How to learn while respecting individual privacy?

An Introduction to Differential Privacy

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Outline

Machine Learning

Modeling Privacy

Exponential Mechanism

Laplace Mechanism

Machine Learning & Customer Data Privacy

Src: https://www.actian.com/company/blog/ laymans-guide-to-machine-learning-and-customer-data-privacy/





References



Office is setting on operations are charged with setting information how indexing are exability and an example for a source and an example of the setting of

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The Algorithmic Foundations of Differential Privacy

Cynthia Dwork and Aaron Roth

now

the essence of knowledge

Near Instance-Optimality in Differential Privacy

Hilal Asi John C. Duchi asi@stanford.edu jduchi@stanford.edu Stanford University

Abstract

We decide you so action of instance optimality in differential prince, implied by cluster of a statistical decises: use of definits a local aminor rate in all decisity of workering unblaced modulations and analogizette the Crame-Base bond, and we cluster that the bond works as consistent and the state of the state

1 Introduction

We study instance-precific optimality for differentially private release of a function $f(\mathbf{z}) \in \mathbf{z}$ dataset $\mathbf{z} \in \mathcal{X}^n$. In contrast to existing notions of optimality for private procedures, which measure mechanism' worst case performance over all instances, we develop instance-perific notions to capture the difficulty of—and potential industrivity of private mechanisms to—the given data, \mathbf{z} , rather than some potential version cose.

The trajectory of differential privacy means had private mechanisms reflects the desire to be adaptive both to be formed μ / be compared on diamost μ at μ and μ models μ and μ with the global resultivity $(\Omega_{2j}) = m_{2j} \frac{1}{1-(\mu_{2j})} \frac{1}{(\mu_{2j})} \frac{1}{($

To understand these phenomenas, we take a two-promped approach, presenting both lower bounds on error and complementary (near) optimal mechanisms. We first consider the desiderata a lower bound should satisfy, following a program Cai and Low [11] develop (see also [16]):





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Abstraction

Computer = machine able to make a few elementary operations on data, that can be combined arbitrarily

Problem = description of the desired output for any given input

[3,2,5,1,4]

le petit chat

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[1,2,3,4,5]

5

 \rightarrow the little cat

Examples:

Two approaches

Classical approach : reduction = describe the sequence of elementary operations that permit to construct the output from the input = computer programming (coding)

Artificial Intelligence = use a computer to build itself the program that will
solve the task
= meta-programming

Machine Learning = feed the computer only with a (large number of) *example pairs* (input, output) = search for the program that is supposed to work best on *new examples*

 \rightarrow for each problem, both approaches are possible \rightarrow but they are more or less efficient...



(Cliché) example: MNIST dataset







$$X \in \mathcal{M}_{n,p}(\mathbb{R}) \qquad Y \in \mathcal{Y}^n$$
Classifier \mathcal{A}_n

1

$$h_n: \mathcal{X} \to \mathcal{Y}$$

Data: *n*-by-*p* matrix *x*

- n examples = points of observations
- p features = characteristics measured for each example

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Analysis

ML algorithm

- takes in input $x \in \mathfrak{M}_{n,p}(\mathbb{R})$
- performs some computations so as to optimize some criterion ex: minimizes the training error of some neural network
- returns the optimal predictor *h_n* for future use

- Simple example: return the average value, the proportion of votes for some candidate, etc.

- More difficult: 2-dim representation of the database, image recognition, automatic translation, etc.



Information leakage

Membership attack

Src: https://www.arxiv-vanity.com/papers/1904.05506/



Model inversion attack [Fredrikson et al. '2015]



Figure 1: An image recovered using a new model inversion attack (left) and a training set image of the victim (right). The attacker is given only the person's name and access to a facial recognition system that returns a class confidence score.

See https://arxiv.org/abs/1610.05820 for more information: Membership Inference Attacks against Machine Learning Models by Reza Shokri, Marco Stronati, Congzheng Song, Vitaly Shmatikov



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Anonymization is not the solution

Linkage attack

[*Simple Demographics Often Identify People Uniquely*, by Latanya Sweeney] showed that gender, date of birth, and zip code are sufficient to uniquely identify the vast majority of Americans.

 \implies By linking these attributes in a supposedly anonymized healthcare database to public voter records, she was able to identify the individual health record of the Governor of Massachussetts.

Differencing attack

Imposing request on many lines is not the solution Example from [Dwork & Roth]:

- How many people in the database have the sickle cell trait?
- How many people, not named X, in the database have the sickle cell trait?



Differential Privacy

Differentially private algorithms make assurance that attackers can learn virtually nothing more about an individual than they would understand if that individual's record were absent from the dataset.

Smoker example

if an individual is openly "smoking" but wants privacy on her medical status,

- a medical study will prove the risk associated with smoking (whether she participates or not)
- a *DP* study will make it impossible to know if she indeed participated or not, even to someone who would have all the remaining information

Fundamental Law of Information Recovery: Need to *randomize* the output.





Triathletes doping status $X_i \stackrel{iid}{\sim} \mathcal{B}(p)$ but they may lie: answer $Y_i \in \{0, 1\}$





Triathletes doping status $X_i \stackrel{iid}{\sim} \mathcal{B}(p)$ but they may lie: answer $Y_i \in \{0, 1\}$

RANDOMIZED RESPONSE: A SURVEY TECHNIQUE FOR ELIMINATING EVASIVE ANSWER BIAS

STANLEY L. WARNER Claremont Graduate School

For various reasons individuals in a sample survey may prefer not to confide to the interviewer the correct answers to certain questions. In such assess the individuals may detet not to reply at all of to reply difficult to assess. In this paper is it aproved that and this is potentially removable through allowing the interviewer to maintain privacy through the device or randomizing his response. A monitorial privacy through the device of randomizing his response. A monitorial response method for estimating a population proportion is presented as an example. Unbiased maximum likelihood estimates are obtained and their vantional estimates under various assumptions about the underlying population.

1. INTRODUCTION

 $\label{eq:resonance} \begin{array}{c} \mathbf{F} \mbox{reasons of modesty, fear of being thought bigoted, or merely a reluc$ funce to confide secrets to strangers, many individuals attempt to evadecertain questions put to them by interviewers. In survey vernacular, thesepeople become the "non-cooperative" group [5, pp. 235-72], either refusingoutright to be surveyed, or consenting to be surveyed but purposely providingwrong nanewers to the questions. In the one case there is the problem of refusalbias [1, pp. 355-61], [2, pp. 33-6], [5, pp. 261-9]; in the other case there is the $problem of response bias [3, pp. 30], [4, pp. 208-235]. \end{array}$

Journal of the American Statistical Association, Mar. 1965, Vol.60, No.309, pp. 63-69

See also Chong, Chun Yin Andy & Chu, Amanda & So, Mike & Chung, Ray. (2019). Asking Sensitive Questions Using the Randomized Response Approach in Public

Health Research: An Empirical Study on the Factors of Illegal Waste Disposal. International Journal of Environmental Research and Public Health.





Triathletes doping status $X_i \stackrel{iid}{\sim} \mathcal{B}(p)$ but they may lie: answer $Y_i \in \{0, 1\}$

Randomized Response [Warner'65]

Flip a coin, then:

- $ightarrow \,$ if tails, answer according to another coin flip
- $ightarrow \,$ if heads, give the right answer

$$\mathbb{P}(Y = 1 | X = x) = 1/4 + x/2$$

$$\frac{\mathbb{P}(\mathbf{Y}=1|\mathbf{X}=1)}{\mathbb{P}(\mathbf{Y}=1|\mathbf{X}=0)} = 3$$

- No triathlete can be prosecuted one cannot condemn 1/4th of the innocent triathletes!
- But still permits to estimate the proportion of dopers by $2\bar{Y}_n 1$.

Cost: for the same precision, requires $\approx 4x$ more data or even more if $x(1-x) \ll 1$





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"smoker example": if $\hat{p}=98\%,$ a lot of information on each triathlete BUT no more than if she had not participated in the study







Randomized algorithm $\mathcal{A}(x)$ = random variable on \mathcal{T} **Def:** Neighboring databases $x \sim x'$ if $\exists i \in \{1, ..., n\}, \forall j \neq i, x_{i,\cdot} = x'_{i,\cdot}$

Differential Privacy

 \mathcal{A} is ϵ -DP if for all $x \sim x'$ and all $\mathcal{S} \subset \mathcal{T}$

$$\mathbb{P}(\mathcal{A}(x) \in S) \leq e^{\epsilon} \mathbb{P}(\mathcal{A}(x') \in S)$$

Equivalently,

- $-\epsilon \le \ln \frac{\mathbb{P}(\mathcal{A}(x)=t)}{\mathbb{P}(\mathcal{A}(x')=t)} \le \epsilon \quad \text{ for all } t \in \mathcal{T}$ • if $\mathcal{A}(x)$ is discrete,







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Differential Privacy

[« Calibrating Noise to Sensitivity », TCC'2006, by Cynthia Dwork, Frank McSherry, Kobbi Nissim et Adam Smith ⇒ Gödel Prize 2017]

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```

$$\mathbb{P}(\mathcal{A}(x) \in S) \leq e^{\epsilon} \mathbb{P}(\mathcal{A}(x') \in S)$$

In the previous example on the DP survey, algorithm $\mathcal{A}(x) = (Y_1, \dots, Y_n)$ is $\ln(3)$ -DP.

Note that it outputs an entire (differentially private), which is unusual: more often, we just want the answer to a query.



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```

A person's privacy cannot be compromised by a statistical release if their data are not in the database. Therefore, with differential privacy, the goal is to give each individual roughly the same privacy that would result from having their data removed. That is, the statistical functions run on the database should not overly depend on the data of any one individual.





Randomized algorithm $\mathcal{A}(x)$ = random variable on \mathcal{T} **Def:** Neighboring databases $x \sim x'$ if $\exists i \in \{1, ..., n\}, \forall j \neq i, x_{i,.} = x'_{i,.}$

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$$\mathbb{P}(\mathcal{A}(x) \in S) \leq e^{\epsilon} \mathbb{P}(\mathcal{A}(x') \in S)$$

An algorithm is said to be differentially private if by looking at the output, one cannot tell whether any individual's data was included in the original dataset or not.

Cryptographic origins (and vocabulary).



Randomized algorithm $\mathcal{A}(x)$ = random variable on \mathcal{T} **Def:** Neighboring databases $x \sim x'$ if $\exists i \in \{1, ..., n\}, \forall j \neq i, x_{i,.} = x'_{i,.}$

Differential Privacy

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```

$$\mathbb{P}(\mathcal{A}(x) \in S) \leq e^{\epsilon} \mathbb{P}(\mathcal{A}(x') \in S)$$

Differential privacy mathematically guarantees that anyone seeing the result of a differentially private analysis will essentially make the same inference about any individual's private information, whether or not that individual's private information is included in the input to the analysis.



Properties

Post-processing

If $\mathcal{A} : \mathcal{M}_{n,p}(\mathbb{R}) \to \mathfrak{M}_1(\mathcal{T})$ is ϵ -DP, then for every $f : \mathcal{T} \to \mathcal{T}'$ algorithm $f \circ \mathcal{A}$ is also ϵ -DP.

Group privacy

If $x \sim x_2 \sim \cdots \sim x_k$, then for all $\mathcal{S} \subset \mathcal{T}$, $\mathbb{P}(\mathcal{A}(x) \in S) \leq e^{k\epsilon} \mathbb{P}(\mathcal{A}(x_k) \in S)$.

"Composition"

If $\mathcal{A}_1 : \mathcal{M}_{n,p}(\mathbb{R}) \to \mathfrak{M}_1(\mathcal{T})$ is ϵ -DP and if $\mathcal{A}_2 : \mathcal{M}_{n,p}(\mathbb{R}) \to \mathfrak{M}_1(\mathcal{T}')$ is ϵ' -DP, then $x \mapsto (\mathcal{A}_1(x), \mathcal{A}_2(x))$ is $(\epsilon + \epsilon')$ -DP.

DP defines privacy not as a binary notion of "was the data of individual exposed or not", but rather a matter of accumulative risk.





They already adopted DP

WINTERDAM Several uses of differential privacy in practice are known to date:

- 2008 U.S. Census Bureau, for showing commuting patterns.
- 2014 Google's RAPPOR, for telemetry such as learning statistics about unwanted software hijacking users' settings.
- 2015 Google, for sharing historical traffic statistics.
- 2016 Apple announced its intention to use differential privacy in iOS 10 to improve its Intelligent personal assistant technology.
- 2017 Microsoft, for telemetry in Windows.
- 2019 Privitar Lens is an API using differential privacy.
- 2019 Sarus provides ML with DP as a service.
- 2020 LinkedIn, for advertiser queries.





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Example: Majority of Binary Observations

n = 2k + 1, target $f(x) = \mathbb{1}\left\{\sum x_i \ge n/2\right\} = \text{median}(x)$.

- $\mathcal{A}(x)$ depends only on $s = \sum x_i \implies \mathbb{P}(\mathcal{A}(x) = 1) =: p(s)$
- By symmetry p(n s) = 1 p(s)
- $p(k+1) \le e^{\epsilon}p(k) = e^{\epsilon}(1-p(k+1)) \implies p(k+1) \le \frac{1}{1+e^{-\epsilon}}$
- More generally, for all s > n/2, $p(s) \le \frac{1}{1 + e^{-(2s-n)\epsilon}}$ and $p(n) \le \frac{1}{1 + e^{-(n+2)\epsilon}}$
- In fact, $p(s) = \frac{1}{1 + e^{-(2s-n)\epsilon/2}}$ is ϵ -DP (see next slide) Better: $p(k+r) = \frac{1}{1 + e^{-r\epsilon}}$ is ϵ -DP: $\frac{p(k+r+1)}{p(k+r)} = \epsilon^{\epsilon} \frac{1 + e^{-r\epsilon}}{e^{\epsilon} + e^{-r\epsilon}} \le e^{\epsilon}$ and similarly for $\frac{p(k+1)}{p(k)}$ and $\frac{1 - p(k+r+1)}{1 - p(k+r)}$.
- Requires $n\gg 1/\epsilon$
- + If $|\mathbf{s} \mathbf{n}/2| \geq 3/\epsilon$, the answer is correct with probability $\geq 95\%$
- But if $|s n/2| \le \sqrt{n}$, the chances are high that the majority in the sample is not the majority in the population

• \implies if $\epsilon \ge 3/\sqrt{n} \iff n \ge 9/\epsilon^2$, ϵ -DP does not really cost any reliability!





More generally: Exponential Mechanism

If \mathcal{T} is discrete, one wants \mathcal{A} to assign a probability to each possible outcome $t \in \mathcal{T}$ that depends on its utility u(x, t) on the data x. The sensibility of u is defined as $\Delta u = \max_{t \in \mathcal{T}} \max_{x \sim x'} |u(x, t) - u(x', t)|$.

Exponential Mechanism

The algorithm
$$\mathcal{A}$$
 defined by $\mathbb{P}(\mathcal{A}(x) = t) = \frac{\exp\left(\frac{\epsilon u(x,t)}{2\Delta u}\right)}{\sum_{t'\in\mathcal{T}} \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)}$ is ϵ -DP.

Previous example: for $u(x, t) = (2t - 1)\left(s - \frac{n}{2}\right) = -u(x, 1 - t)$,

$$\mathbb{P}(\mathcal{A}(x)=1) = \frac{\exp\left(\frac{\left(s-\frac{n}{2}\right)\epsilon}{2}\right)}{\exp\left(\frac{\left(s-\frac{n}{2}\right)\epsilon}{2}\right) + \exp\left(-\frac{\left(s-\frac{n}{2}\right)\epsilon}{2}\right)} = \frac{1}{1 + \exp\left(-\left(s-\frac{n}{2}\right)\epsilon\right)}$$



Proof

For every $t \in \mathcal{T}$ and $x \sim x'$,

$$\begin{split} \mathbb{P}(\mathcal{A}(x) = t) &= \frac{\exp\left(\frac{\epsilon u(x,t)}{2\Delta u}\right)}{\sum_{t' \in \mathcal{T}} \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)} \middle/ \frac{\exp\left(\frac{\epsilon u(x',t)}{2\Delta u}\right)}{\sum_{t' \in \mathcal{T}} \exp\left(\frac{u(x',t')\epsilon}{2\Delta u}\right)} \\ &= \exp\left(\frac{\epsilon (u(x,t) - u(x',t))}{2\Delta u}\right) \frac{\sum_{t' \in \mathcal{T}} \exp\left(\frac{(u(x',t') - u(x,t))\epsilon}{2\Delta u}\right) \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)}{\sum_{t' \in \mathcal{T}} \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)} \\ &\leq \exp\left(\frac{\epsilon}{2}\right) \frac{\sum_{t' \in \mathcal{T}} \exp\left(\frac{\epsilon}{2}\right) \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)}{\sum_{t' \in \mathcal{T}} \exp\left(\frac{u(x,t')\epsilon}{2\Delta u}\right)} = \exp(\epsilon) \;. \end{split}$$



Bound on the resulting utility

Theorem

For every database *x* the exponential mechanism satisfies:

$$\mathbb{P}\left(u(x,\mathcal{A}(x)) \leq u(x,f(x)) - \frac{2\Delta u}{\epsilon} \ln \frac{|\mathcal{T}|}{\delta}\right) \leq \delta.$$

Proof: for any $t \in \mathcal{T}$ such that $u(x, t) \leq u(x, f(x)) - 2\Delta u \epsilon^{-1} \ln (|\mathcal{T}|/\delta)$,

$$\mathbb{P}(\mathcal{A}(x) = t) \leq \frac{\exp\left(\frac{\epsilon\left(u\left(x, f(x)\right) - \frac{2\Delta u}{\epsilon} \ln \frac{|\mathcal{T}|}{\delta}\right)\right)}{2\Delta u}\right)}{\exp\left(\frac{\epsilon u(x, f(x))}{2\Delta u}\right)} = \frac{\delta}{|\mathcal{T}|}$$

and there are at most ${\mathcal T}$ of them.



Bound on the resulting utility

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For every database *x* the exponential mechanism satisfies:

$$\mathbb{P}\left(u(x,\mathcal{A}(x)) \leq u(x,f(x)) - \frac{2\Delta u}{\epsilon} \ln \frac{|\mathcal{T}|}{\delta}\right) \leq \delta.$$

Equivalently, for every $\mathbf{v} > 0$

$$\mathbb{P}\left(u(x,\mathcal{A}(x)) \leq u(x,f(x)) - v\right) \leq |\mathcal{T}|e^{-\frac{\epsilon v}{2\Delta_u}}$$

In the example:

$$\mathbb{P}(\mathcal{A}(x) \neq f(x)) = \mathbb{P}\left(u(x, \mathcal{A}(x)) \leq u(x, f(x)) - 2u(x, f(x))\right) \leq 2\exp\left(-\frac{u(x, f(x))\epsilon}{\Delta u}\right) = 2e^{-\left|s-\frac{n}{2}\right|\epsilon}.$$



Lower Bound for Discrete Mechanisms

 $\mathcal A$ (discrete) is said to be unbiased if for x and $t\in\mathcal T$,

$$\mathbb{P}(\mathcal{A}(x) = t) \leq \mathbb{P}(\mathcal{A}(x) = f(x))$$
.

Inverse Sensibility

The inverse sensibility of function *f* on data *x* to output $t \in T$ is defined as

$$\mathcal{D}_f(x,t) = \min\left\{k : \exists x \sim x_1 \sim ... \sim x_k \text{ and } f(x_k) = t\right\}.$$

Lower bound

For every unbiased ϵ -DP mechanism \mathcal{A} , $\mathbb{P}(\mathcal{A}(x) = f(x)) \leq \frac{1}{\sum_{t \in \mathcal{T}} e^{-2D_f(x,t)\epsilon}}$.



Proof

Let $t \in \mathcal{T}$ and let x' be such that $x \sim x_1 \sim ... \sim x_{\mathcal{D}_f(x,t)-1} \sim x'$ and f(x') = t. Differential privacy implies

$$\frac{\mathbb{P}(\mathcal{A}(x)=t)}{\mathbb{P}(\mathcal{A}(x)=f(x))} = \frac{\mathbb{P}(\mathcal{A}(x)=t)}{\mathbb{P}(\mathcal{A}(x')=t)} \frac{\mathbb{P}(\mathcal{A}(x')=t=f(x'))}{\mathbb{P}(\mathcal{A}(x')=f(x))} \frac{\mathbb{P}(\mathcal{A}(x')=f(x))}{\mathbb{P}(\mathcal{A}(x)=f(x))} \ge e^{-\mathcal{D}_f(x,t)\epsilon} \times 1 \times e^{-\mathcal{D}_f(x,t)\epsilon}$$

and hence

$$1 = \sum_{t \in \mathcal{T}} \frac{\mathbb{P}(\mathcal{A}(x) = t)}{\mathbb{P}(\mathcal{A}(x) = f(x))} \mathbb{P}(\mathcal{A}(x) = f(x)) \ge \sum_{t \in \mathcal{T}} e^{-2\mathcal{D}_f(x,t)\epsilon} \mathbb{P}(\mathcal{A}(x) = f(x)) .$$

In the previous example, $\mathcal{D}_f(x, 1 - f(x)) = \left|s - \frac{n}{2}\right| + \frac{1}{2}$ and this yields: $\mathbb{P}(\mathcal{A}(x) = f(x)) \leq \frac{1}{1 + e^{-(|2s-n|+1)\epsilon}}$.

The Exponential Mechanism above is almost optimal: it has $\mathbb{P}(\mathcal{A}(x) = f(x)) = \frac{1}{1 + e^{-\frac{|2s-n|\epsilon}{2}}}$.



Inverse Sensitivity Mechanism

The inverse sensibility D_f yields a good candidate utility function for an exponential mechanism! In fact, $\Delta D_f = 1$, hence the ϵ -DP Inverse Sensitivity Mechanism (ISM)

$$\mathbb{P}(\mathcal{A}(x) = t) = \frac{e^{-\epsilon \mathcal{D}_f(x,t)/2}}{\sum_{t' \in \mathcal{T}} e^{-\mathcal{D}_f(x,t')\epsilon/2}} .$$

Remark: if $|\mathcal{T}| = 2$ the denominator 2 is not needed:

$$\mathbb{P}(\mathcal{A}(x) = t) = \frac{e^{-\epsilon \mathcal{D}_f(x,t)}}{\sum_{t' \in \mathcal{T}} e^{-\mathcal{D}_f(x,t')\epsilon}}$$

is ϵ -DP. Previous example: here $\mathcal{D}_f(x, 1 - f(x)) = \left|s - \frac{n}{2}\right| + \frac{1}{2}$, the ISM $p(s) = \frac{1}{1 + e^{-(s-k-1\{s \le k\})\epsilon}}$ is an ϵ -DP mechanism slightly better than the exponential mechanism above.



Near-Optimality of the Inverse Senbility Mechanism

1/4-Optimality of the ISM

The ISM A is "more accurate" than any $\epsilon/4$ -DP algorithm A':

 $\mathbb{P}(\mathcal{A}'(x) = f(x)) \leq \mathbb{P}(\mathcal{A}(x) = f(x))$.

Proof: Since $\mathcal{D}_f(x, f(x)) = 0$, $\mathbb{P}(\mathcal{A}(x) = f(x)) = 1/\sum_{t \in \mathcal{T}} e^{-\mathcal{D}_f(x, t)\epsilon/2}$. Recall the lower bound: for every unbiased ϵ -DP mechanism \mathcal{A}' , $\mathbb{P}(\mathcal{A}'(x) = f(x)) \leq 1/\sum_{t \in \mathcal{T}} e^{-2D_f(x, t)\epsilon}$.

Remark: if $|\mathcal{T}| = 2$, the ISM is 1/2-optimal.



Continuous Exponential Mechanism

If continuous \mathcal{T} , taking $\mathcal{A}(x)$ with density

$$f_x(t) = rac{\exp\left(rac{\epsilon u(x,t)}{2\Delta u}
ight)}{\int_{\mathcal{T}} \exp\left(rac{u(x,t')\epsilon}{2\Delta u}
ight) dt'}$$

also yields an $\epsilon\text{-}\mathsf{DP}$ mechanism.

- The ISM is hence a very good candidate in theory.
- It is reminiscent of statistical physics "Gibbs law" (thermodynamics).
- It can be hard to sample from.
- In fact, the discrete case is already computationally challenging when the output space ${\cal T}$ is "big".
- Research question: does approximate sampling preserve differential privacy?



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Example: estimate the mean

Here
$$x \in \mathfrak{M}_{n,1}(\{0,1\})$$
 and $f(x) = \bar{x}$. $\mathcal{T} = \{0, \frac{1}{n}, \dots, \frac{n-1}{n}, 1\}$.
Inverse sensibility function: $\mathcal{D}_f(x,t) = n|t-\bar{x}|$.
ISM: $\mathbb{P}(\mathcal{A}(x) = t) = \frac{e^{-\epsilon n|t-\bar{x}|/2}}{\sum_{j=0}^n e^{-\epsilon n|j/n-\bar{x}|/2}}$.
 $\mathbb{P}\left(\mathcal{A}(x) \ge \bar{x} + \frac{k}{n}\right) = \frac{\sum_{j=k}^n e^{-\epsilon j/2}}{2\sum_{j=0}^n e^{-\epsilon j/2}} = \frac{e^{-\epsilon k/2}}{2}$ and $\mathbb{P}\left(\mathcal{A}(x) \le \bar{x} - \frac{k}{n}\right) = \frac{e^{-\epsilon k/2}}{2}$

 \implies up to the discretization, $\mathcal{A}(x)$ has the same distribution as

 $\mathcal{A}'(x) = \bar{x} + Y$

where $Y \sim \operatorname{Lap}\left(\frac{2}{n\epsilon}\right)$ has density $\ell_{\frac{2}{n\epsilon}}(x) = \frac{n\epsilon}{4}e^{-\frac{n\epsilon|x|}{2}}$ since $\mathbb{P}(\mathcal{A}'(x) \ge \bar{x} + k/n) = \frac{e^{-\epsilon k/2}}{2}$.

⇒ very simple mechanism: just add (well-calibrated) Laplace noise!



Laplace mechanism

The L^1 -sensitivity of a function $f: \mathfrak{M}_{n,p}(\mathbb{R}) \to \mathbb{R}^k$ is defined as

 $\Delta f = \max_{x \sim x'} \|f(x) - f(x')\|_1 .$

Example: if $f(x) = n^{-1} \sum x_i$ and $x_i \in [0, 1]$, then $\Delta(f) = 1/n$. Lap (σ) has density $\ell_{\sigma}(x) = \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}}$.

Laplace Mechanism

The Laplace mechanism for $f: \mathfrak{M}_{n,p}(\mathbb{R}) \to \mathbb{R}^k$ defined by

$$\mathcal{A}(x) = f(x) + (Y_1, \dots, Y_k), \qquad Y_i \stackrel{iid}{\sim} \operatorname{Lap}\left(\frac{\Delta f}{\epsilon}\right) \;.$$

is ϵ -differentially private.



Proof

For $t \in \mathbb{R}^k$ let

$$p_{x}(t) = \prod_{j=1}^{k} \frac{\epsilon}{2\Delta f} \exp\left(-\frac{\epsilon |t_{j} - f(x)_{j}|}{\Delta f}\right) = \left(\frac{\epsilon}{2\Delta f}\right)^{k} \exp\left(-\frac{\epsilon ||t - f(x)||_{1}}{\Delta f}\right)$$

be the density of $\mathcal{A}(x)$. Then for every $x \sim x'$ and every $t \in \mathbb{R}^k$,

$$\frac{p_{x}(t)}{p_{x'}(t)} = \frac{\exp\left(-\frac{\epsilon \left\|t-f(x)\right\|_{1}}{\Delta f}\right)}{\exp\left(-\frac{\epsilon \left\|t-f(x')\right\|_{1}}{\Delta f}\right)} = \exp\left(-\frac{\epsilon \left(\left\|t-f(x)\right\|_{1}-\left\|t-f(x')\right\|_{1}\right)}{\Delta f}\right)$$
$$\leq \exp\left(\frac{\epsilon \left\|f(x)-f(x')\right\|_{1}}{\Delta f}\right) \leq \exp(\epsilon)$$

since by definition $||f(x) - f(x')||_1 \le \Delta f$.





Examples

- Counting queries accuracy independent of *n*
- Histogram queries accuracy independent of *n* and of the number of bins!
- Most common value of a variable
- Noisy max accuracy depends on the number of values, but not on *n*
- Estimate average (or sum)

▲ Finite precision arithmetic \implies possible privacy leaks For example, one can have $\mathbb{P}(\mathcal{A}(x) = t) > 0$ while $\mathbb{P}(\mathcal{A}(x') = t) = 0$ for two neighbors $x \sim x'$ because of rounding.



Lower bound for continuous mechanisms

We consider a target function $f: \mathfrak{M}_{n,p}(\mathbb{R}) \to \mathbb{R}$. For every $x \in \mathfrak{M}_{n,p}(\mathbb{R})$, let $\Delta_f(x,k) = \sup \{ |f(x_k) - f(x)| : x \sim x_1 \sim \cdots \sim x_k \}$ and $\Delta_f(x) = \Delta_f(x, 1)$.

Lower bound

 \mathcal{A} is unbiased if for every x, $\mathbb{E}[\mathcal{A}(x)] = f(x)$. Then, for every $x \in \mathfrak{M}_{n,p}(\mathbb{R})$,

$$\mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right] \ge \sup_k \frac{\Delta_f(x,k)^2/4}{1 + e^{2k\epsilon}}$$

and in particular if $\Delta_{f}(x,k) = k \Delta_{f}(x)$, then for $\epsilon \leq 1/2$

$$\mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right] \ge \frac{\Delta_f(x)^2}{68\epsilon^2}$$

For the average $f(x) = \overline{x}_n$, $\Delta_f(x, k) = k\Delta_f(x) = k/n$ and $\mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right] \ge \frac{1}{68n^2\epsilon^2}$.



Proof

Let $k \ge 1$ and let x_k be such that $x \sim x_1 \sim \cdots \sim x_k$ and $|f(x_k) - f(x)| = \Delta_f(x, k)$. By definition,

$$\mathbb{E}\left[\left(\mathcal{A}(x_k)-t\right)^2\right] = \int_{\mathcal{T}} (s-t)^2 d\mathbb{P}_{\mathcal{A}(x_k)}(s) \leq \int_{\mathcal{T}} (s-t)^2 e^{k\epsilon} d\mathbb{P}_{\mathcal{A}(x)}(s) = e^{k\epsilon} \mathbb{E}\left[\left(\mathcal{A}(x)-t\right)^2\right]$$

and hence

$$\frac{\mathbb{E}\left[\left(\mathcal{A}(x_{k})-f(x_{k})\right)^{2}\right]}{\mathbb{E}\left[\left(\mathcal{A}(x)-f(x)\right)^{2}\right]}=\frac{\mathbb{E}\left[\left(\mathcal{A}(x_{k})-f(x_{k})\right)^{2}\right]}{\mathbb{E}\left[\left(\mathcal{A}(x_{k})-f(x)\right)^{2}\right]}\frac{\mathbb{E}\left[\left(\mathcal{A}(x_{k})-f(x)\right)^{2}\right]}{\mathbb{E}\left[\left(\mathcal{A}(x_{k})-f(x)\right)^{2}\right]}\leq 1\times e^{k\epsilon}$$

since ${\mathcal A}$ is unbiased. Therefore, by the Bienaymé-Chebishev inequality

$$\begin{split} & \mathbb{P}\left(\left|\mathcal{A}(x) - f(x)\right| \geq \frac{\Delta_f(x,k)}{2}\right) \leq \frac{4 \mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right]}{\Delta_f(x,k)^2} \text{ and} \\ & \mathbb{P}\left(\left|\mathcal{A}(x) - f(x_k)\right| \geq \frac{\Delta_f(x,k)}{2}\right) \leq e^{k\epsilon} \mathbb{P}\left(\left|\mathcal{A}(x_k) - f(x_k)\right| \geq \frac{\Delta_f(x,k)}{2}\right) \leq \frac{4e^{2k\epsilon} \mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right]}{\Delta_f(x,k)^2} \text{ .} \\ & \text{But since } |f(x_k) - f(x)| = \Delta_f(x,k), \\ & 1 \leq \mathbb{P}\left(\left|\mathcal{A}(x) - f(x)\right| \geq \frac{\Delta_f(x,k)}{2}\right) + \mathbb{P}\left(\left|\mathcal{A}(x) - f(x_k)\right| \geq \frac{\Delta_f(x,k)}{2}\right) \leq \frac{4\left(1 + e^{2k\epsilon}\right) \mathbb{E}\left[\left(\mathcal{A}(x) - f(x)\right)^2\right]}{\Delta_f(x,k)^2} \end{split}$$

The second statement is obtained by the choice $k = \lfloor 1/(2\epsilon) \rfloor$, noting that $1/(2\epsilon) \le k \le 1/\epsilon$.



Working Research: Multi-quantile Estimation

Differentially Private Quantiles

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Abstract

Quantiles are often used for summarizing and understanding data. If that data is sensitive, it may be necessary to compute quantiles in a way that is differentially private, providing theoretical guarantees that the result does not reveal private information. However, when multiple quantiles are needed, existing differentially private algorithms frage poorly: they either compute quantiles individually, splitting the privacy budget, or summarize the entire distribution, wasting effort. In either case the result is reduced accuracy. In this work we propose an instance of the exponential mechanism that simultaneously estimates exactly mean prior model of the order of the efficient inplementation that returns estimates of all m quantiles from $O(m) + m^2 n)$. Experiments show that our method significantly outperforms the current state of the art on both real and synthetic data while remaining efficient enough to be practical.

1 Introduction

Quantiles are a widespread method for understanding real-world data, with example applications ranging from income [29] to birth weight [8] to standardized test scores [16]. At the same time, the individual contributing data may require that these quantiles not reveal too much information about individual contributions. As a toy example, suppose that an individual joins a company that has exactly two salaries, and half of current employees have one salary and half have another. In this case, publishing the exact median company salary will reveal the new employee's salary.

Differential privacy [14] offers a solution to this problem. Informally, the distribution over a



