrow X}[\mathcal{A}(a) = 1] - \Pr_{a \leftarrow Y}[\mathcal{A}(a) = 1]|$ . Since  $\Delta(X, Y) = 0$ , we directly obtain that  $\Pr[X = a] = \Pr[Y = a]$  for all  $a \in S$ , or in other words,  $X$  and  $Y$  are identically distributed. As a result,  $\Pr_{a \leftarrow X}[\mathcal{A}(a) = 1] = \Pr_{a \leftarrow Y}[\mathcal{A}(a) = 1]$  and thus  $\text{Adv}_{\mathcal{A}}(X, Y) = 0$ .

In the next question, we will prove the *data processing inequality* for the statistical distance.

2. Let  $X, Y$  be two random variables over a common set  $A$ .

(a) Let  $f : A \rightarrow S$  be a deterministic function with domain  $S$ . Show that

$$\Delta(f(X), f(Y)) \leq \Delta(X, Y).$$

 We write the definition of  $\Delta$ .

$$\Delta(f(X), f(Y)) = \frac{1}{2} \sum_{s \in S} |\Pr(f(X) = s) - \Pr(f(Y) = s)|$$

Then decompose the event  $f(X) = s$  into something more explicit.

$$\Delta(f(X), f(Y)) = \frac{1}{2} \sum_{s \in S} \left| \sum_{a \in f^{-1}(s)} \Pr(X = a) - \sum_{a \in f^{-1}(s)} \Pr(Y = a) \right|$$

Now use the triangle inequality.

$$\Delta(f(X), f(Y)) \leq \frac{1}{2} \sum_{s \in S} \sum_{a \in f^{-1}(s)} |\Pr(X = a) - \Pr(Y = a)|$$

Finally, recall that  $\sqcup_{s \in S} f^{-1}(s) = A$ , and this ends the proof.

(b) Let  $Z$  be another random variable with domain  $\mathcal{Z}$ , statistically independent from  $X$  and  $Y$ . Show that

$$\Delta((X, Z), (Y, Z)) = \Delta(X, Y).$$

 Once again, we write the definition of the statistical distance.

$$\begin{aligned} \Delta((X, Z), (Y, Z)) &= \sum_{(a,z) \in A \times \mathcal{Z}} |\Pr(X = a \wedge Z = z) - \Pr(Y = a \wedge Z = z)| \\ &= \sum_{(a,z) \in A \times \mathcal{Z}} |\Pr(Z = z) \cdot (\Pr(X = a) - \Pr(Y = a))| \\ &= \sum_{z \in \mathcal{Z}} \Pr(Z = z) \cdot \sum_{a \in A} |\Pr(X = a) - \Pr(Y = a)|. \end{aligned}$$

And this is exactly  $\Delta(X, Y)$ .

- (c) Let  $f$  be a (possibly probabilistic) function with domain  $S$ . Define  $f'$  a deterministic function and  $R$  a random variable independent from  $X$  and  $Y$  such that for any input  $x$ , we have  $f'(x, R) = f(x)$ . The random variable  $R$  is the internal randomness of  $f$ . Using  $f'$  and  $R$ , show that  $\Delta(f(X), f(Y)) = \Delta(f'(X, R), f'(Y, R)) \leq \Delta(X, Y)$ .

$\text{☞}$  We apply the two previous results:  $\Delta(f(X), f(Y)) \leq \Delta((X, R), (Y, R)) = \Delta(X, Y)$ .

3. Show that for any (possibly probabilistic) adversary  $\mathcal{A}$ , we have  $\text{Adv}_{\mathcal{A}}(X, Y) \leq \Delta(X, Y)$ .

$\text{☞}$  This follows from the definition of the advantage, and from the above property ( $\mathcal{A}$  is a function):

$$\text{Adv}_{\mathcal{A}}(X, Y) = |\Pr[\mathcal{A}(X) = 1] - \Pr[\mathcal{A}(Y) = 1]| = \frac{1}{2} \sum_{b \in \{0,1\}} |\Pr[\mathcal{A}(X) = b] - \Pr[\mathcal{A}(Y) = b]| = \Delta(\mathcal{A}(X), \mathcal{A}(Y)) \leq \Delta(X, Y).$$

4. Assuming the existence of a secure PRG  $G : \{0, 1\}^s \rightarrow \{0, 1\}^n$ , show that  $\Delta(G(U(\{0, 1\}^s)), U(\{0, 1\}^n))$  can be much larger than  $\max_{\mathcal{A}} \text{Adv}_{\mathcal{A}}(G(U(\{0, 1\}^s)), U(\{0, 1\}^n))$ .

$\text{☞}$  By definition,

$$\begin{aligned} \Delta(G(U(\{0, 1\}^s)), U(\{0, 1\}^n)) &= \frac{1}{2} \sum_{a \in \{0, 1\}^n} |\Pr[G(U(\{0, 1\}^s)) = a] - \Pr[U(\{0, 1\}^n) = a]| \\ &= \frac{1}{2} \left( \sum_{\substack{a \in \{0, 1\}^n \\ a \notin G(\{0, 1\}^s)}} \left| 0 - \frac{1}{2^n} \right| + \sum_{\substack{a \in \{0, 1\}^n \\ a \in G(\{0, 1\}^s)}} \left| \Pr[G(U(\{0, 1\}^s)) = a] - \frac{1}{2^n} \right| \right) \\ &= \frac{1}{2} - \frac{\#G(\{0, 1\}^s)}{2 \cdot 2^n} + \frac{1}{2} \sum_{\substack{a \in \{0, 1\}^n \\ a \in G(\{0, 1\}^s)}} \left( \Pr[G(U(\{0, 1\}^s)) = a] - \frac{1}{2^n} \right) \\ &= 1 - \frac{\#G(\{0, 1\}^s)}{2^{n+1}} - \frac{\#G(\{0, 1\}^s)}{2^{n+1}} \\ &\geq 1 - 2^{s-n}. \end{aligned}$$

At line 3, we use the fact that for  $a \in G(\{0, 1\}^s)$ , we have  $\Pr[G(U(\{0, 1\}^s)) = a] \geq 1/2^s \geq 1/2^n$  (because at least one element  $b$  is such that  $G(b) = a$  and as  $b$  is chosen uniformly in  $\{0, 1\}^s$ , this happens with probability at least  $2^{-s}$ ). We also use the fact that  $\sum_{\substack{a \in \{0, 1\}^n \\ a \in G(\{0, 1\}^s)}} \Pr[G(U(\{0, 1\}^s)) = a] = 1$ .

For the last inequality, we use the fact that  $\#G(\{0, 1\}^s) \leq 2^s$ . As in the lecture we assumed  $n \gg s$ , then in particular as soon as  $n > s + 1$ , the statistical distance will be greater than  $1/2$ .

On the contrary, as  $G$  is a secure PRG, then by definition  $\max_{\mathcal{A}} \text{Adv}_{\mathcal{A}}(G(U(\{0, 1\}^s)), U(\{0, 1\}^n))$  is negligible, i.e. much smaller than  $1/2$ .

## Exercise 2.

About the advantage definition

We consider two distributions  $D_0$  and  $D_1$  over  $\{0, 1\}^n$ .

1. Recall the definitions that were given in class for the notions of *distinguisher*, *advantage* and *indistinguishability* of  $D_0$  and  $D_1$ .

$\text{☞}$  To sum up the behavior of a distinguisher  $\mathcal{A}$ , two experiments  $\text{Exp}_b, b \in \{0, 1\}$  can be defined as follows.

$\mathcal{C}$	$\mathcal{A}$
sample $x \leftarrow D_b$ send $x$ to $\mathcal{A}$	compute a bit $b'$ output $b'$

Then, we consider the advantage  $\text{Adv}(\mathcal{A}) = |\Pr[\mathcal{A} \xrightarrow{\text{Exp}_0} 1] - \Pr[\mathcal{A} \xrightarrow{\text{Exp}_1} 1]|$ ; the distributions  $D_0$  and  $D_1$  are said to be indistinguishable if  $\text{Adv}(\mathcal{A})$  is negligible for any PPT  $\mathcal{A}$ .

2. Consider a distinguishing game involving two experiments  $\text{Exp}_0, \text{Exp}_1$  in which the adversary is interacting either  $\text{Exp}_b$  for  $b \leftarrow U(\{0, 1\})$ . We define two notions of advantages:

$$\text{Adv}_1(\mathcal{A}) = |\Pr[\mathcal{A} \xrightarrow{\text{Exp}_0} 1] - \Pr[\mathcal{A} \xrightarrow{\text{Exp}_1} 1]| ,$$

and

$$\text{Adv}_2(\mathcal{A}) = |2 \Pr[\mathcal{A} \xrightarrow{\text{Exp}_b} b] - 1| .$$

Show that  $\text{Adv}_1(\mathcal{A}) = \text{Adv}_2(\mathcal{A})$ .


 It follows from the following computation:

$$\begin{aligned} \text{Adv}_2(\mathcal{A}) &= |2 \Pr[\mathcal{A} \xrightarrow{\text{Exp}_b} b] - 1| \\ &= |2(\Pr[\mathcal{A} \xrightarrow{\text{Exp}_b} 1|b=1] \cdot \Pr[b=1] + \Pr[\mathcal{A} \xrightarrow{\text{Exp}_b} 0|b=0] \cdot \Pr[b=0]) - 1| \\ &= |2(\Pr[\mathcal{A} \xrightarrow{\text{Exp}_1} 1] \cdot \frac{1}{2} + \Pr[\mathcal{A} \xrightarrow{\text{Exp}_0} 0] \cdot \frac{1}{2}) - 1| \\ &= |\Pr[\mathcal{A} \xrightarrow{\text{Exp}_1} 1] + (1 - \Pr[\mathcal{A} \xrightarrow{\text{Exp}_0} 1]) - 1| \\ &= \text{Adv}_1(\mathcal{A}) . \end{aligned}$$

### Exercise 3.

*A weird distinguisher...*

We consider two distributions  $D_0$  and  $D_1$  over  $\{0,1\}^n$ . You found a distinguisher  $\mathcal{A}$  on internet. However, you cannot find anywhere in the documentation its performances!

- Assuming that you have access to as many samples as you like from  $D_0$  and  $D_1$  (you can for instance assume that you can sample yourself from these distributions), how would you estimate the advantage of  $\mathcal{A}$ ? *Hint: use the Chernoff Bound:  $\Pr(|X - np| \geq nt) \leq 2 \exp(-2nt^2)$ , where  $X$  follows a binomial distribution with parameters  $(n, p)$ .*  Run  $N$  times Exp 0 and Exp 1 for a number  $N$  to be determined later. This gives us  $b_1^{(1)}, \dots, b_1^{(N)}$  and  $b_2^{(1)}, \dots, b_2^{(N)}$ ,  $2N$  results. Define

$$\bar{b}_1 := \frac{\sum_{i=1}^N b_1^{(i)}}{N} \text{ and } \bar{b}_2 := \frac{\sum_{i=1}^N b_2^{(i)}}{N} .$$

Then let  $p_b$  be the probability that  $\mathcal{A}$  outputs 1 at the end of Exp  $b$ . The Chernoff bound gives

$$\Pr(|\bar{b}_i - p_i| \geq \varepsilon) \leq 2 \exp(-2N\varepsilon^2),$$

for any accuracy  $\varepsilon > 0$ . Then notice the following sequence of inequalities:

$$\text{Adv}(\mathcal{A}) = |p_1 - p_0| \leq |p_1 - \bar{b}_1| + |\bar{b}_1 - \bar{b}_0| + |\bar{b}_0 - p_0| \leq 2\varepsilon + |\bar{b}_1 - \bar{b}_0|,$$

where the last inequality holds with probability at least  $1 - 4 \exp(-2N\varepsilon^2)$ . The same sequence can be written by reversing the roles of  $p_b$  and  $\bar{b}_p$ . This gives us  $|\text{Adv}(\mathcal{A}) - |\bar{b}_1 - \bar{b}_0|| \leq 2\varepsilon$  with probability at least  $1 - 4 \exp(-2N\varepsilon^2)$ .

Assuming that you want to compute the advantage with accuracy  $\frac{1}{\lambda^c}$  and probability 0.95, set  $\varepsilon := \frac{1}{2\lambda^c}$  and  $N$  such that  $1 - 4 \exp(N/(2\lambda^{2c})) \geq 0.95$  i.e.  $N/\lambda^{2c} \geq 2 \ln(80) \approx 8.76$ .

By convention, you want to design a distinguisher such that it outputs 1 when it thinks the sample comes from  $D_1$  and 0 otherwise. However, because of the definition of advantage, it is also possible to design distinguishers that do the reverse, and still have the same advantage. For instance, the above distinguisher  $\mathcal{A}$  may often be "wrong". This could be troublesome if your aim is to use its output to do further computations. Luckily, there exists a way to transform  $\mathcal{A}$  into a distinguisher that is more often right than wrong, whatever it previously did.

- The definition of advantage given in class may be called Absolute Advantage, for the purpose of this exercise. In this question, we define the Positive Advantage of  $\mathcal{A}$  as


$$\text{Adv}_p(\mathcal{A}) := \Pr(\mathcal{A} \xrightarrow{\text{Exp}_1} 1) - \Pr(\mathcal{A} \xrightarrow{\text{Exp}_0} 1).$$

Given a distinguisher  $\mathcal{A}$  with Absolute Advantage  $\varepsilon$ , we build a distinguisher  $\mathcal{A}'$  that does the following:

- Upon receiving a sample  $y \leftrightarrow D_b$ , it runs  $b' \leftarrow \mathcal{A}(y)$ .
- It samples  $x_0 \leftrightarrow D_0$  and  $x_1 \leftrightarrow D_1$  and runs  $b_0 \leftarrow \mathcal{A}(x_0)$  and  $b_1 \leftarrow \mathcal{A}(x_1)$ .
- It returns  $b'$  if  $b_0 = 0$  and  $b_1 = 1$ . It returns  $1 - b'$  if  $b_0 = 1$  and  $b_1 = 0$ .

4. In any other cases, it returns a uniform bit.

Prove that the Positive Advantage of  $\mathcal{A}'$  is  $\epsilon^2$ .

 The probability that  $\mathcal{A}$  outputs 1 in experience  $\text{Exp } b$  is  $p_1(1-p_0)p_b + p_0(1-p_1)(1-p_b) + \frac{1}{2}(p_0p_1 + (1-p_0)(1-p_1))$ . The positive advantage of  $\mathcal{A}'$  is then:

$$\begin{aligned}\text{Adv}_P(\mathcal{A}') &= p_1(1-p_0)(p_1-p_0) + p_0(1-p_1)(p_0-p_1) \\ &= (p_1-p_0) \cdot (p_1(1-p_0) - p_0(1-p_1)) \\ &= (p_1-p_0) \cdot (p_1 - p_0p_1 - p_0 + p_0p_1) \\ &= \epsilon^2.\end{aligned}$$