

## Solutions for Exercise sheet 3: stopping times and Markov property

**Solution 1** — *Brownian motion on the circle.*

**Solution 2** — *The set of zeros of  $B$  is perfect.*

Almost surely 0 is an accumulation point of  $Z$  (lecture). By countable union, and strong Markov, every first 0 after a rational is an accumulation point of  $Z$ . If  $Z$  had an isolated point, it would be a first 0 after a rational. Hence it couldn't be isolated in  $Z$ .

**Solution 3** — . (1) Denote by  $I = \min_{0 \leq s \leq 1-t} (B_{t+s} - B_t)$  and  $S = \max_{0 \leq s \leq 1-t} (B_{t+s} - B_t)$ . Both  $I$  and  $S$  are independent of  $\mathcal{F}_t$ .

$$\begin{aligned} \mathbb{P}(L < t) &= \mathbb{P}(B_t > 0 \cap I < -B_t \cup B_t < 0 \cap S > -B_t) \\ &= 2\mathbb{P}(B_t < 0 \cap S > |B_t|) \\ &= \mathbb{P}(B_t \neq 0, S > |B_t|) = \mathbb{P}(S > |B_t|) = \mathbb{P}(|\tilde{B}_{1-t}| > |B_t|) \\ &= \mathbb{P}\left(\left|\frac{Z}{\tilde{Z}}\right| < \sqrt{(1-t)/t}\right) \end{aligned}$$

where  $Z, \tilde{Z}$  are two independent standard Gaussians. Then

$$\dots = \mathbb{P}(|\arg(\tilde{Z} + iZ)| < \arcsin(\sqrt{t})) = 2 \arcsin(\sqrt{t}).$$

(2) For now let  $\tilde{A} = \inf\{t \in [0, 1], B_t = \min_{[0,1]} B\}$ . Then  $\mathbb{P}(\tilde{A} > t) = \mathbb{P}(S > \max_{[0,t]} B - B_t)$ . Once again  $S$  is independent of  $\max_{[0,t]} B - B_t \in \mathcal{F}_t$ , whose distribution is known to be equal to that of  $|B_t|$ . We end up with  $\mathbb{P}(\tilde{A} > t) = \mathbb{P}(L < t)$ . By symmetry of the arcsine distribution, we have shown that  $\tilde{A} \stackrel{d}{=} L$ .

To show that  $A$  is well-defined, consider  $\tilde{\tilde{A}} = \sup\{t \in [0, 1], B_t = \min_{[0,1]} B\}$ . By time reversal and symmetry of the arcsine distribution,  $\tilde{\tilde{A}} \stackrel{d}{=} \tilde{A}$ . At the same time,  $\tilde{\tilde{A}} \leq \tilde{A}$  almost surely. This implies that  $\tilde{A} = \tilde{\tilde{A}}$  almost surely.

**Solution 4** — *Markov processes derived from Brownian motion.* (1) Let  $B = (B^{(1)}, B^{(2)})$ . We have that  $(C_{a+} - C_a)$  is constructed from  $B_{T_{a+}^{(1)+}} - B_{T_a^{(1)}}$  the same way  $C$  is constructed from  $B$ . Hence by the strong Markov property of  $B$ ,  $(C_{a+} - C_a) \stackrel{d}{=} C$ , and  $(C_{a+} - C_a) \perp\!\!\!\perp \mathcal{F}_{T_{a+}^{(1)}} \supset \sigma(C_u, u \leq a)$ .

- (2)  $C$  is càdlàg because  $T_+$  is. By independence of  $B^{(1)}$  and  $B^{(2)}$  it jumps almost surely when  $T_+$  jumps.
- (3) Firstly,  $\text{Cov}(X_t, X_s) = e^{-|t-s|}$ . So at each time  $t$ ,  $X_t$  is a standard Gaussian.

**Solution 5** — *All hypotheses matter.*

Take  $S = 3$  and  $T$  to be the first zero after 3. Of course the problem is that  $\mathbb{E}[T] = \infty$ .

**Solution 6** — *Brownian gambler's ruin.*

Let  $a < 0 < b$  and  $T$  be the hitting time of  $\{a, b\}$ .

- (1) We may show that  $T$  is integrable to apply Wald's second lemma. Here's a way to do it by comparison with a geometric variable. Let  $x \leq |a| \wedge |b|$ .

$$\begin{aligned} \mathbb{P}(T \geq n) &\leq \mathbb{P}(\forall k \leq n-1, \quad |B_{k+1} - B_k| < 2x) \\ &= \prod_{k=0}^{n-1} \mathbb{P}(|B_1^{(k)}| < 2x) = \rho^n \end{aligned}$$

Where  $\rho < 1$ . Hence  $T$  is integrable, and we can apply Wald's second lemma. We get

$$E[T] = \mathbb{E}[B_T^2] = \frac{-a}{b-a} b^2 + a^2 \frac{b}{b-a} = -ab$$

- (2) Let  $M = \sup_{0 \leq t \leq T_{a,b}} B_t$ . Let  $c \in [0, b]$ . We denote  $\tilde{B}_t = B_{t+T_c} - B_{T_c}$ .

$$\begin{aligned} \mathbb{P}(M \geq c \mid T_a \leq T_b) &= \frac{\mathbb{P}(T_c \leq T_a \leq T_b)}{\mathbb{P}(T_a \leq T_b)} \\ &= \frac{\mathbb{P}(T_c \leq T_a, \tilde{T}_{a-c} \leq \tilde{T}_{b-c})}{\mathbb{P}(T_a \leq T_b)} \\ &= \frac{\mathbb{P}(T_c \leq T_a) \mathbb{P}(\tilde{T}_{a-c} \leq \tilde{T}_{b-c})}{\mathbb{P}(T_a \leq T_b)} \\ &= \frac{\frac{-a}{c-a} \frac{b-c}{b-a}}{\frac{b}{b-a}} = \frac{-a(b-c)}{b(c-a)}. \end{aligned}$$

**Solution 7** — *Girsanov theorem and hitting times with drift.*

Let  $B$  be a brownian motion, and for  $\lambda \in \mathbb{R}$ , denote  $M_t^\theta = e^{\theta B_t - \theta^2 t/2}$ . You have shown that  $M^\lambda$  is a  $(\mathcal{F}_t)_t$ -martingale, and used it to show that  $\mathbb{E}[e^{-\lambda T_b}] = \mathbb{E}[e^{-|b|\sqrt{2\lambda}}]$ .

- (1) We check that

$$\mathbb{P}_{\theta, T}(\Omega) := \mathbb{E}[M_T^\theta] = \mathbb{E}[e^{\theta B_T}] e^{-\theta^2 T/2} = 1.$$

We use characteristic functions of fdms to characterize distribution of a process. Let  $0 \leq t_1 \leq \dots \leq t_k \leq T$  and  $u_1, \dots, u_k \in R$

$$\begin{aligned}
\mathbb{E}_{\theta, T}[e^{i(u_1 B_{t_1} + \dots + u_k B_{t_k})}] &= \mathbb{E}[e^{i(u_1 B_{t_1} + \dots + u_k B_{t_k})} e^{\theta B_T - \theta^2 T/2}] \\
&= \mathbb{E}[e^{i(u_1 B_{t_1} + \dots + u_k B_{t_k} - i\theta B_T)}] e^{\theta^2 T/2} \\
&= \mathbb{E}[e^{-\text{Var}(u_1 B_{t_1} + \dots + u_k B_{t_k} - i\theta B_T)/2}] e^{\theta^2 T/2} \\
&= \mathbb{E}[e^{-\text{Var}(u_1 B_{t_1} + \dots + u_k B_{t_k})/2 + \theta^2 T/2 + i\theta(u_1 t_1 + \dots + u_k t_k)}] e^{\theta^2 T/2} \\
&= \mathbb{E}[e^{i(u_1 B_{t_1} + \dots + u_k B_{t_k})}] e^{i\theta(u_1 t_1 + \dots + u_k t_k)} \\
&= \mathbb{E}[e^{i(u_1 (B_{t_1} + \theta t_1) + \dots + u_k (B_{t_k} + \theta t_k))}]
\end{aligned}$$

- (2) No, Brownian motion with and without drift are not absolutely continuous to each other over  $\mathbb{R}_+$ .
- (3) Let  $T_b^\theta$  be the hitting time of  $b$  by  $(B_t + \theta t)_t$ . Then using question 1 and optional stopping,

$$\begin{aligned}
\mathbb{E}[e^{-\lambda T_b^\theta} \mathbf{1}_{T_b^\theta < U}] &= \mathbb{E}[e^{-\lambda T_b} \mathbf{1}_{T_b < U} M_U^\theta] \\
&= \mathbb{E}[e^{-\lambda T_b} \mathbf{1}_{T_b < U} M_{T_b}^\theta] \\
&= \mathbb{E}[e^{-\lambda T_b} \mathbf{1}_{T_b < U} e^{\theta b - \theta^2 T_b/2}] \\
&= e^{\theta b} \mathbb{E}[e^{-(\lambda - \theta^2/2) T_b} \mathbf{1}_{T_b < U}]
\end{aligned}$$

Using dominated convergence, we get

$$\mathbb{E}[e^{-\lambda T_b^\theta} \mathbf{1}_{T_b^\theta < \infty}] = e^{\theta b} \mathbb{E}[e^{-(\lambda - \theta^2/2) T_b}] = e^{-|b|\sqrt{\theta^2 + 2\lambda} + b\theta}.$$

- (4) Then taking  $\lambda = 0$ ,  $\mathbb{P}(T_b^\lambda < \infty) = e^{-|b\theta| + b\theta} = e^{2b\theta \wedge 0}$ . We observe, that

$$\mathbb{E}[e^{-\lambda T_b^\theta} \mid T_b^\theta < \infty] = e^{-|b|\sqrt{\theta^2 + 2\lambda} - |b\theta|}.$$

It is independent on the sign of  $b$ . So a Brownian motion with negative drift, conditioned on hitting a positive level, will behave as a Brownian motion with the reverse (positive) drift.

**Solution 8** — *The binary splitting martingale.* (1) We write

$$\begin{aligned}
X_{n+1} - X_n &= \mathbb{E}[X - X_n \mid \mathcal{G}_n] \\
&= \mathbb{E}[(X - X_n) \mathbf{1}_{X > X_n} \mid \mathcal{G}_n] \mathbf{1}_{X > X_n} + \mathbb{E}[(X - X_n) \mathbf{1}_{X < X_n} \mid \mathcal{G}_n] \mathbf{1}_{X < X_n}.
\end{aligned}$$

where we used the fact that the sign of  $(X - X_n)$  is  $\mathcal{G}_n$ -measurable. The first term is almost surely positive, the second one is almost surely negative, and almost surely only one of them is nonzero. Hence they almost surely they form a decomposition of  $X_{n+1} - X_n$  into a positive and negative part. Then

$$\begin{aligned}
|X_{n+1} - X_n| &= \mathbb{E}[(X - X_n) \mathbf{1}_{X > X_n} \mid \mathcal{G}_n] \mathbf{1}_{X > X_n} - \mathbb{E}[(X - X_n) \mathbf{1}_{X < X_n} \mid \mathcal{G}_n] \mathbf{1}_{X < X_n} \\
&= \mathbb{E}[|X - X_n| \mid \mathcal{G}_n].
\end{aligned}$$

- (2) We deduce  $\mathbb{E}[|X_n - X|] = \mathbb{E}[|X_{n+1} - X_n|]$ , and this last expression goes to 0 as  $(X_n)_n$  is  $L^1$ -convergent. Thus  $|X_n - X|$  goes to 0 in  $L^1$  and by uniqueness (up to a.s. equality) of the  $L^1$  limit we get that  $X_\infty = X$  a.s. Hence  $X_n$  converges a.s. and  $L^1$  to  $X$ .

**Solution 9** — *Martingales derived from  $B$ .*

Those martingales are the derivative w.r.t  $\lambda$  of the exponential martingale.