

## Solutions for Exercise sheet 7: Donsker's invariance principle (v3)

version 3: typo fixed thanks to Romain Durand!

**Solution 1** — *Recurrence and Donsker.*

We cancelled this exercise but might come back to it later.

**Solution 2** — *Skorokhod's embedding.*

To be updated. Let  $X$  be a centered random variable with variance 1.

- (1) This is done by repetetively applying Skorokhod's embedding and the strong Markov property.
- (2) The strong LLN tells us that almost surely, for every  $t \in \mathbb{Q}$ ,  $\phi_n(t) \rightarrow t$  as  $n \rightarrow \infty$ . Then Dini's theorem gives that  $\phi_n$  converges to the identity uniformly on every compact.
- (3) Denote  $B^n : t \mapsto n^{-1/2}B_{nt}$ . Remark that all  $B^n$  are Brownian motions. Now if  $t$  is a multiple of  $1/n$ ,  $\tilde{S}_t^n = n^{-1/2}S_{nt} = n^{-1/2}B_{n^{-1}T_{[nt]}} = B^n(\phi_n(t))$ . Then for  $0 \leq t \leq T$ , there is  $u \in [0, T] \cap n^{-1}\mathbb{Z}$ ,  $|u - t| < 1/n$ . As a result,

$$\begin{aligned} |\tilde{S}_t^n - B_t^n| &\leq |\tilde{S}_u^n - B_u^n| + |\tilde{S}_u^n - \tilde{S}_t^n| + |B_u^n - B_t^n| \\ &\leq |B^n \circ \phi_n(u) - B_u^n| + n^{-1/2}|X_{nu}| + |B_u^n - B_t^n| \end{aligned}$$

Taking the sup,

$$\|\tilde{S}^n - B^n\|_{[0,T]} \leq m_{2T}(B^n, \|\phi_n - id\|_{[0,T]}) + \infty \mathbf{1}_{\|\phi_n - id\| > 1} + n^{-1/2} \sup_{1 \leq i \leq nT} X_i + m_T(B^n, n^{-1}),$$

and each term indeed goes to 0 in probability. For the third one, thanks to the second moment hypothesis, we have by Chebychev's inequality  $\mathbb{P}(X > \varepsilon n^{1/2}) = o(n)$ .

- (4) We have  $B^n \rightarrow B$  in distribution (actually the distribution is constant!) and  $d(\tilde{S}^n, B^n) \rightarrow 0$  in probability. Then a classic generalization of Slutsky's lemma tells us that  $\tilde{S}^n \rightarrow B$  in distribution.

**Solution 3** — *Donsker's theorem for bridges.*

In this exercise, let  $b(n, p, k)$  denote the probability that a binomial of parameters  $(n, p)$  equals  $k$ , and  $f$  denote the standard Gaussian density. We will make use of the following *local limit theorem*, which is a refinement of the central limit theorem.

**Theorem.** (De Moivre–Laplace) As  $n \rightarrow \infty$ ,

$$\sup_{k \in \mathbb{Z}} \left| \sqrt{p(1-p)}nb(n, p, k) - f \left( \frac{k - np}{\sqrt{p(1-p)n}} \right) \right| = o(n^{-1/2}).$$

Recall (a few sessions back) that the Brownian bridge  $\beta$  (from 0 to 0) has the following property: for every integrable function  $H$  and  $\varepsilon > 0$ ,

$$\mathbb{E} [H(\beta_{|[0,1-\varepsilon]})] = \mathbb{E} \left[ H(B_{|[0,1-\varepsilon]}) \frac{\varepsilon^{-1/2} f(B_{1-\varepsilon})}{f(0)} \right].$$

Define the simple random walk  $S$  and its interpolated and rescaled version  $\tilde{S}^n$ . Our goal is to show that the distribution of  $\tilde{S}^{2n}$  **given that it is a bridge** (i.e.  $\tilde{S}_{2n}^{2n} = 0$ ), converges to that of  $\beta$  as  $n \rightarrow \infty$ .

(1) For  $j_0, j_1, \dots, j_{2k_n}$  fixed integers,

$$\begin{aligned} & \mathbb{P}(S_0 = j_0, \dots, S_{2k_n} = j_{2k_n} \mid S_{2n} = 0) \\ &= \frac{\mathbb{P}(S_0 = j_0, \dots, S_{2k_n} = j_{2k_n}) \mathbb{P}(S_{2n} = 0 \mid S_0 = j_0, \dots, S_{2k_n} = j_{2k_n})}{\mathbb{P}(S_{2n} = 0)} \\ &= \mathbb{P}(S_0 = j_0, \dots, S_{2k_n} = j_{2k_n}) \frac{\mathbb{P}(S_{2n} = 0 \mid S_{2k_n} = j_{2k_n})}{\mathbb{P}(S_{2n} = 0)} \\ &= \mathbb{P}(S_0 = j_0, \dots, S_{2k_n} = j_{2k_n}) \frac{b(\frac{1}{2}, 2n - 2k_n, n - k_n + \frac{j_{2k_n}}{2} \tilde{S}_{k_n/n})}{b(\frac{1}{2}, 2n, n)} \end{aligned}$$

Now since  $H(\tilde{S}_{|[0, k_n/n]}^{2n})$  is a deterministic function of  $S_0, \dots, S_{2k_n}$ , we get the desired result.

(2) It is clear that the central binomial coefficient bounds all the others. Hence the bound of  $A_n$  by  $b(\frac{1}{2}, 2n - 2k_n, n - k_n)/b(\frac{1}{2}, 2n, n)$ . By de Moivre-Laplace, this sequence converges to  $\frac{f(0)}{\sqrt{\varepsilon}f(0)}$  so is bounded uniformly in  $n$  (by  $C_\varepsilon$ , say). Then for  $s, t$  with  $|s - t| < 1/2$ , we can without loss of generality assume that  $s < t < 3/4$  (otherwise reverse everything). Then  $\mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4 \mid \tilde{S}_{2n} = 0] \leq C_{3/4} \mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4]$ . You know from the proof of Donsker's theorem that this  $\mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4]$  is bounded by  $c|s - t|^{1+\gamma}$  with  $c, \gamma > 0$ . We get

$$\forall s, t, |t - s| < 1/2, \quad \mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4 \mid \tilde{S}_{2n} = 0] \leq C_{3/4} c |t - s|^{1+\gamma}$$

If  $|t - s| \geq 1/2$ , then  $\mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4 \mid \tilde{S}_{2n} = 0]$  is uniformly bounded, (by the  $L^4$  triangle inequality we can reuse the case  $|t - s| < 1/2$ , and bound by some constant  $D$  independent of  $n, t - s$ ). Hence we get

$$\forall s, t, \quad \mathbb{E}[|\tilde{S}_t^{2n} - \tilde{S}_s^{2n}|^4 \mid \tilde{S}_{2n} = 0] \leq D 2^{1+\gamma} C_{3/4} c |t - s|^{1+\gamma}$$

proving a tightness bound for the random walk under the conditioned measure.

(3) From now on we suppose  $k_n/n = 1 - \varepsilon + o(1/n)$ , for instance by taking  $k_n = n^{-1}\lfloor n(1 - \varepsilon) \rfloor$ .

$$\begin{aligned}
A_n &= \frac{b(\frac{1}{2}, 2n - 2k_n, n - k_n + \frac{\sqrt{2n}}{2}\tilde{S}_{k_n/n})}{b(\frac{1}{2}, 2n, n)} \\
&= \frac{1}{\sqrt{1 - \frac{k_n}{n}}} \frac{\sqrt{2n - 2k} b(\frac{1}{2}, 2n - 2k_n, n - k_n + \frac{\sqrt{2n}}{2}\tilde{S}_{k_n/n})}{\sqrt{2n} b(\frac{1}{2}, 2n, n)} \\
&= (\varepsilon^{-1/2} + o(n^{-1})) \frac{f\left(\frac{\sqrt{2n}}{2}\tilde{S}_{k_n/n}^2 \cdot \frac{2}{\sqrt{2n-2k_n}}\right) + o(n^{-1/2})}{f(0) + o(n^{-1/2})} = \frac{\varepsilon^{-1/2} f(\varepsilon^{-1/2}\tilde{S}_{k_n/n}^2)}{f(0)} + o(n^{-1/2})
\end{aligned}$$

Consider some f.d.m. Without loss of generality we can always assume that it contains 1. Hence set  $0 \leq t_1 < \dots < t_r = 1$  and  $G : \mathbb{R}^r \rightarrow \mathbb{R}$  continuous with compact support. Take  $\varepsilon$  such that  $1 - \varepsilon > t_{r-1}$ . Now consider only  $n$  large enough so that  $k_n/n > t_{r-1}$ . where  $H$  is some continuous functional. Hence we can use question 2. Thus

$$\begin{aligned}
&\mathbb{E}[G(\tilde{S}_{t_1}^{2n}, \dots, \tilde{S}_{t_r}^{2n}) \mid \tilde{S}_{2n} = 0] \\
&= \mathbb{E}[G(\tilde{S}_{t_1}^{2n}, \dots, \tilde{S}_{t_{r-1}}^{2n}, 0) \mid \tilde{S}_{2n} = 0] \quad (\text{a.s. under } \mathbb{P}(\cdot \mid \tilde{S}_{2n} = 0), S^{2n} = 0) \\
&= \mathbb{E}\left[G(\tilde{S}_{t_1}^{2n}, \dots, \tilde{S}_{t_{r-1}}^{2n}, 0) A_n\right] \quad (\text{question 2, the integrand is a function of } \tilde{S}_{[0, k_n/n]}^{2n}) \\
&= \|G\|_\infty o(n^{-1/2}) + \mathbb{E}\left[G(\tilde{S}_{t_1}^{2n}, \dots, \tilde{S}_{t_{r-1}}^{2n}, 0) \frac{\varepsilon^{-1/2} f(\varepsilon^{-1/2}\tilde{S}_{k_n/n}^{2n})}{f(0)}\right] \quad (\text{first part of the question}) \\
&= o(1) + o(1) + \mathbb{E}\left[G(B_{t_1}, \dots, B_{t_{r-1}}, 0) \frac{\varepsilon^{-1/2} f(\varepsilon^{-1/2}B_{1-\varepsilon})}{f(0)}\right] \quad (\text{unconditioned Donsker}) \\
&= o(1) + \mathbb{E}\left[G(\beta_{t_1}, \dots, \beta_{t_{r-1}}, 0)\right] \quad (\text{absolute continuity property of the bridge}) \\
&= o(1) + \mathbb{E}\left[G(\beta_{t_1}, \dots, \beta_{t_{r-1}}, \beta_{t_r})\right] \quad (\beta_1 = 0 \text{ almost surely})
\end{aligned}$$

(4) By the usual criterion for convergence in distribution of functions, we are done.

**Solution 4** — *The binary splitting martingale.* (1) We write

$$\begin{aligned}
X_{n+1} - X_n &= \mathbb{E}[X - X_n \mid \mathcal{G}_n] \\
&= \mathbb{E}[(X - X_n) \mathbf{1}_{X > X_n} \mid \mathcal{G}_n] \mathbf{1}_{X > X_n} + \mathbb{E}[(X - X_n) \mathbf{1}_{X < X_n} \mid \mathcal{G}_n] \mathbf{1}_{X < X_n}.
\end{aligned}$$

where we used the fact that the sign of  $(X - X_n)$  is  $\mathcal{G}_n$ -measurable. The first term is almost surely positive, the second one is almost surely negative, and almost surely only one of them is nonzero. Hence they almost surely they form a decomposition of  $X_{n+1} - X_n$  into a positive and negative part. Then

$$\begin{aligned}
|X_{n+1} - X_n| &= \mathbb{E}[(X - X_n) \mathbf{1}_{X > X_n} \mid \mathcal{G}_n] \mathbf{1}_{X > X_n} - \mathbb{E}[(X - X_n) \mathbf{1}_{X < X_n} \mid \mathcal{G}_n] \mathbf{1}_{X < X_n} \\
&= \mathbb{E}[|X - X_n| \mid \mathcal{G}_n].
\end{aligned}$$

- (2) We deduce  $\mathbb{E}[|X_n - X|] = \mathbb{E}[|X_{n+1} - X_n|]$ , and this last expression goes to 0 as  $(X_n)_n$  is  $L^1$ -convergent. Thus  $|X_n - X|$  goes to 0 in  $L^1$  and by uniqueness (up to a.s. equality) of the  $L^1$  limit we get that  $X_\infty = X$  a.s. Hence  $X_n$  converges a.s. and  $L^1$  to  $X$ .